

Continuous Probability Density Functions

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Outline

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Inverse Gamma

Discussion

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Properties and Summaries

Assume that the continuous random variable (RV) x can take on values

$$x \in (a, b)$$

then, the probability density function (PDF) is given by

$$f(x|\theta) \text{ defined for } x \in (a, b)$$

where x can be defined within an infinite interval

and θ are any parameters that the PDF depends on.

Properties and Summaries

The accumulation of probability from left to right
is called the cumulative distribution function (CDF)

$$F(x | \theta) = \int_{t=-\infty}^x f(t | \theta) dt$$

and

$$1) \quad 0 \leq f(x | \theta)$$

$$2) \quad \int_x f(x | \theta) = 1 \quad .$$

Properties and Summaries

Given an arbitrary continuous probability density

$f(x|\theta)$, we want to compute quantitative population

summaries of it such as:

population mean, $\mu = \int_{x=-\infty}^{\infty} xf(x|\theta)dx$

population variance, $\sigma^2 = \int_{x=-\infty}^{\infty} (x - \mu)^2 f(x|\theta)dx$

population standard deviation, $\sigma = \sqrt{\sigma^2}$.

Properties and Summaries

Given an arbitrary continuous probability density

$f(x|\theta)$, we want to compute quantitative population

summaries of it such as:

population median $\tilde{x}$, $\int_{x=-\infty}^{\tilde{x}} f(x|\theta)dx = \frac{1}{2}$

population mode $\hat{x}$, $\left. \frac{\partial}{\partial x} f(x|\theta) \right|_{\hat{x}} = 0$.

Provided f is differentiable.

Max if 2nd derivative neg at point.

Check boundary points for max.

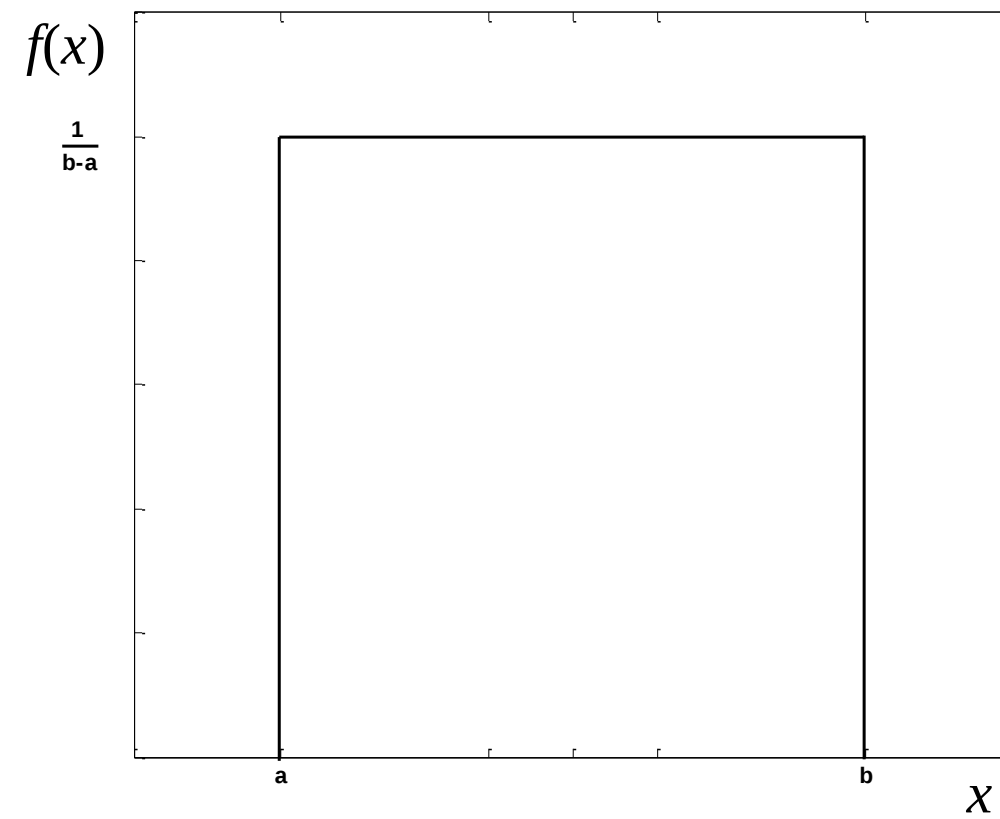
May need to be numerically calculated.

Uniform and Beta PDF

A random variable x has a continuous uniform distribution,
 $x \sim \text{uniform}(a, b)$ if

$$f(x|a, b) = \frac{1}{b-a}, \quad x \in [a, b]$$

where, $a < x < b$, $a < b$.



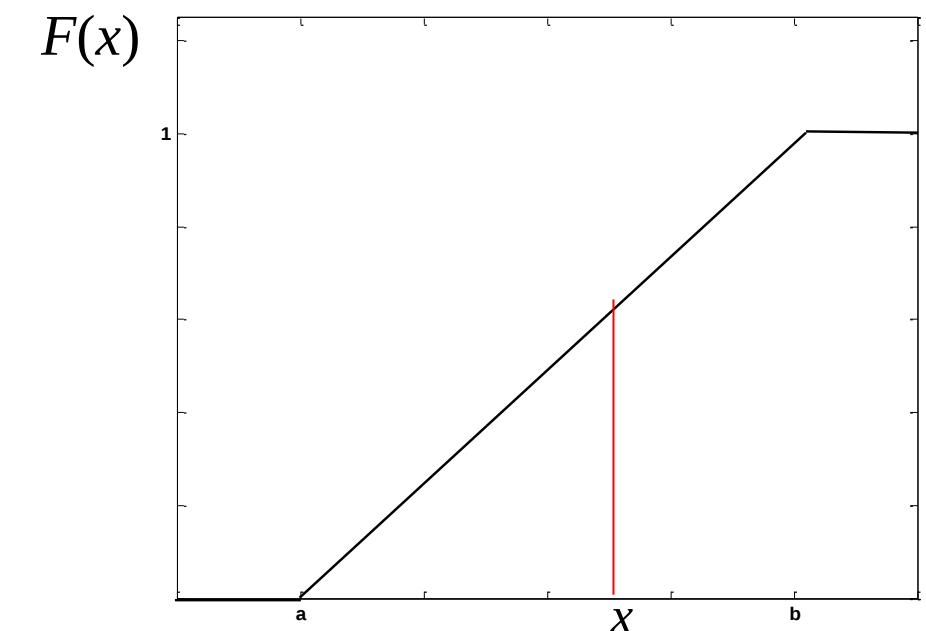
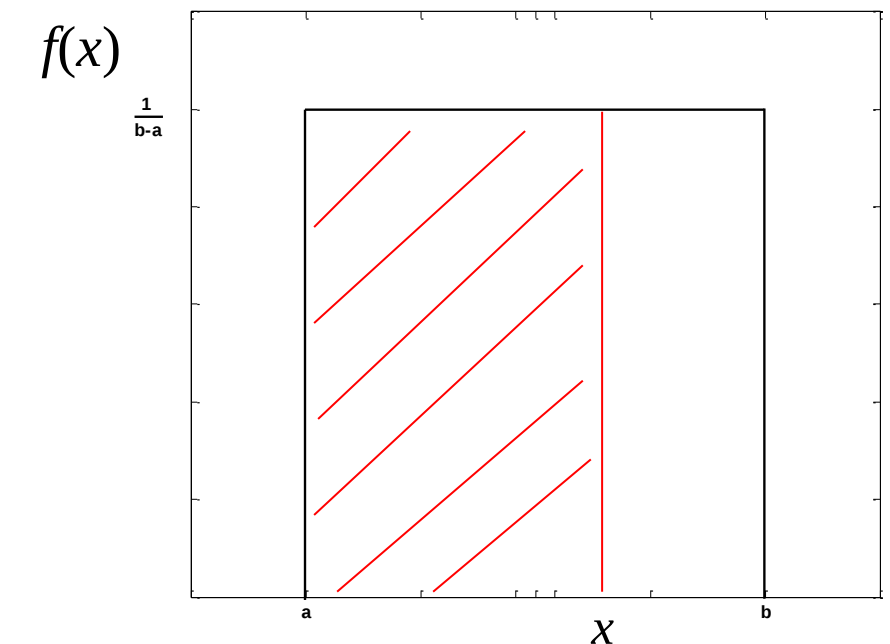
Uniform and Beta PDF

The CDF of the continuous uniform distribution is

$$F(x|\theta) = \int_{t=-\infty}^x f(t|\theta)dt$$

$$F(x|a,b) = \int_{t=a}^x \frac{1}{b-a}dt$$

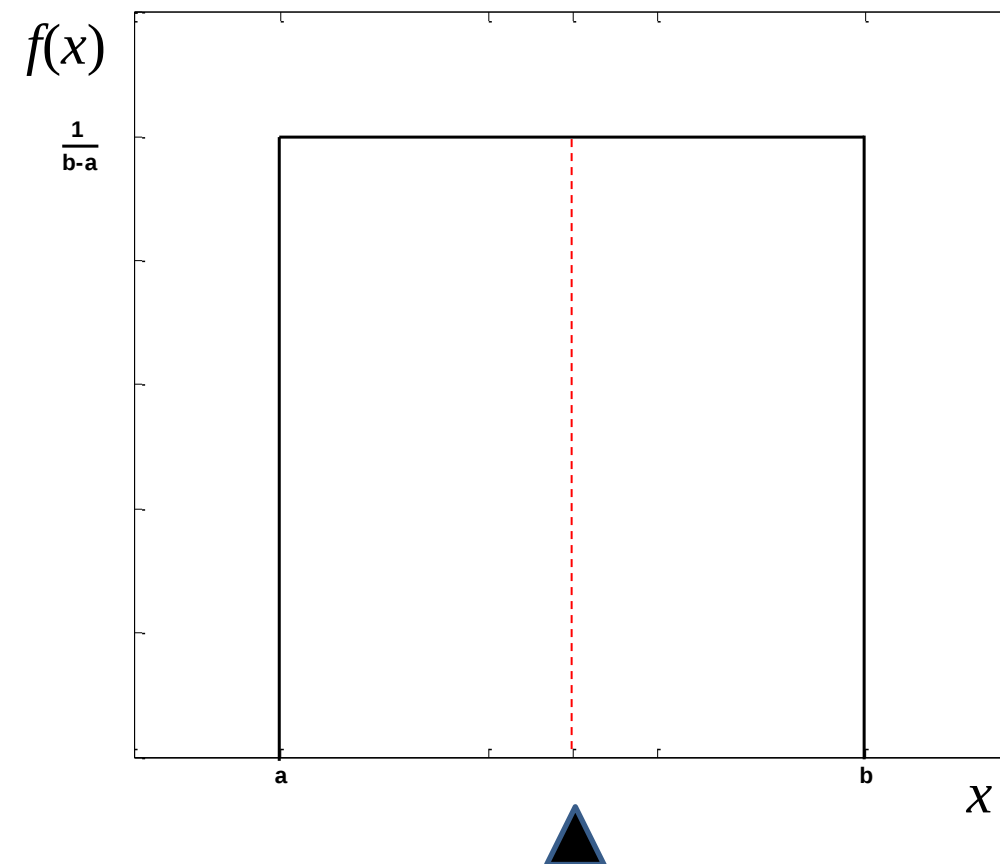
$$= \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & x \in [a,b] \\ 1 & x > b \end{cases}$$



Uniform and Beta PDF

It can be shown that

$$\begin{aligned}\mu &= \int_x x f(x|\theta) dx \\ &= \int_{x=a}^b x \frac{1}{b-a} dx \\ &= \frac{b+a}{2}\end{aligned}$$



Uniform and Beta PDF

It can be shown that

median

$$\int_{x=a}^{\tilde{x}} f(x|\theta) dx = \frac{1}{2}$$

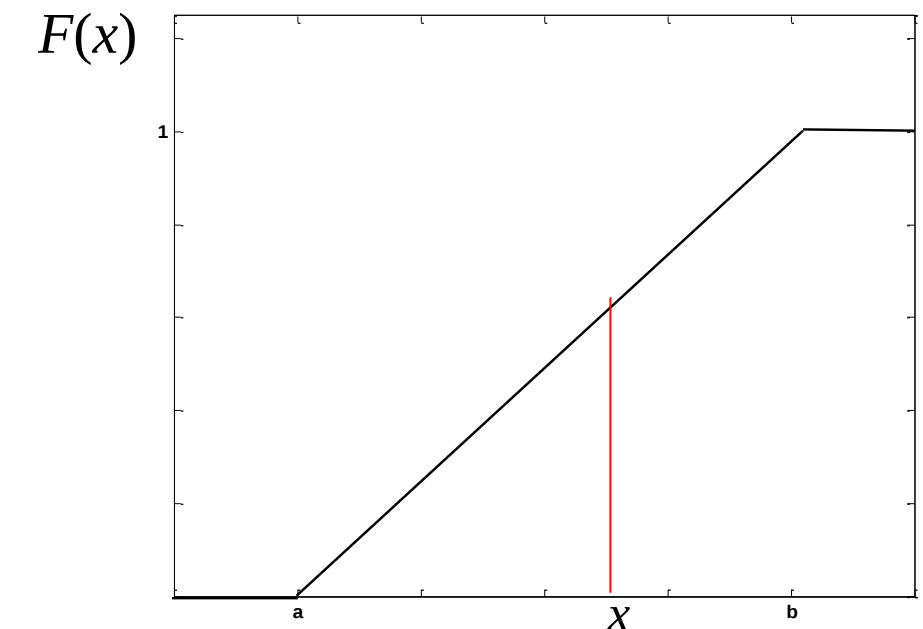
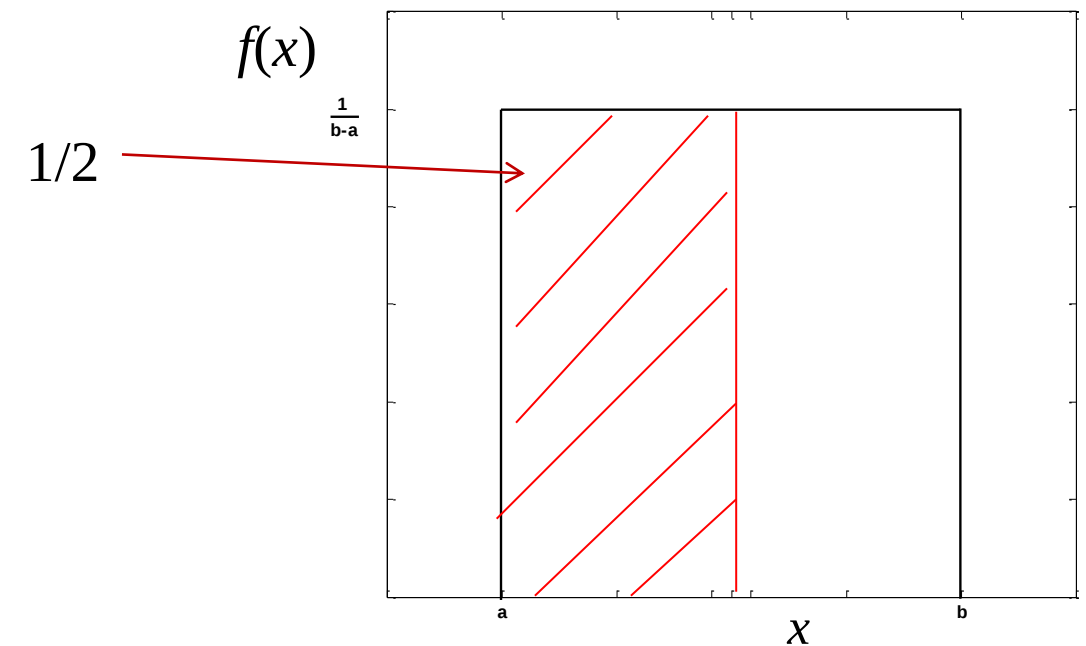
$$\tilde{x} = \frac{b+a}{2}$$

mode

$$\left. \frac{\partial}{\partial x} f(x|\theta) \right|_{\hat{x}} = 0$$

$$\hat{x} = \text{any value} \in [a, b]$$

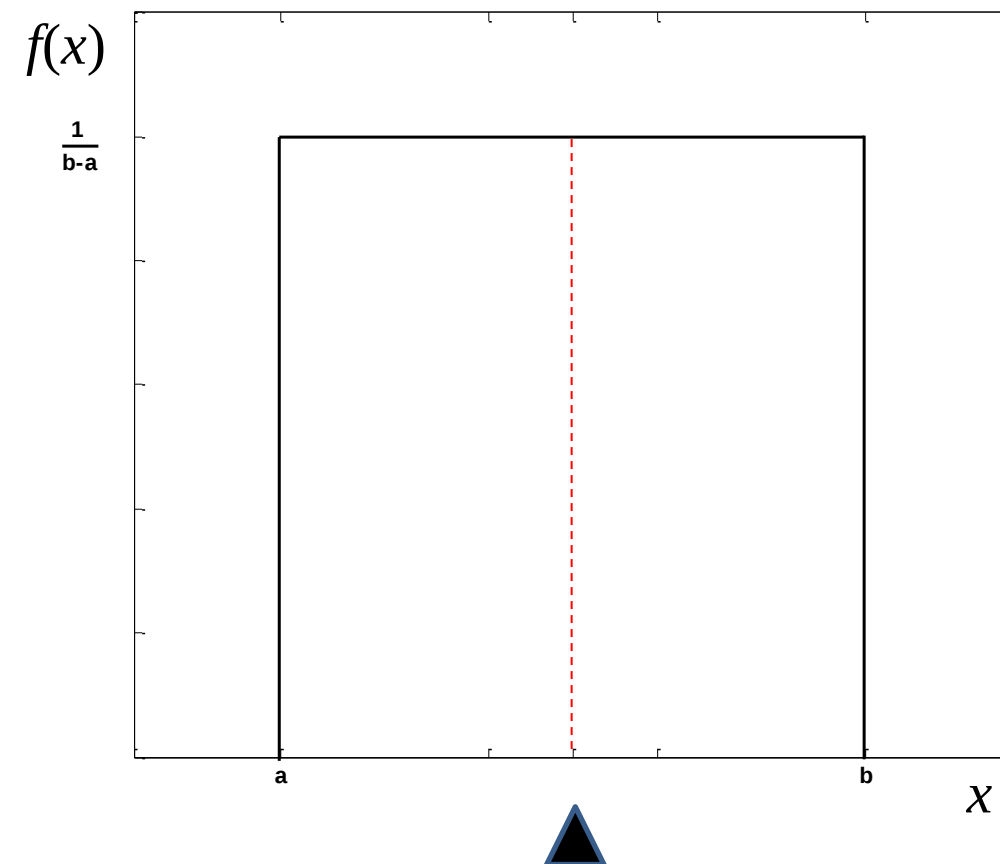
flat distribution



Uniform and Beta PDF

that

$$\begin{aligned}\sigma^2 &= \int_x (x - \mu)^2 f(x|\theta) dx \\ &= \int_{x=a}^b (x - \mu)^2 \frac{1}{b-a} dx \\ &= \frac{(b-a)^2}{12} \\ \sigma &= \frac{(b-a)}{\sqrt{12}}\end{aligned}$$



Uniform and Beta PDF

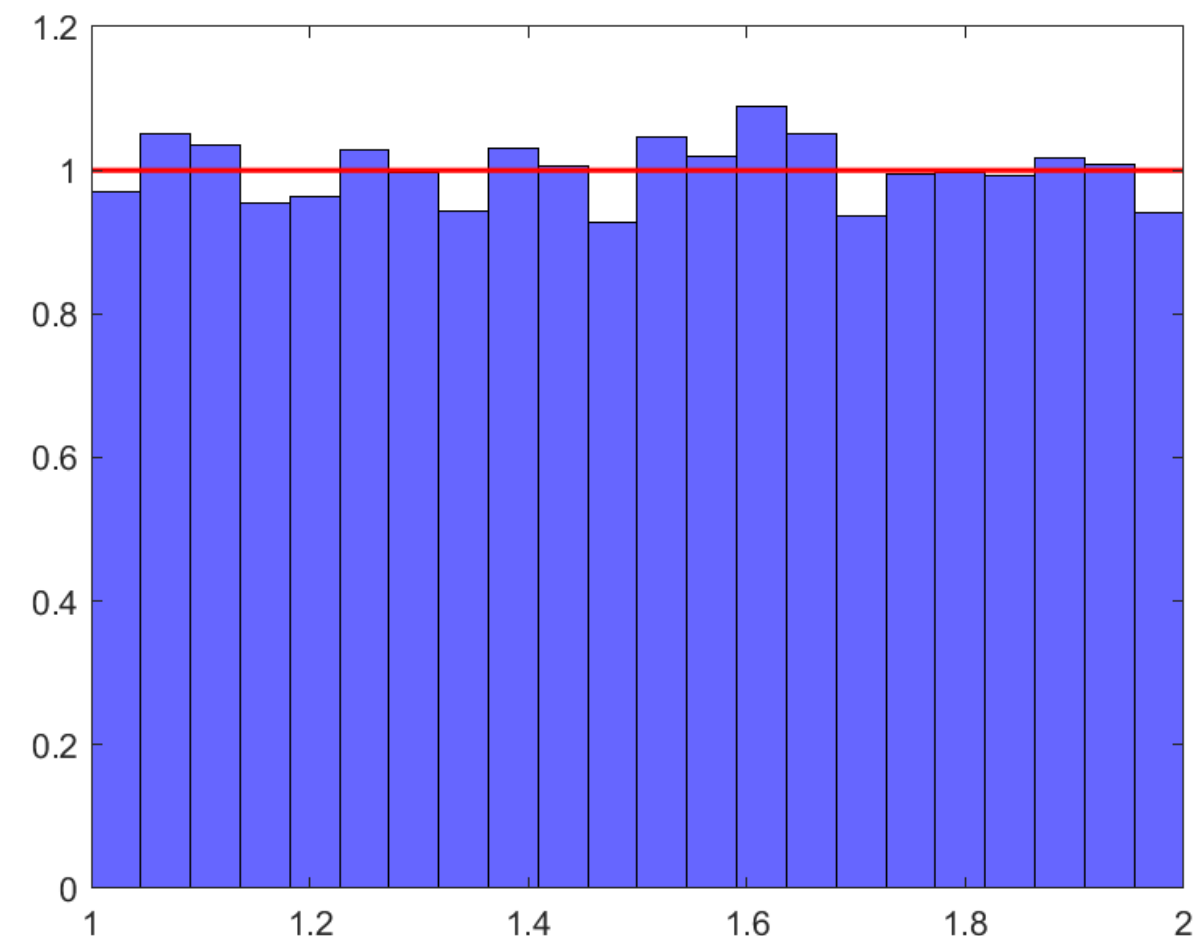
```
rng('default'), a=1;;b=2;;num=10^4;
x=a+(b-a)*rand(num,1);
figure;
histogram(x,'BinLimits',[1,2],'normalization','pdf','FaceColor','blue')
hold on
line([1,2],[1,1],'Color','red','LineWidth',1.5)
xlim([1,2])
```

	True	Simulated
μ	1.5	1.4996
σ^2	0.0833	0.0829

Can also find and plot ECDF.

```
themean=(b-a)/2
thevar=(b-a)^2/12
[mean(x),var(x)]
```

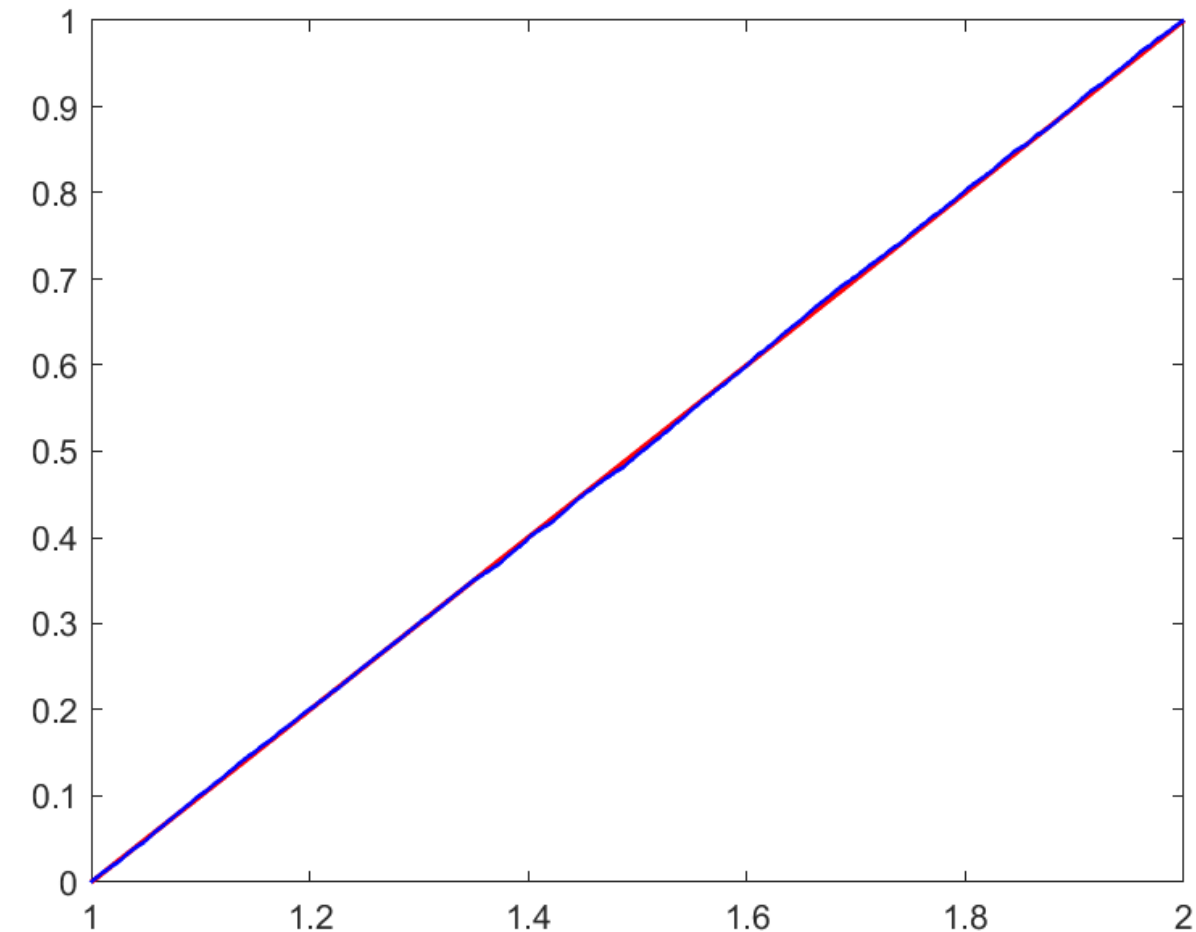
$$f(x|a,b) = \frac{1}{b-a}$$



Uniform and Beta PDF

```
y=cdf('unif',(a:.01:b),a,b);  
figure;  
plot((1:.01:2),y,'Color','red','LineWidth',1.5)  
[F,xx]=ecdf(x);  
hold on  
stairs(xx,F,'Color','blue','LineWidth',1.5)
```

$$F(x|a,b) = \frac{1}{b-a}x$$



Uniform and Beta PDF

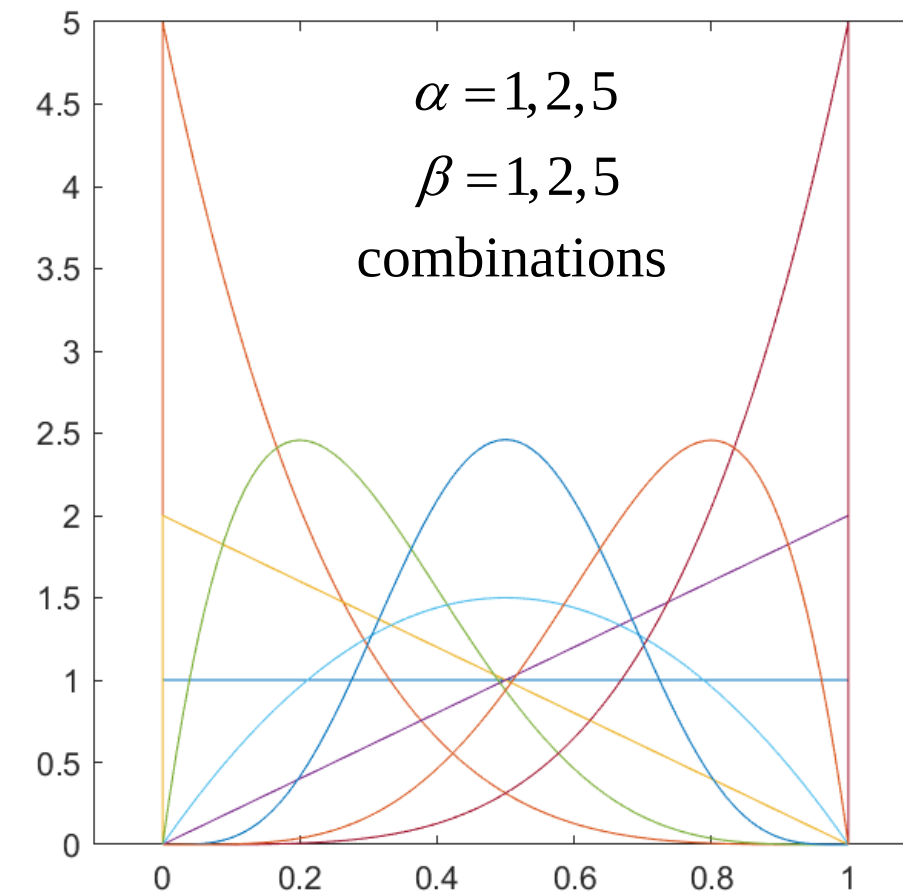
A random variable x has a continuous Beta distribution, $x \sim \text{Beta}(\alpha, \beta)$ if

$$f(x | \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1},$$

where, $\alpha, \beta > 0$, $x \in [a, b]$.

And can take on many shapes.

$$\Gamma(a) = (a-1)! \text{ for integer } a$$



Uniform and Beta PDF

There is no closed form analytic formula for the CDF $F(x|\alpha,\beta)$.

It can be shown that the mean is

$$\begin{aligned}\mu &= \int_x x f(x|\theta) dx \\ &= \int_{x=0}^1 x \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} dx \\ &= \frac{\alpha}{\alpha + \beta}\end{aligned}$$

Uniform and Beta PDF

It can be shown that the variance is

$$\begin{aligned}\sigma^2 &= \int_x (x - \mu)^2 f(x | \theta) dx \\ &= \int_{x=0}^1 (x - \mu)^2 \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} dx \\ &= \frac{\alpha\beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}\end{aligned}$$

Uniform and Beta PDF

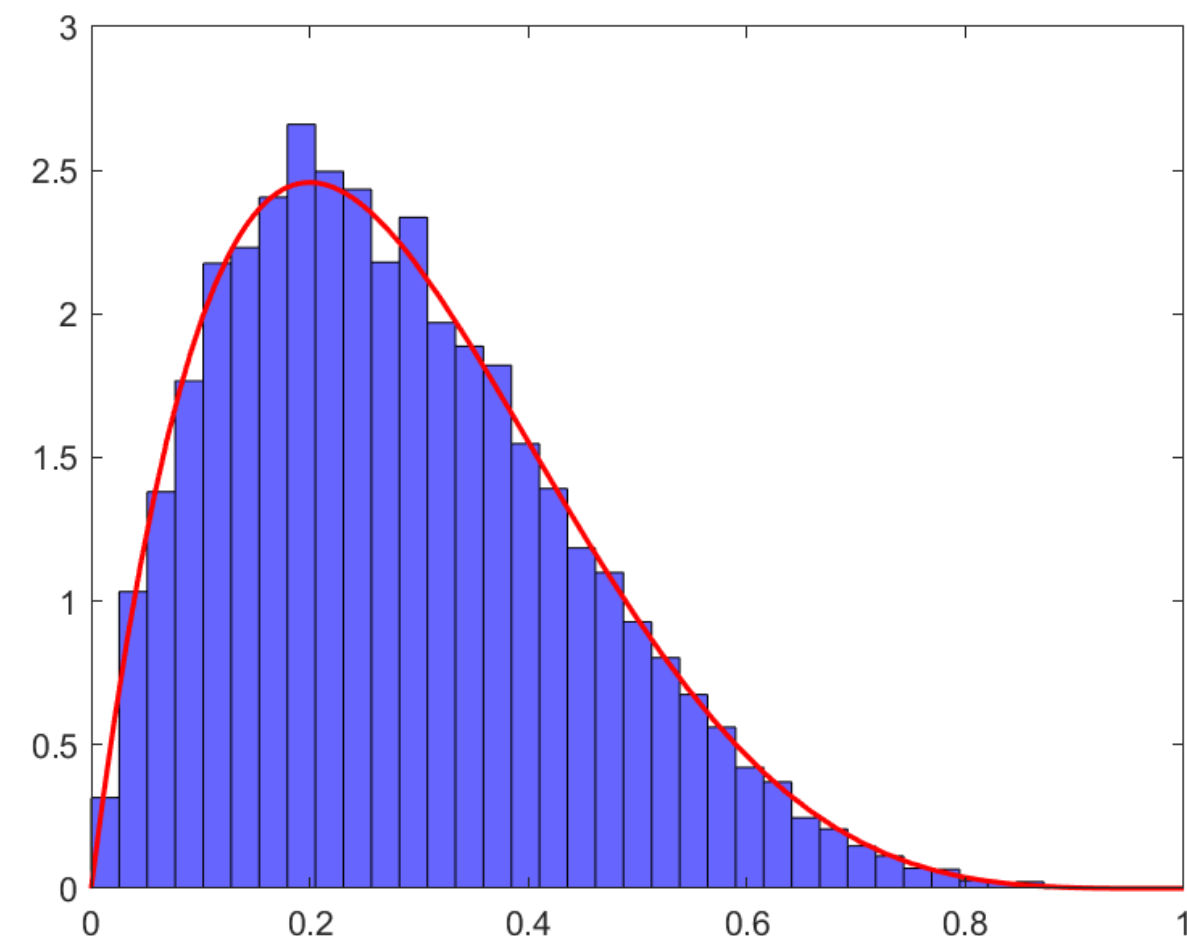
```
rng('default'), a=2; b=5; num=10^4;
x=betarnd(a,b,[num,1]);
figure;
histogram(x,'BinLimits',[0,1],'normalization','pdf','FaceColor','blue')
hold on
fx = @(x) factorial(a+b-1)/factorial(a-1)/factorial(b-1)...
.*x.^(a-1).*(1-x).^(b-1);
fplot(fx,[0,1],'r','LineWidth',1.5)
```

	True	Simulated
μ	0.2857	0.2858
σ^2	0.0255	0.0251

Can also find and plot ECDF.

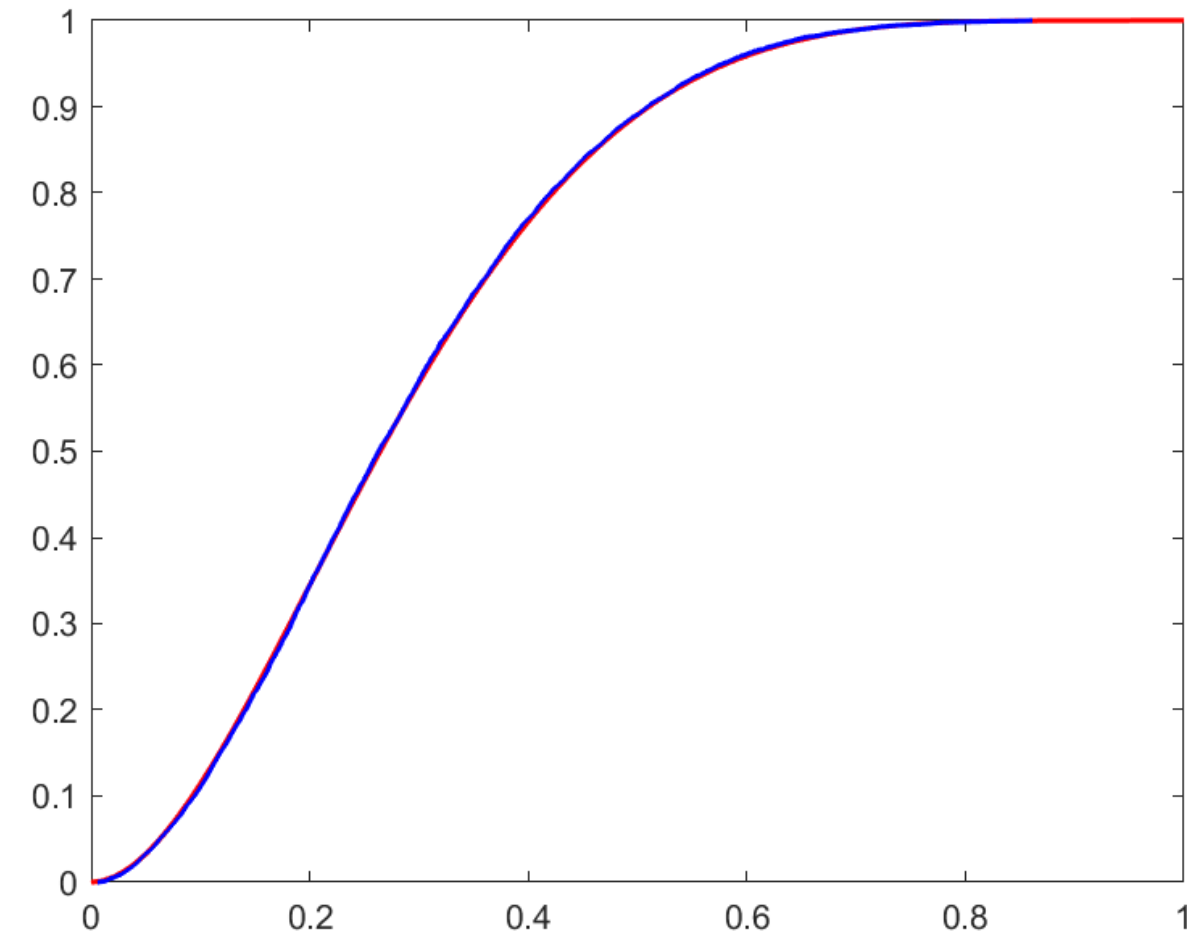
```
themean=a/(a+b)
thevar=(a*b)/((a+b)^2*(a+b+1))
[mean(x),var(x)]
```

$$f(x|\alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$



Uniform and Beta PDF

```
y=cdf('beta',(0:.01:1),a,b);  
figure;  
plot((0:.01:1),y,'Color','red','LineWidth',1.5)  
[F,xx]=ecdf(x);  
hold on  
stairs(xx,F,'Color','blue','LineWidth',1.5)
```

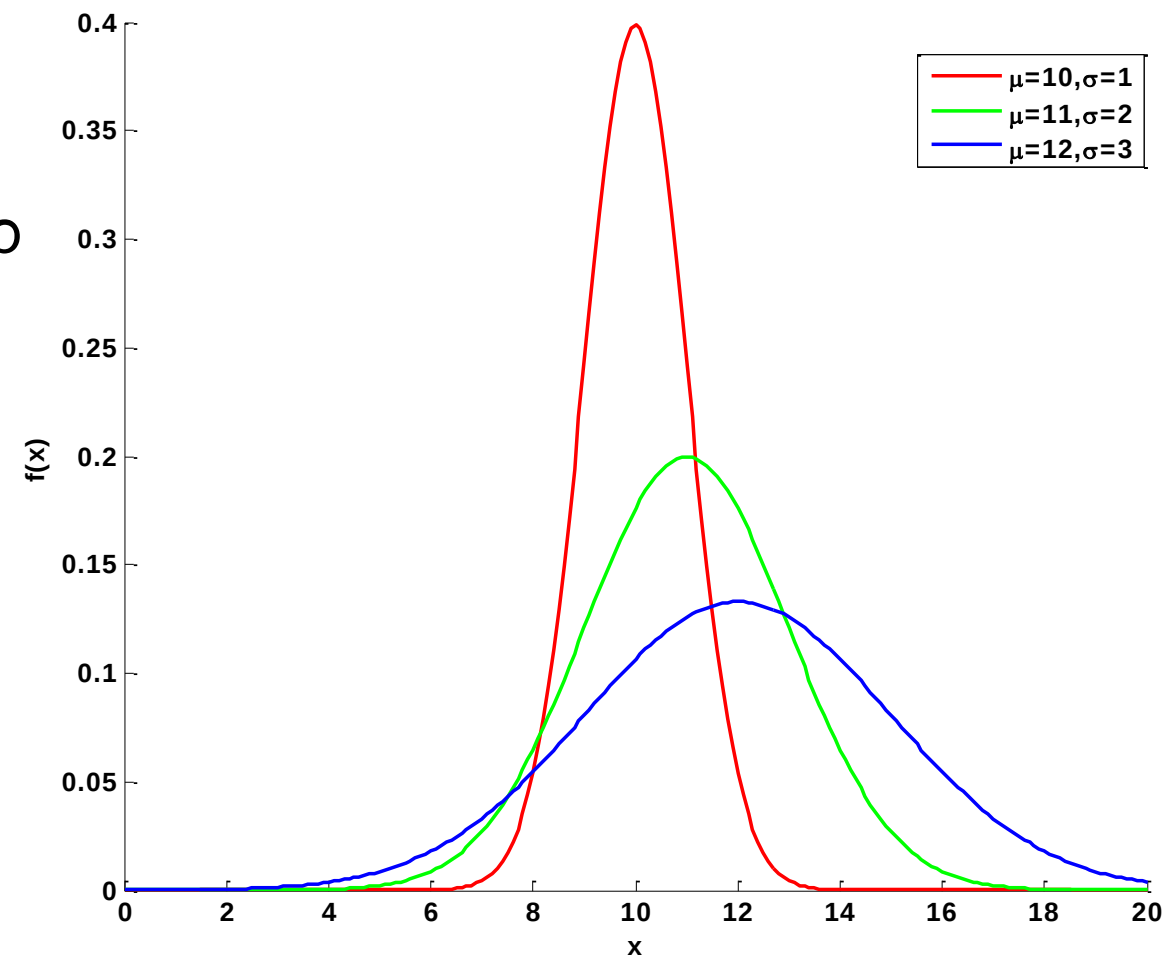


Normal PDF

A random variable x has a continuous normal distribution, $x \sim \text{normal}(\mu, \sigma^2)$ if

$$f(x | \mu, \sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}, \quad -\infty < x < +\infty$$

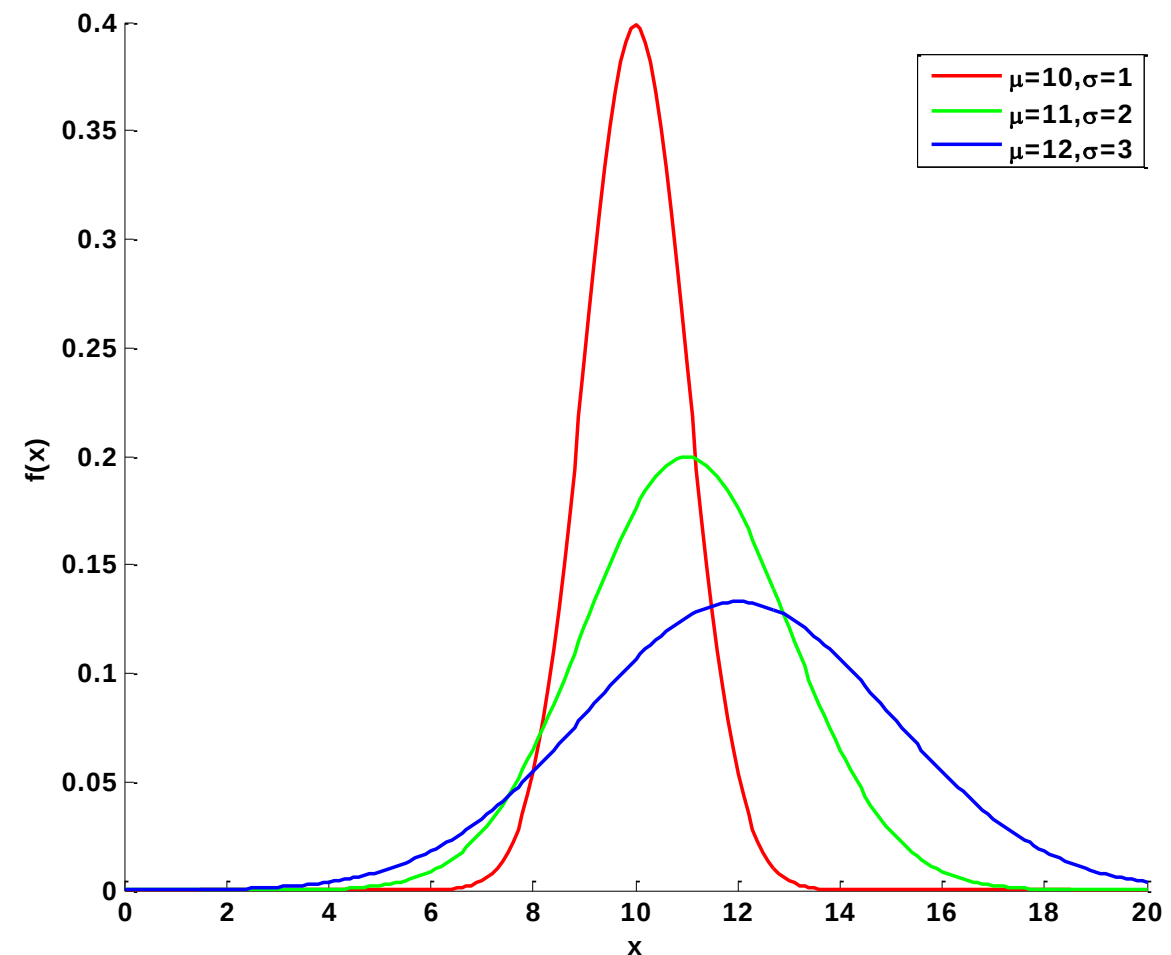
where, $-\infty < \mu < +\infty$ and $\sigma > 0$.



Normal PDF

```
x=(0:1:20)';  
mu=[10,11,12];, sigma=[1,2,3];  
figure(1)  
hold on  
for count=1:length(sigma)  
    y = normpdf(x,mu(count),sigma(count));  
    if count==1  
        plot(x,y,'r','LineWidth',2)  
    elseif count==2  
        plot(x,y,'g','LineWidth',2)  
    elseif count==3  
        plot(x,y,'b','LineWidth',2)  
    end  
end  
xlim([0 20])
```

$$f(x | \mu, \sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$



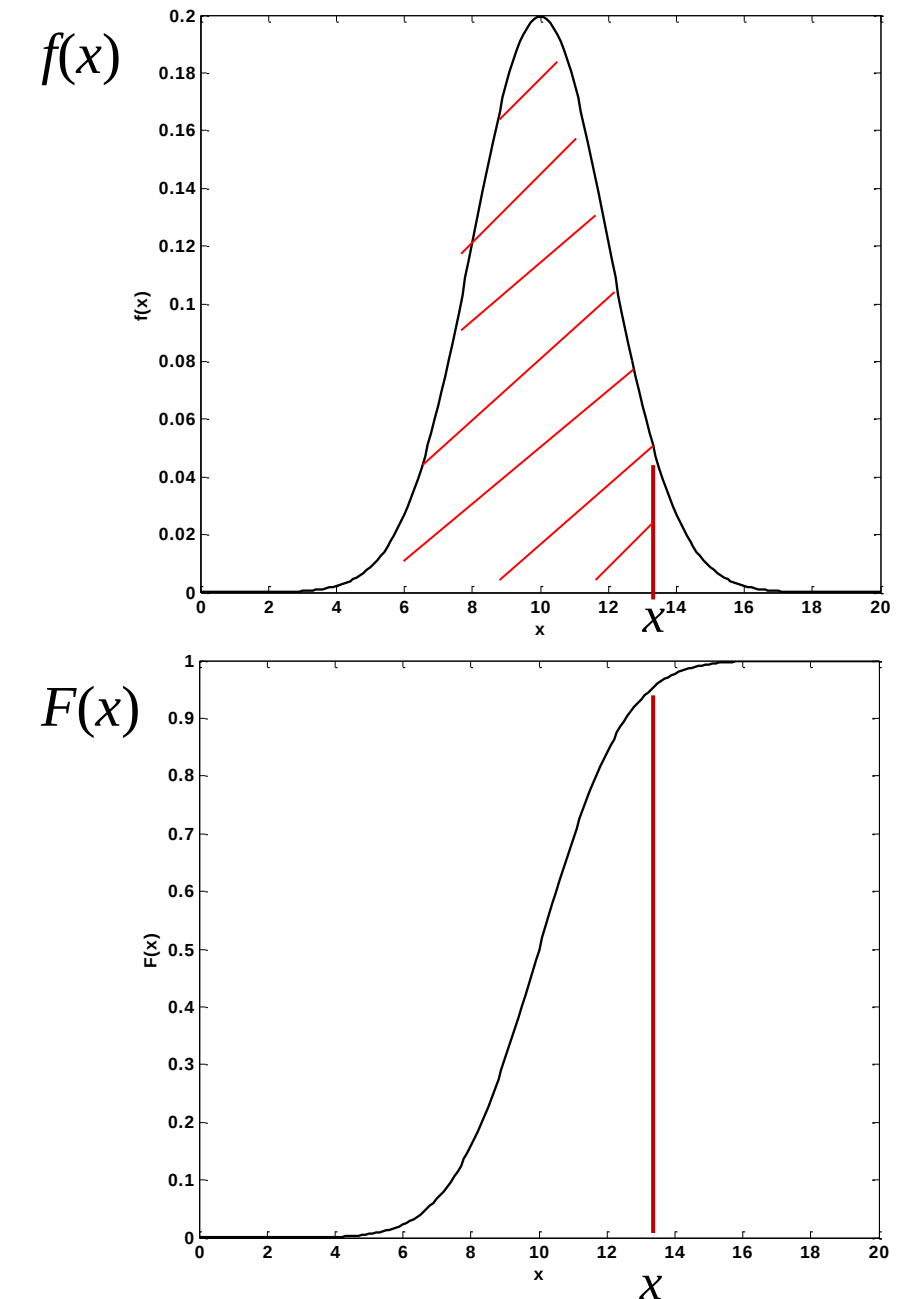
Normal PDF

The CDF of the continuous normal distribution is

$$F(x|\theta) = \int_{t=-\infty}^x f(t|\theta)dt$$

$$F(x|\mu, \sigma^2) = \int_{t=-\infty}^x \frac{e^{-\frac{(t-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}} dt$$

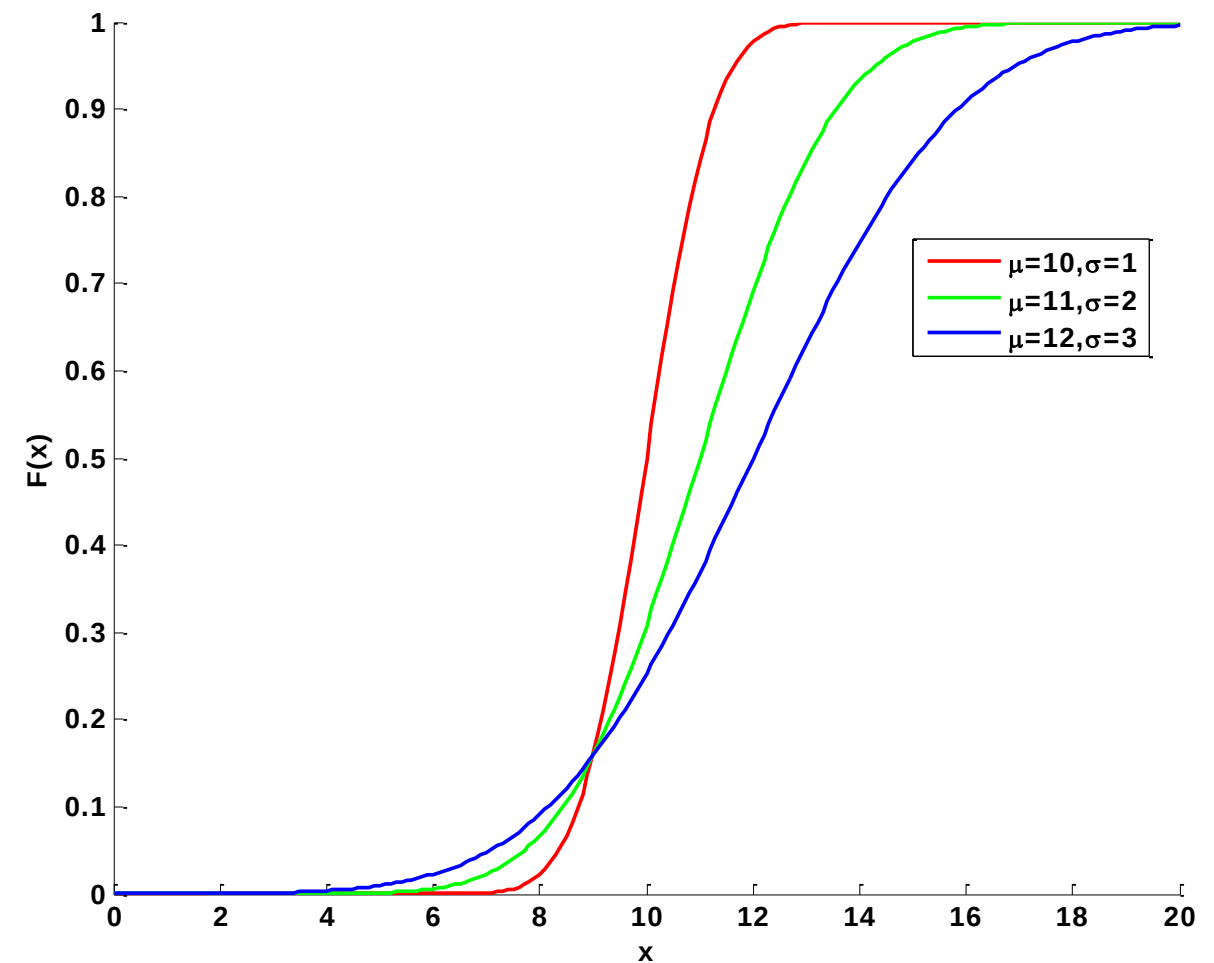
No closed form analytic solution.



Normal PDF

```
x=(0:1:20)';
mu=[10,11,12];, sigma=[1,2,3];
figure(1)
hold on
for count=1:length(sigma)
    y = normcdf(x,mu(count),sigma(count));
    if count==1
        plot(x,y,'r','LineWidth',2)
    elseif count==2
        plot(x,y,'g','LineWidth',2)
    elseif count==3
        plot(x,y,'b','LineWidth',2)
    end
end
end
xlim([0 20])
```

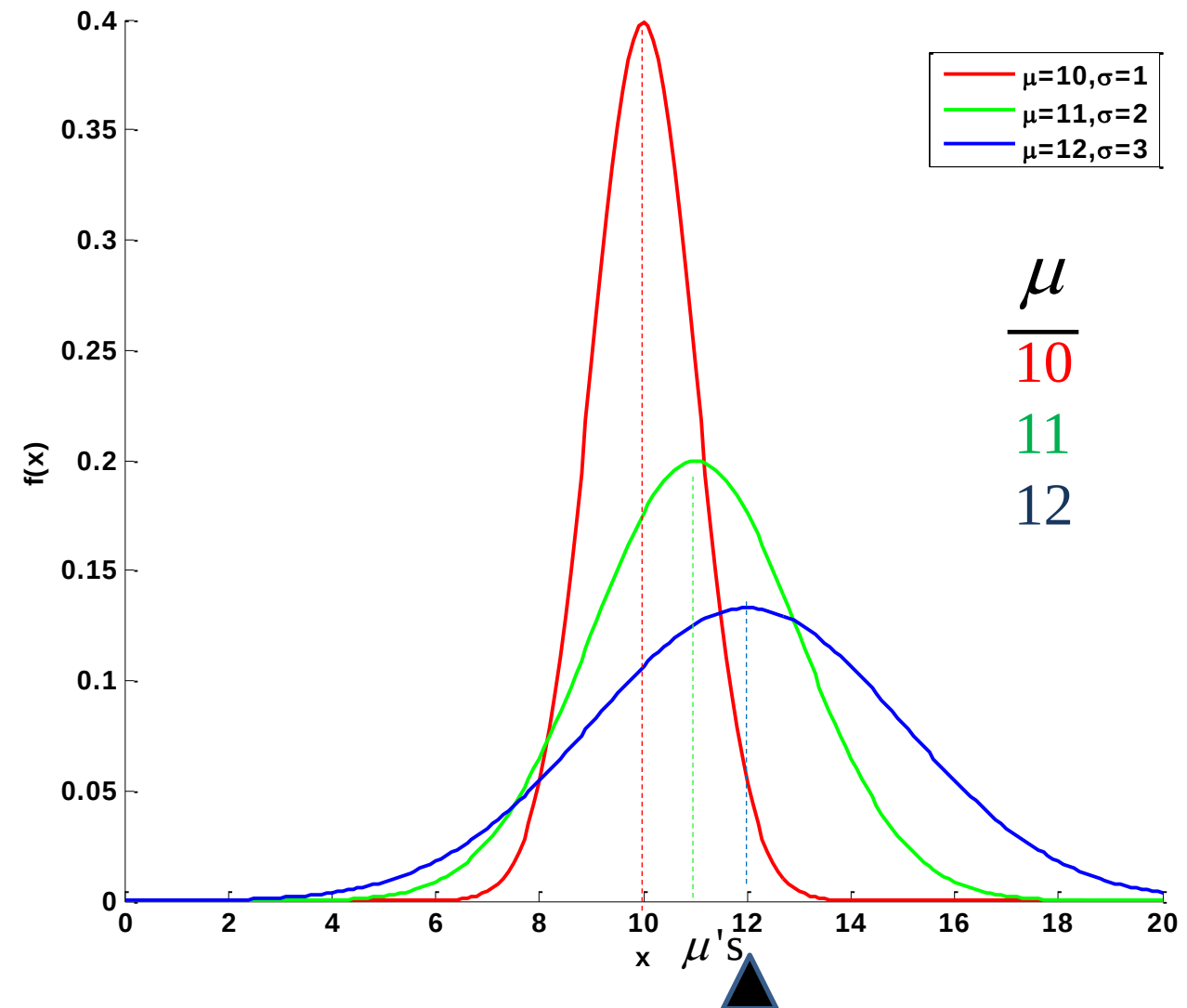
$$f(x | \mu, \sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$



Normal PDF

It can be shown that

$$\begin{aligned}\mu &= \int_{-\infty}^{\infty} xf(x|\theta)dx \\ &= \int_{-\infty}^{\infty} x \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}} dx\end{aligned}$$



Normal PDF

It can be shown that

median

$$\int_{x=-\infty}^{\tilde{x}} f(x|\theta)dx = \frac{1}{2}$$

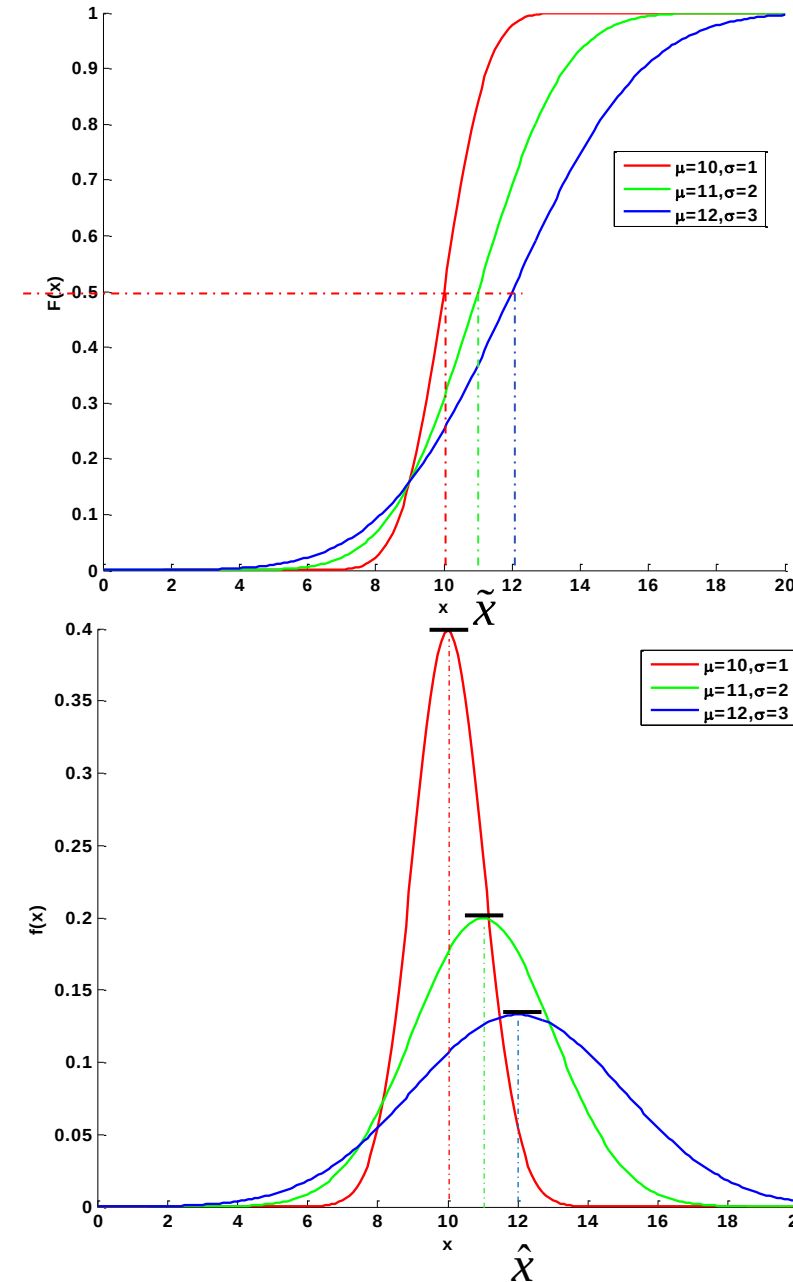
$$\tilde{x} = \mu$$

mode

$$\left. \frac{\partial}{\partial x} f(x|\theta) \right|_{\hat{x}} = 0$$

$$\hat{x} = \mu$$

1/2



$$\frac{\tilde{x}}{10}$$

$$\frac{\tilde{x}}{11}$$

$$\frac{\tilde{x}}{12}$$

$$\frac{\hat{x}}{10}$$

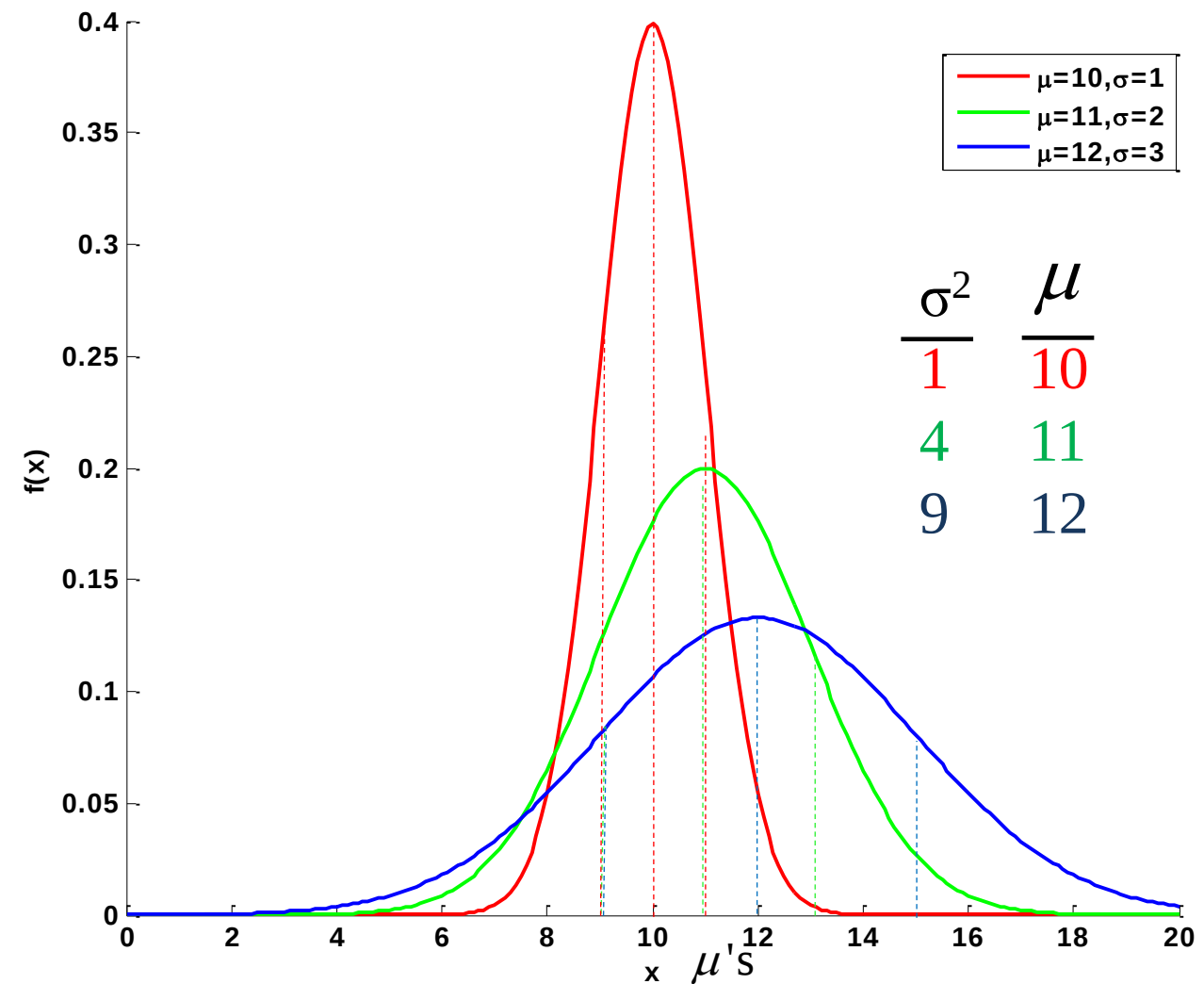
$$\frac{\hat{x}}{11}$$

$$\frac{\hat{x}}{12}$$

Normal PDF

that

$$\begin{aligned}\sigma^2 &= \int_{-\infty}^{\infty} (x - \mu)^2 f(x | \theta) dx \\ &= \int_{-\infty}^{\infty} (x - \mu)^2 \frac{e^{-\frac{(x - \mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}} dx\end{aligned}$$



Normal PDF

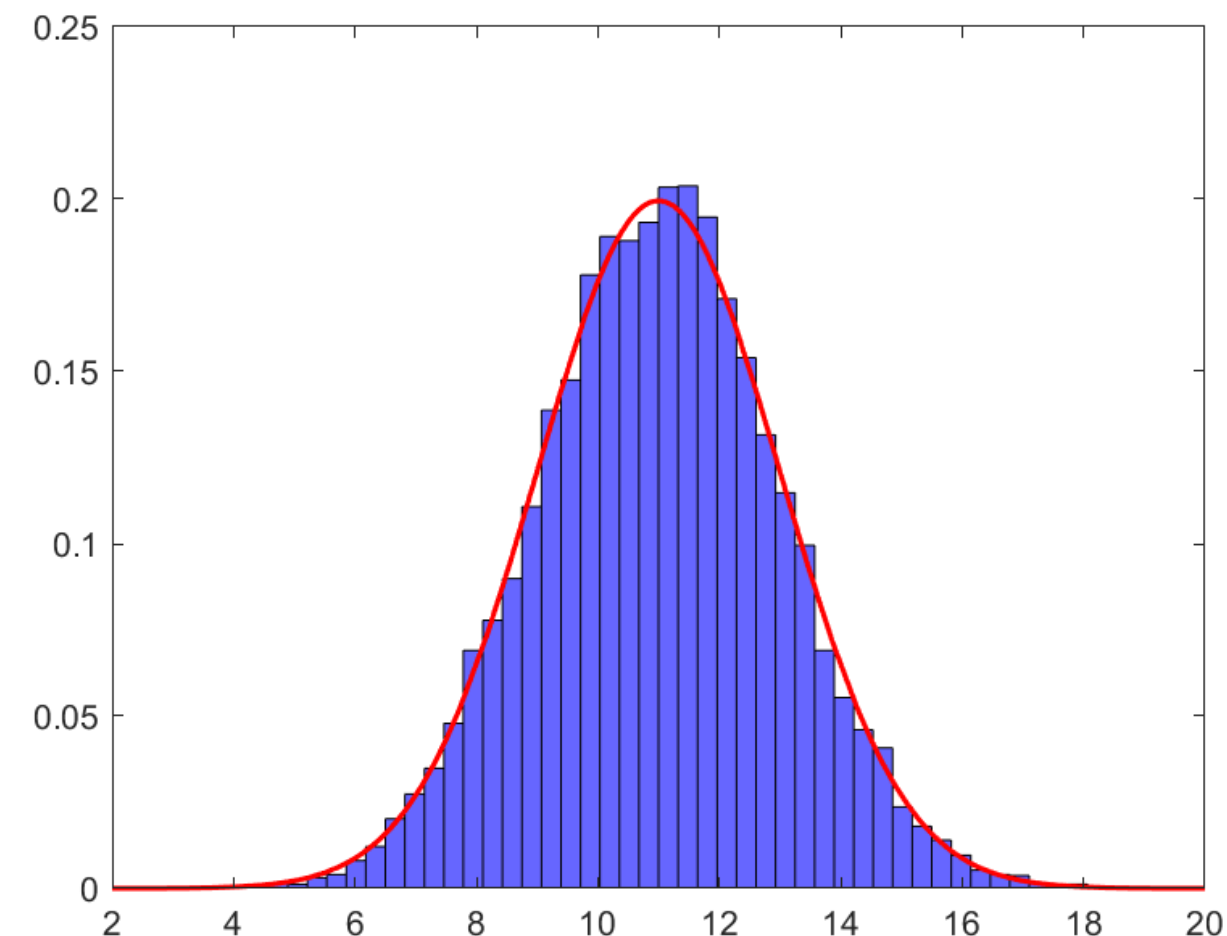
```
rng('default'), mu=11;sigma=2;num=10^4;
x=sigma*randn(num,1)+mu;
figure;
histogram(x,'BinLimits',[2,20],'normalization','pdf','FaceColor','blue')
hold on
fx = @(z) 1/sqrt(2*pi*sigma^2).*exp(-(z-mu).^2/(2*sigma^2));
fplot(fx,[0,22],'r','LineWidth',1.5)
xlim([2,20])
```

	True	Simulated
μ	11	11.0033
σ^2	4	3.9321

Can also find and plot ECDF.

```
themean=mu
thevar=sigma^2
[mean(x),var(x)]
```

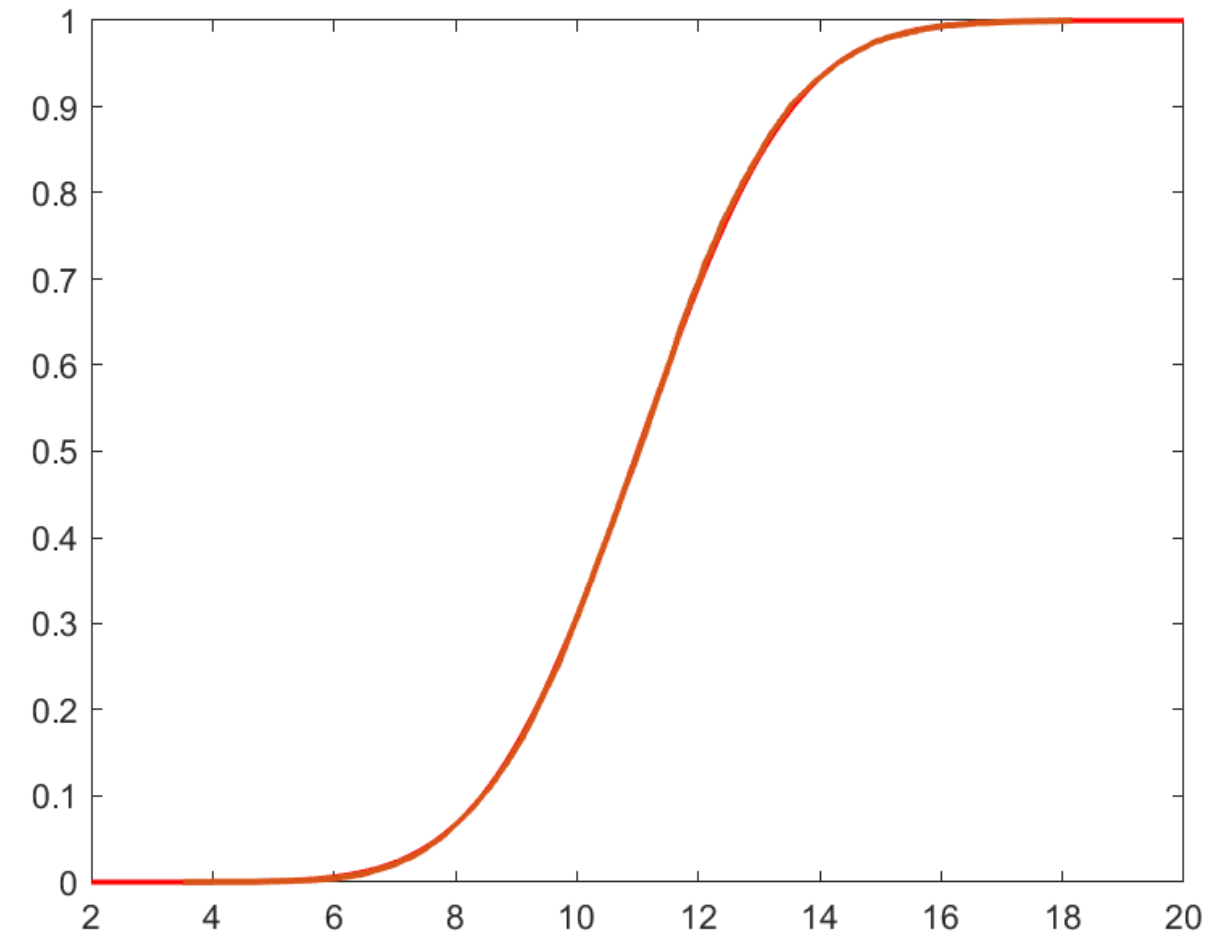
$$f(x|\mu,\sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$



Normal PDF

```
y=cdf('Normal',(2:.01:20),mu,sigma);  
figure;  
plot((2:.01:20),y,'Color','red','LineWidth',1.5)  
[F,xx]=ecdf(x);  
hold on  
stairs(xx,F,'Color','blue','LineWidth',1.5)
```

$$f(x|\mu,\sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$



Inverse Gamma PDF

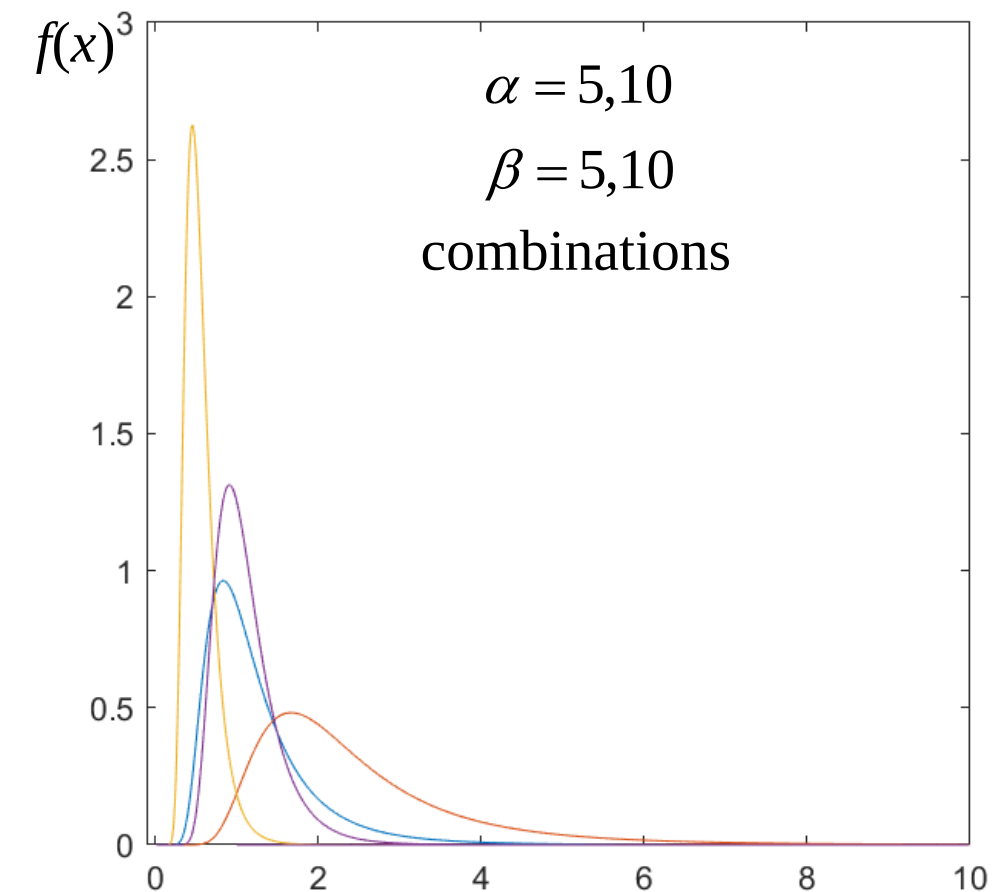
A random variable x has a continuous Inverse Gamma distribution, $x \sim \text{Inv}\Gamma(\alpha, \beta)$ if

$$f(x | \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x},$$

where, $\alpha, \beta > 0$, $0 < x < \infty$.

And can take on many shapes.

$$\Gamma(a) = (a-1)! \text{ for integer } a$$



Inverse Gamma PDF

There is no closed form analytic formula for the CDF $F(x | \alpha, \beta)$.

It can be shown that the mean is

$$\begin{aligned}\mu &= \int_0^\infty x f(x | \theta) dx \\ &= \int_0^\infty x \frac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x} dx \\ &= \frac{\beta}{\alpha - 1}\end{aligned}$$

Inverse Gamma PDF

It can be shown that the variance is

$$\begin{aligned}\sigma^2 &= \int_x (x - \mu)^2 f(x | \theta) dx \\ &= \int_{x=0}^{\infty} (x - \mu)^2 \frac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x} dx \\ &= \frac{\beta^2}{(\alpha - 1)^2 (\alpha - 2)}\end{aligned}$$

Inverse Gamma PDF

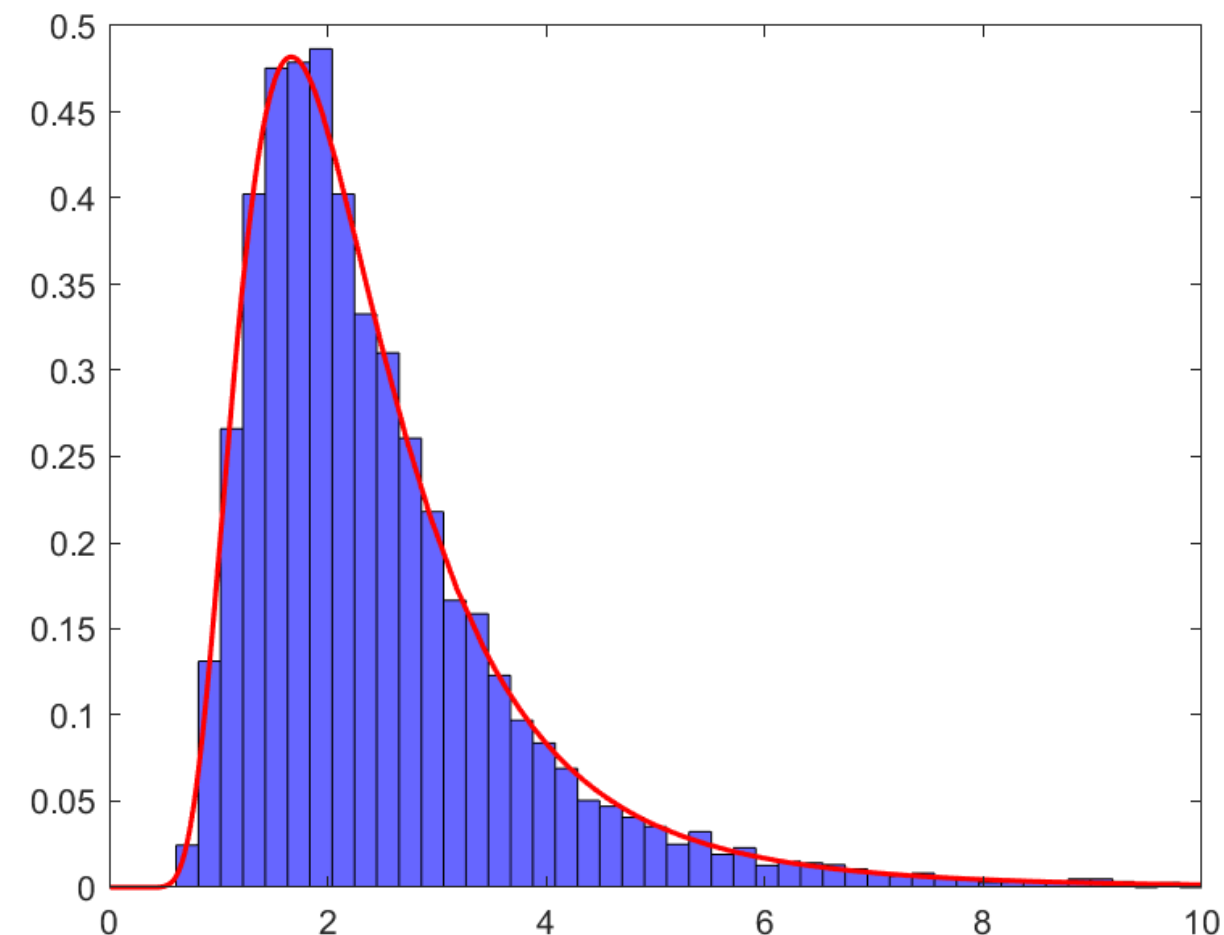
```
rng('default'), a=5; b=10; c=1/b; num=10^4;
x=1./gamrnd(a,c,[num,1]);
figure;
histogram(x,'BinLimits',[0,10],'normalization','pdf','FaceColor','blue')
hold on
fx = @(z) b.^a/factorial(a-1).*z.^(-a-1).*exp(-b./z);
fplot(fx,[0,10],'r','LineWidth',1.5)
xlim([0,10])
```

	True	Simulated
μ	2.5000	2.4877
σ^2	2.0833	1.9690

Can also find and plot ECDF.

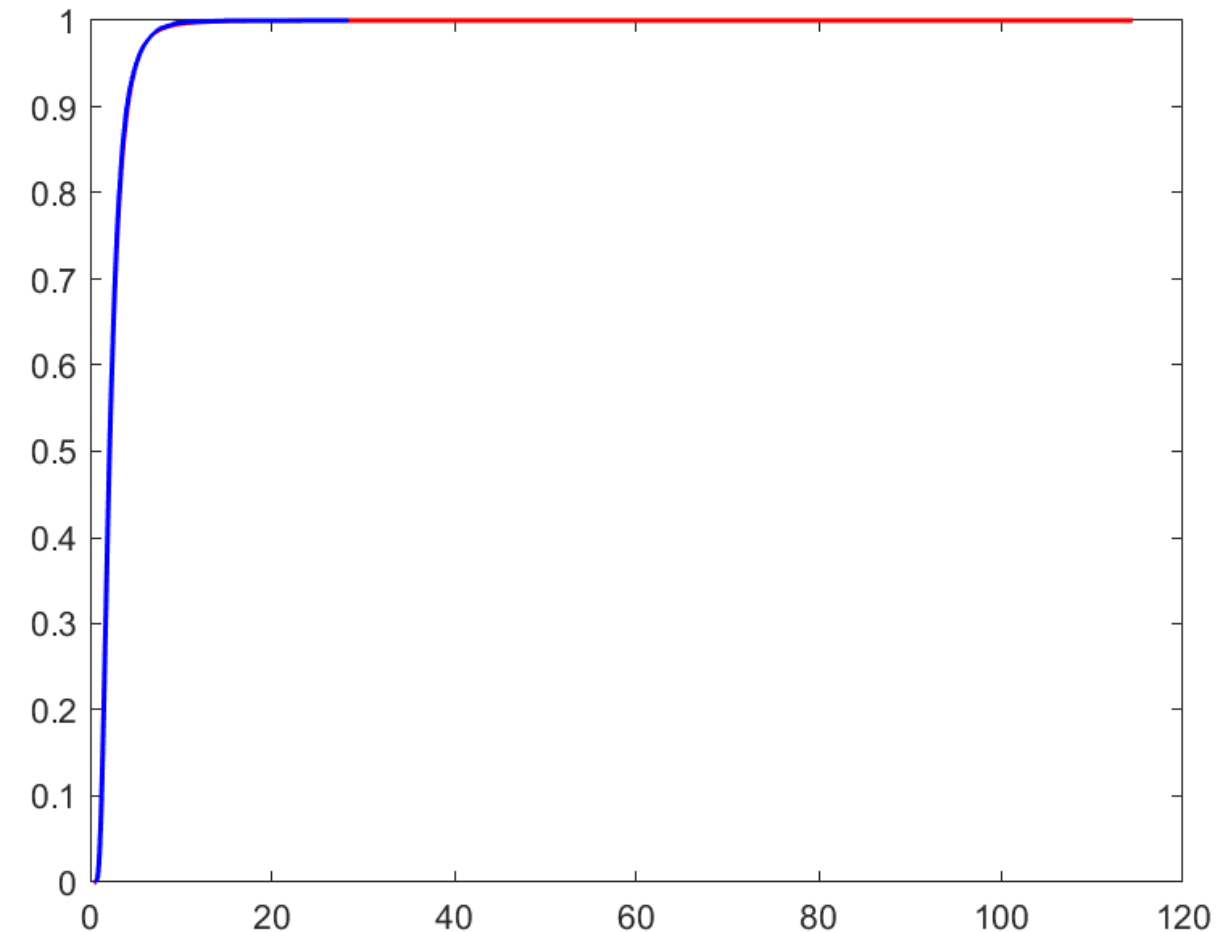
```
themean=b/(a-1)
thevar=(b^2)/((a-1)^2*(a-2))
[mean(x),var(x)]
```

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{-\alpha-1} e^{-\beta/x}$$



Inverse Gamma PDF

```
y=1./gamrnd(a,c,[5*10^6,1]);  
[Fy,yy]=ecdf(y);  
figure;  
stairs(yy,Fy,'Color','red','LineWidth',1.5)  
[F,xx]=ecdf(x);  
hold on  
stairs(xx,F,'Color','blue','LineWidth',1.5)
```



Questions?

Homework 3

$$f(x|\alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

1. Use Matlab to generate $n=1000$ observations from the Beta PDF with $\alpha=3$ and $\beta=7$. Change α, β to observe shapes.
 - a. Calculate the sample mean.
 - b. Calculate the sample variance.
 - c. Calculate the median.
 - d. Make a histogram.
 - e. Make an empirical CDF and calculate the 95th percentile

Homework 3

$$f(x | \mu, \sigma^2) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma^2}}$$

2. Use Matlab to generate $n=1000$ observations from the normal PDF with $\mu=100$ and $\sigma^2=5$.
 - a. Calculate the sample mean.
 - b. Calculate the sample variance.
 - c. Calculate the median.
 - d. Make a histogram.
 - e. Make an empirical CDF and calculate the 95th percentile.

Homework 3

$$f(x | \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$$

3. Assume y has a gamma distribution with parameters with $\alpha=3$, $\beta=7$, and $c=1/\beta$. Change α, β to observe shapes.
- a. Generate $n=1000$ random y observations.
Calculate the sample mean and variance.
Make a histogram. Compare to the Gamma PDF.
 - b*. Using pencil and paper transform to x as $x=1/y$.
 - c. Take $x=1/y$ for each x random observation in part a.
 - d. Repeat a. but now for y 's and compare to inverse gamma PDF.

*For students in MSSC 5790