

# The Distribution of Magnitude and Complex Voxel Values in MRI

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Waisman Laboratory  
for Brain Imaging  
and Behavior

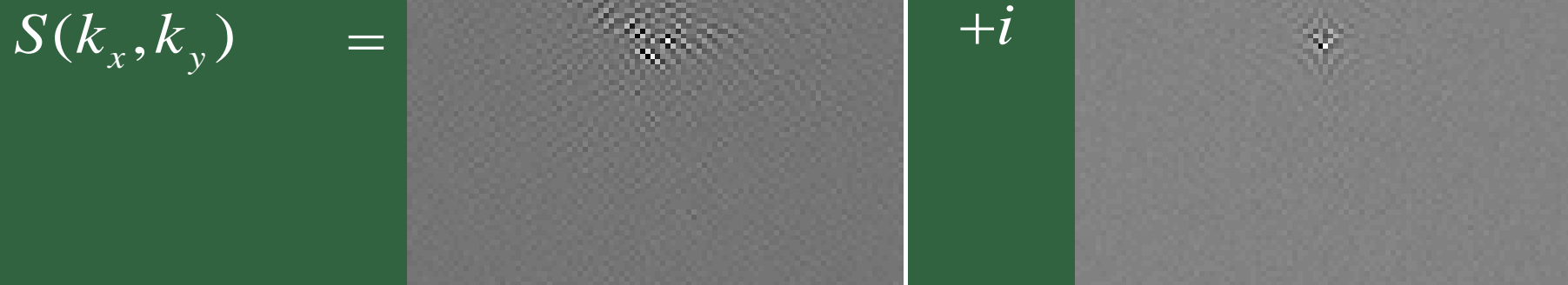
# OUTLINE

1. Image Reconstruction
2. Statistics-Ricean & Normal
3. Estimation-Ricean & Normal
4. Estimation-Bivariate Normal
5. Discussion

**Reconstruction:**

Ideally measure complex-valued FT of the object.

$$S(k_x, k_y) = S_R(k_x, k_y) + i S_I(k_x, k_y)$$



Complex: 96×96

Real: 96×96

Imaginary: 96×96

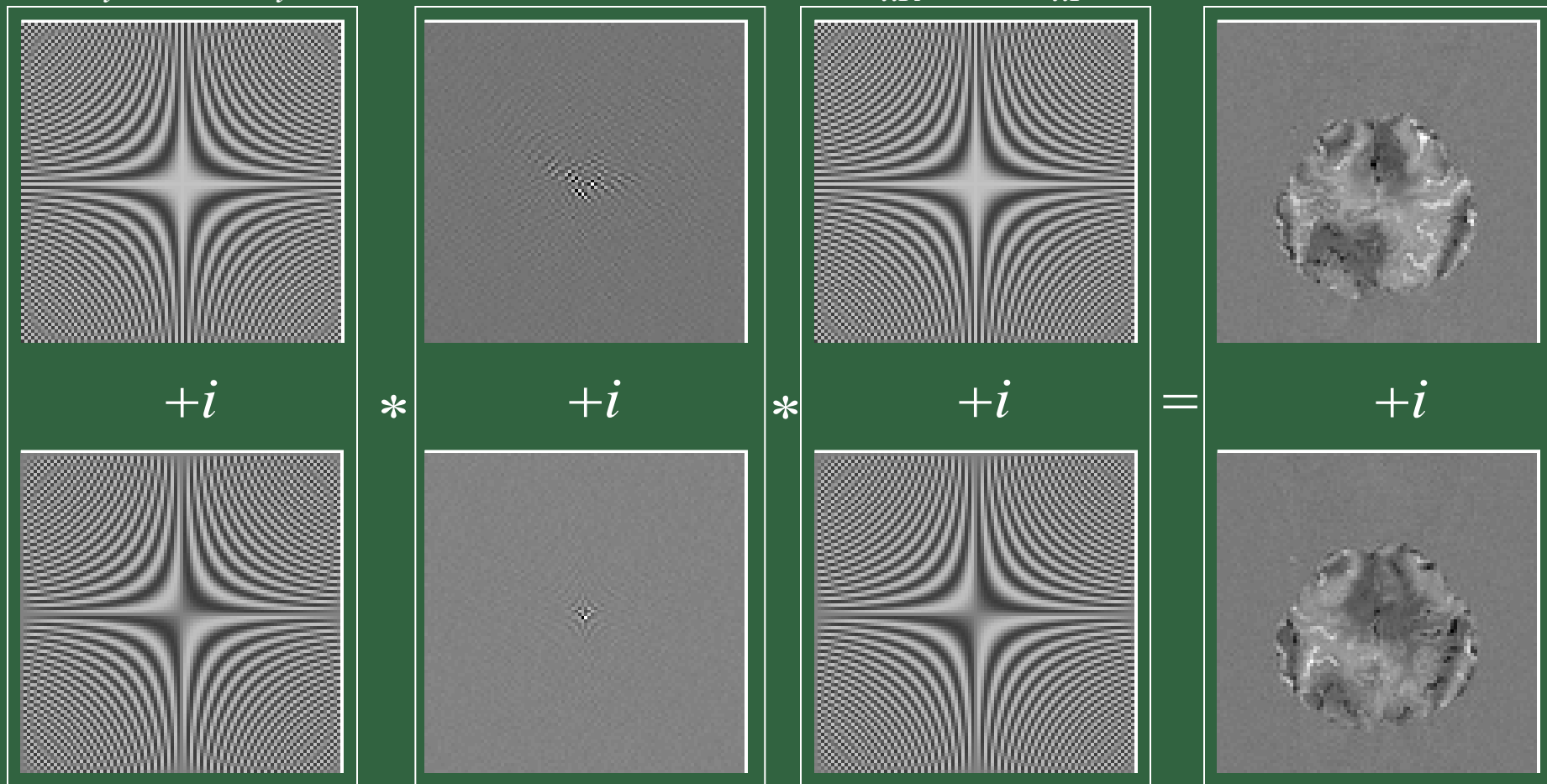
Actual data!

$p = 9216$   
# of voxels

# Reconstruction:

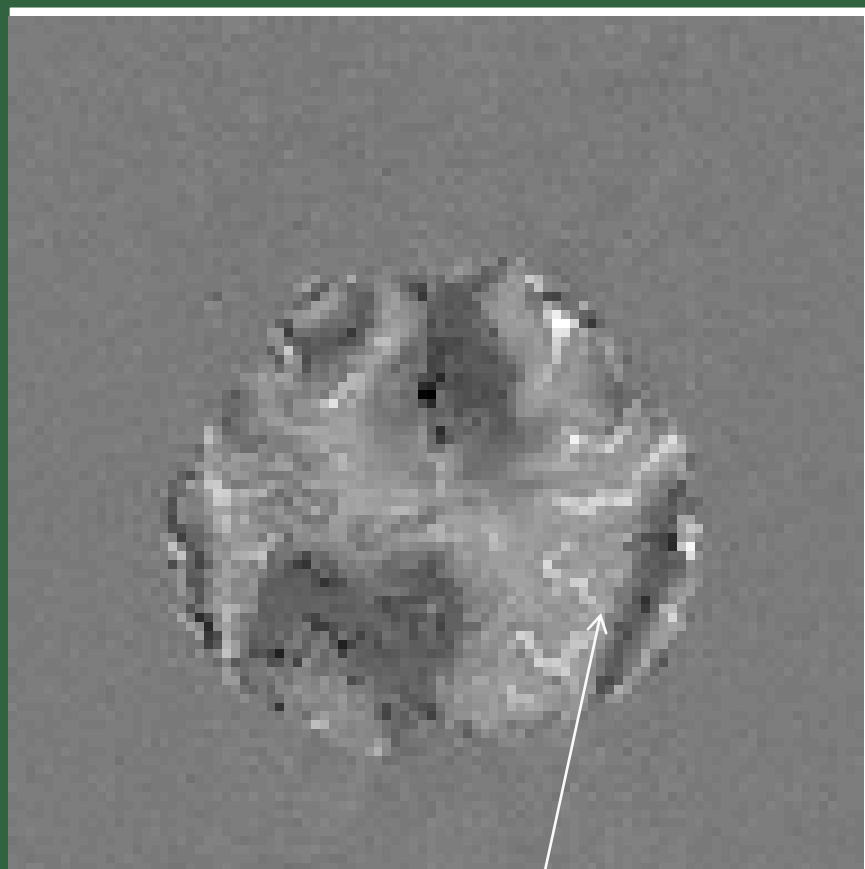
By complex-valued inverse FT of the object.

$$(\Omega_{yR} + i\Omega_{yI}) * (S_R + iS_I) * (\Omega_{xR} + i\Omega_{xI})^T = (Y_R + iY_I)$$



## Reconstruction:

Due to imperfect reconstruction (noise,  $T_2^*$ ,  $\Delta B$ , ...), image is complex-valued,  $Y_C(x, y) = Y_R(x, y) + iY_I(x, y)$ .



Real Image

$y_R$



Imaginary Image

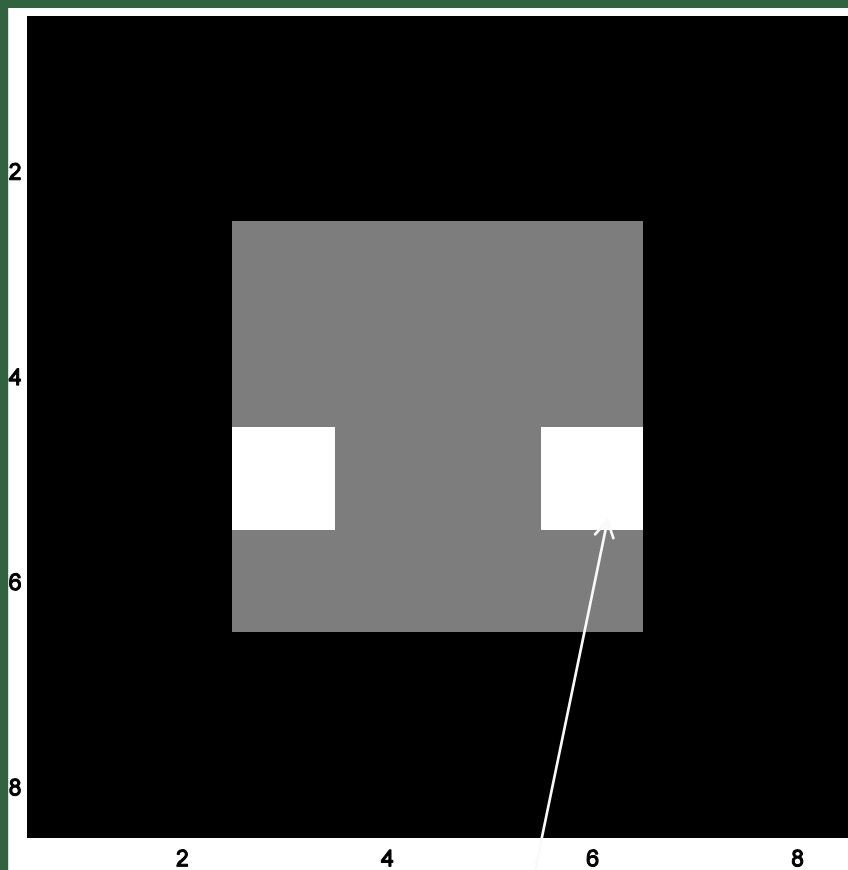
$y_I$

given voxel

## Reconstruction:

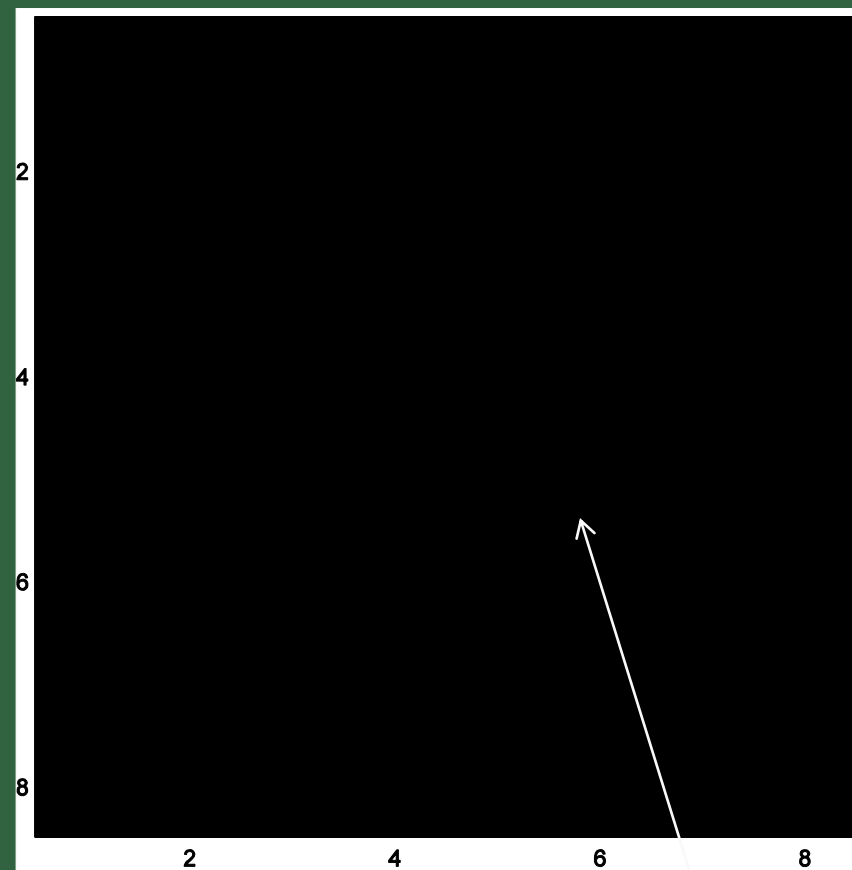
Toy Example 8×8, image is complex-valued,

$$Y_C(x, y) = Y_R(x, y) + iY_I(x, y).$$



Real Image

$y_R$



Imaginary Image

$y_I$

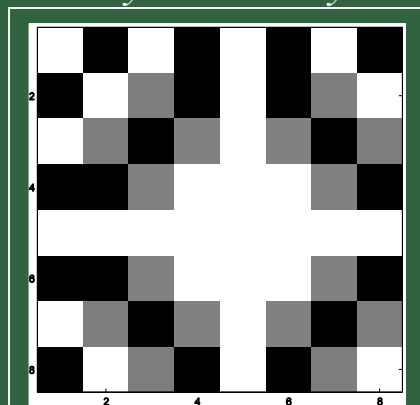
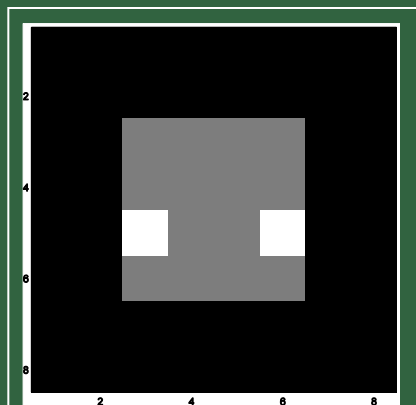
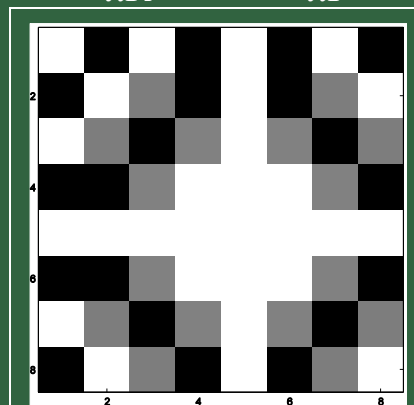
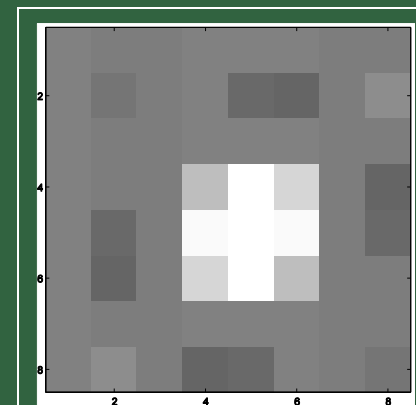
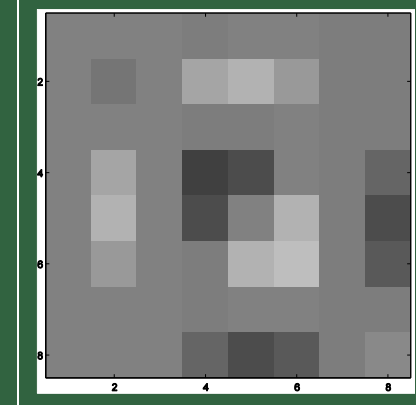
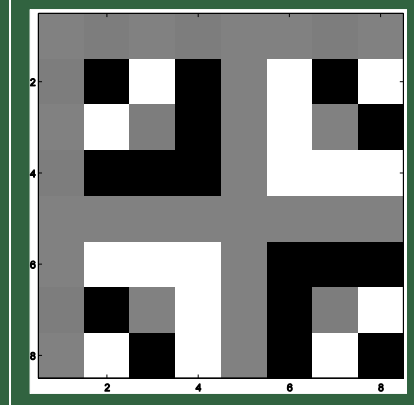
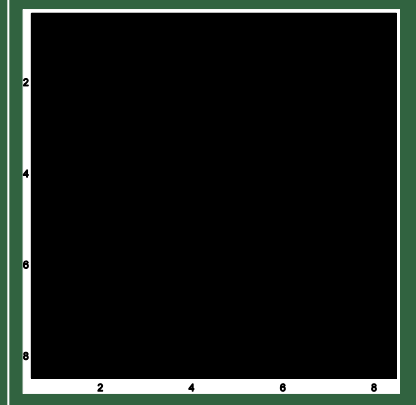
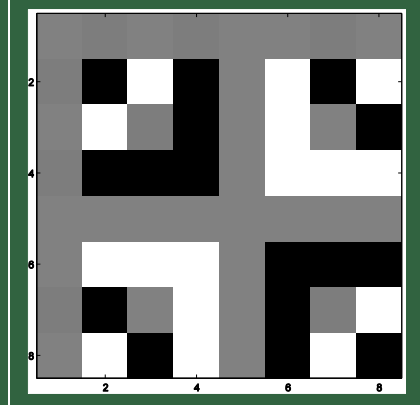
given voxel

# Reconstruction:

By complex-valued forward FT of the object.

$$\Omega \bar{\Omega} = I$$

$$(\bar{\Omega}_{yR} + i\bar{\Omega}_{yI}) * (Y_R + iY_I) * (\bar{\Omega}_{xR} + i\bar{\Omega}_{xI})^T = (S_R + iS_I)$$


 $+i$ 

 $+i$ 

 $+i$ 

 $+i$ 


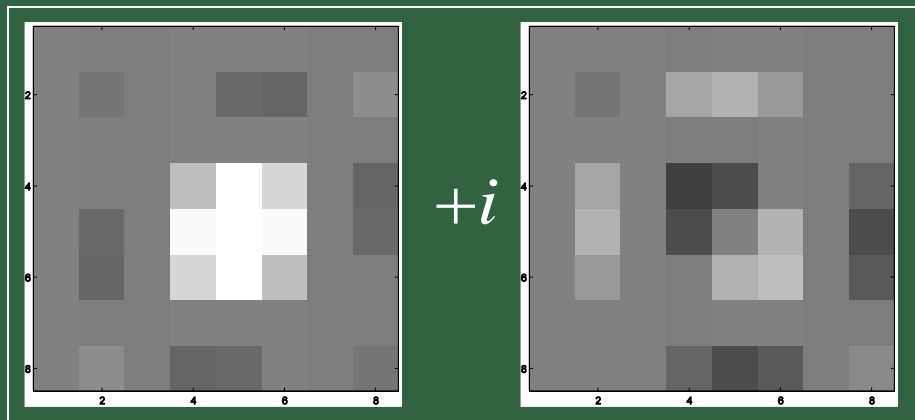
# Reconstruction:

On Cartesian grid

$S_R$

$+i$

$S_I$

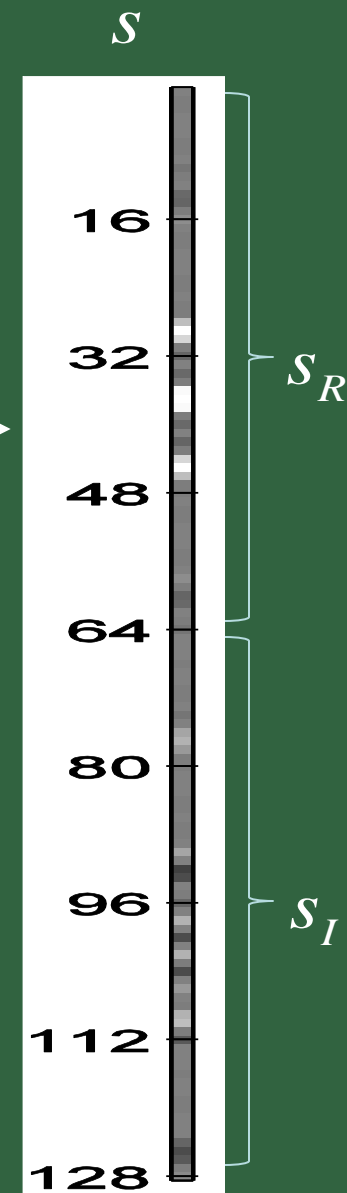
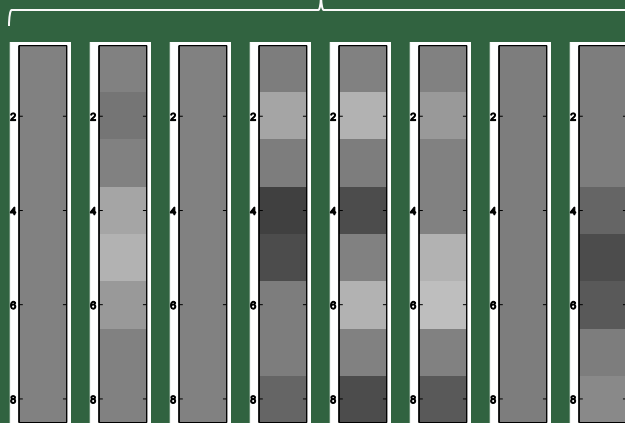
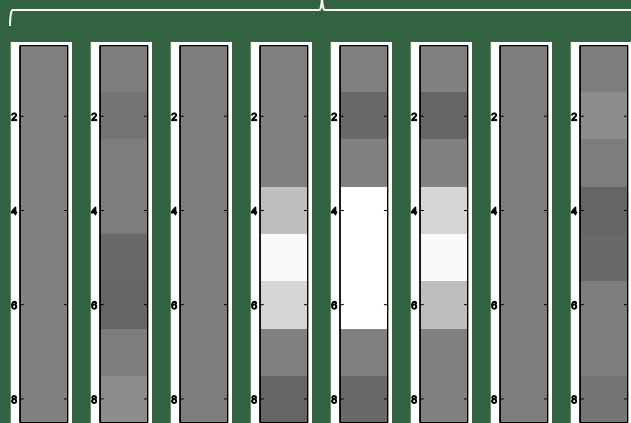


$+i$

stack rows of  $S_R$  on rows of  $S_I$

take rows of  $S_R$

take rows of  $S_I$





# Reconstruction:

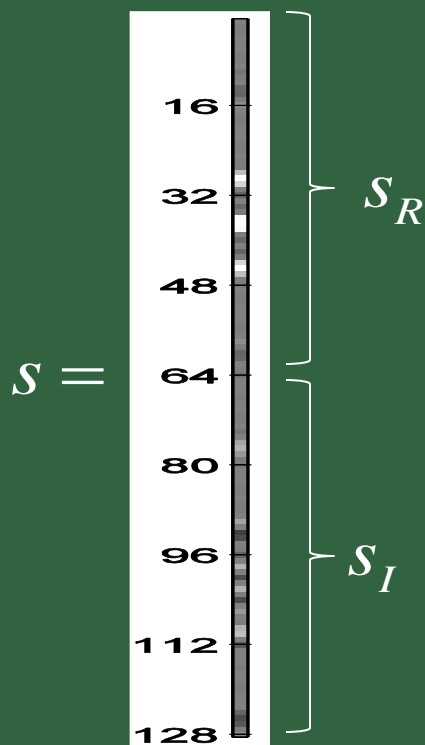
$$S = \begin{pmatrix} S_R \\ S_I \end{pmatrix}$$

$2p \times 1$

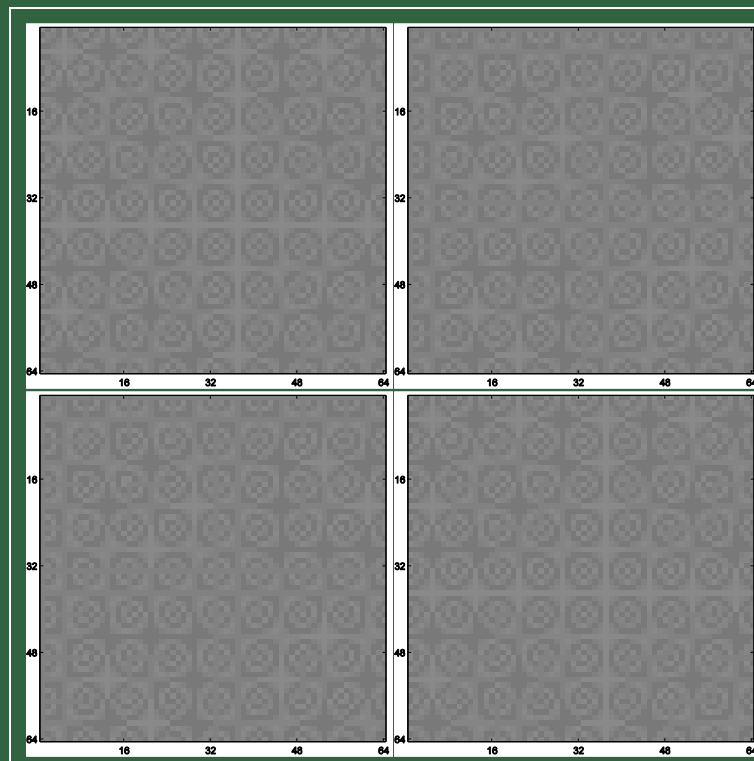
reconstruction matrix

$$\Omega = \begin{bmatrix} \Omega_R & -\Omega_I \\ \Omega_I & \Omega_R \end{bmatrix}$$

k-space vector



$$\Omega =$$



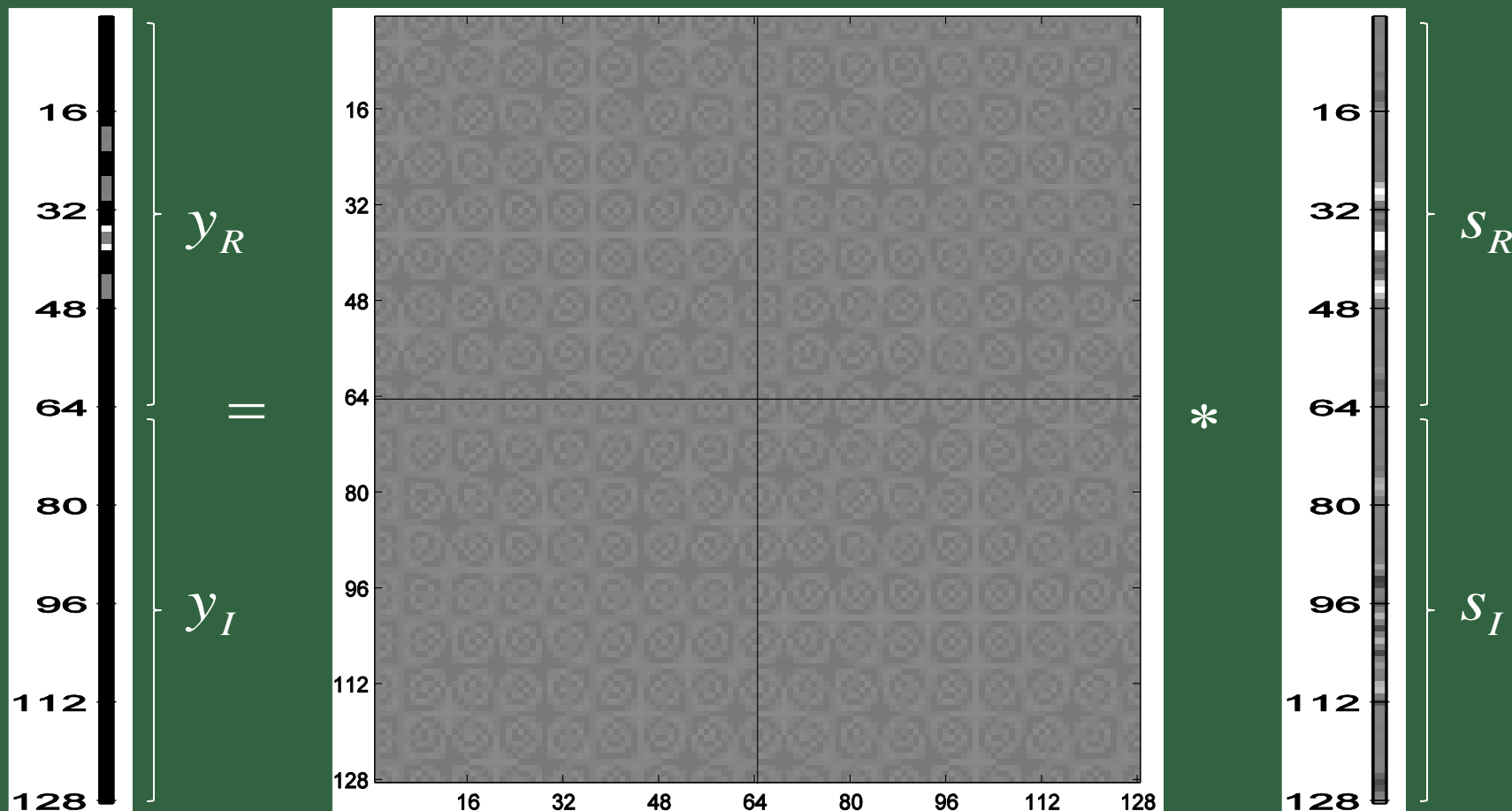
$$\Omega_R = [(\Omega_{yR} \otimes \Omega_{xR}) - (\Omega_{yI} \otimes \Omega_{xI})]$$

$$\Omega_I = [(\Omega_{yR} \otimes \Omega_{xI}) + (\Omega_{yI} \otimes \Omega_{xR})]$$

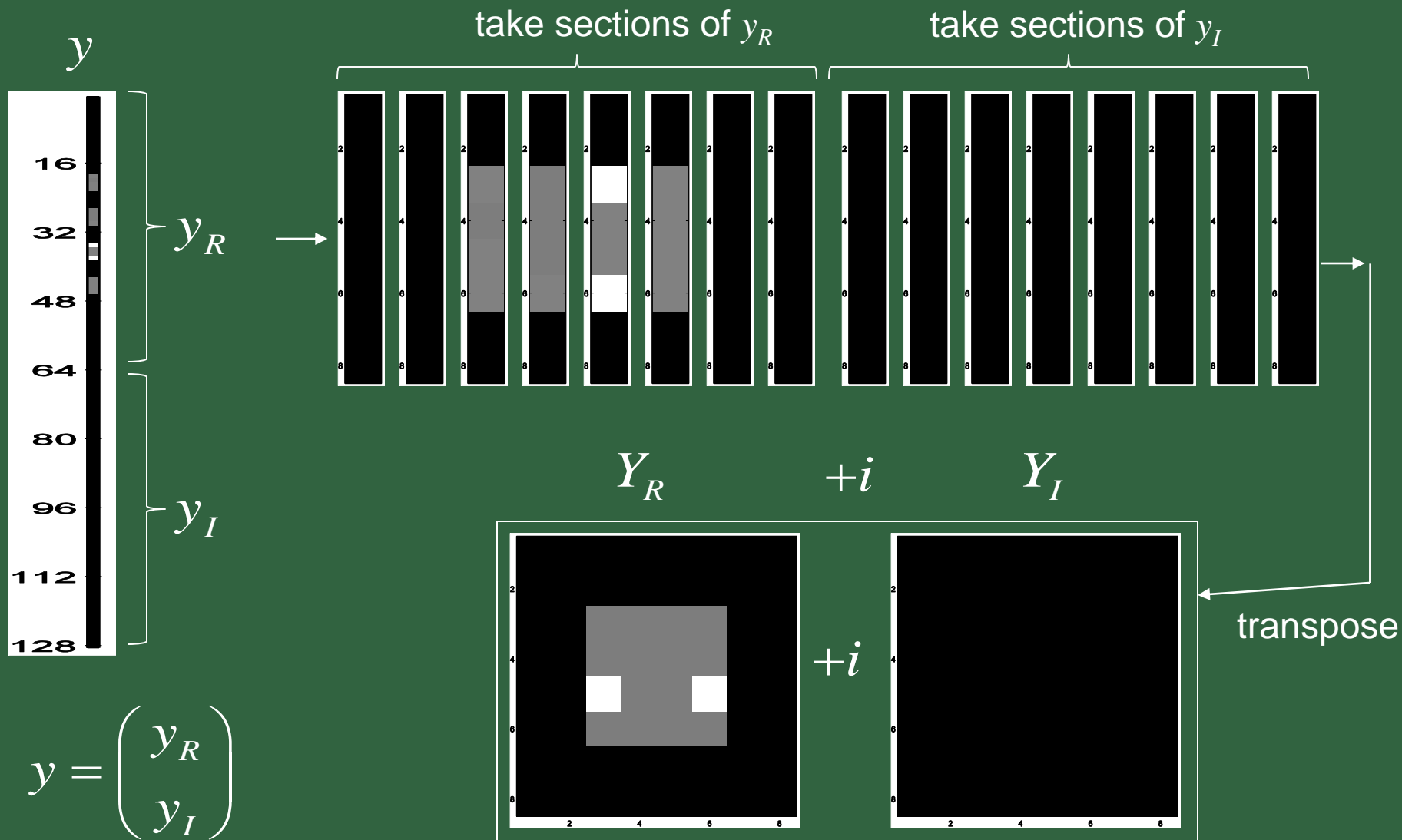
# Reconstruction:

Inverse FT reconstruction can be equivalently described as:

$$y = \Omega * S$$



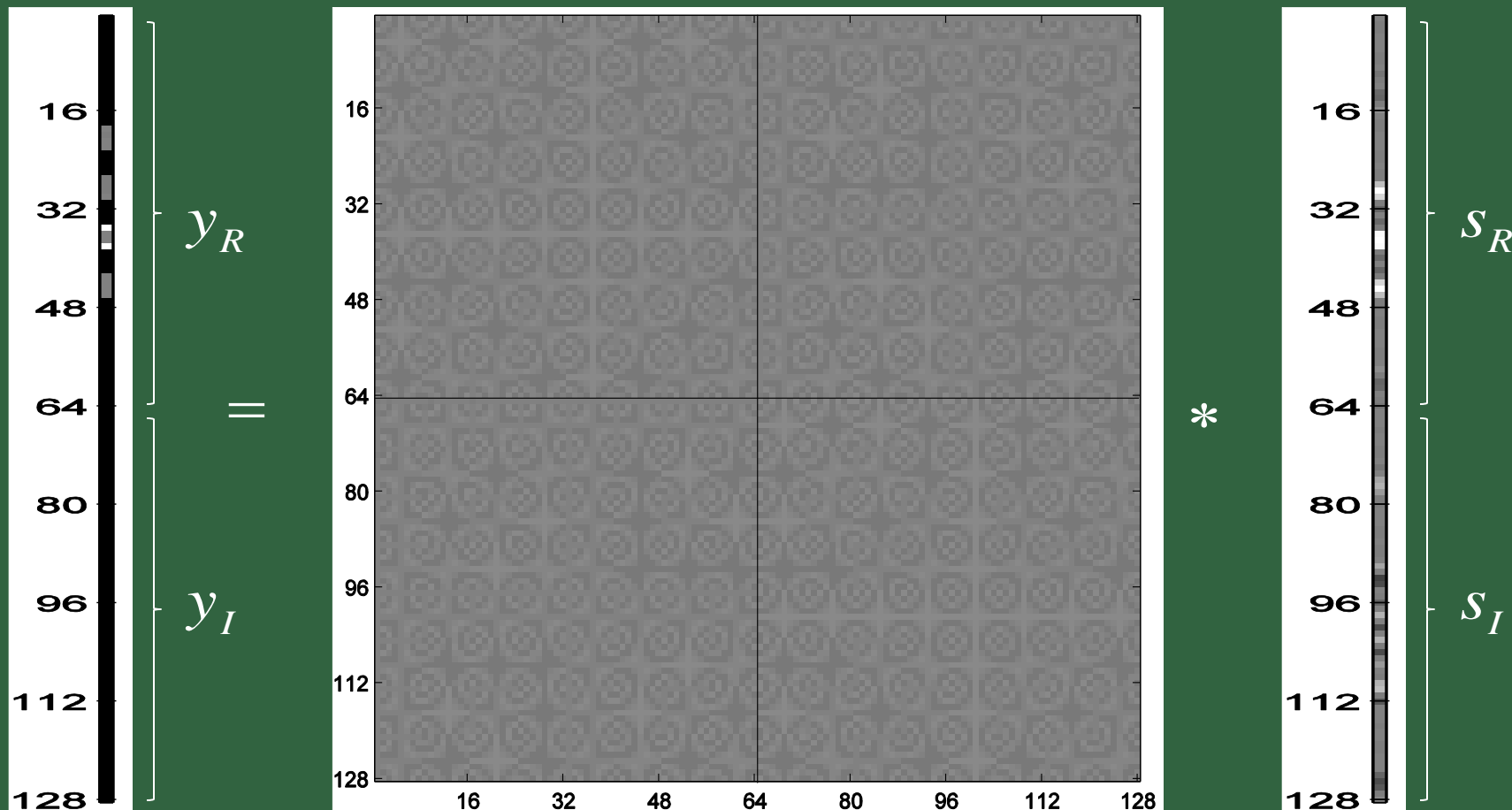
# Reconstruction:



# Reconstruction:

Inverse FT reconstruction can be performed as:

$$y = \Omega * S$$



## Statistics: Expectation and Covariance.

If  $E(s) = s_0$ , then for  $y = \Omega s$ ,  $E(y) = E(\Omega s) = \Omega s_0$ .

If  $\text{cov}(s) = \Gamma$ , then for  $y = \Omega s$ ,  $\text{cov}(y) = \text{cov}(\Omega s) = \Omega \Gamma \Omega'$ .

This means that with  $\Gamma = \sigma_k^2 I$ ,

and because  $\Omega \Omega' = \sigma^2 I$  where  $\sigma^2 = (\sigma_k^2 / p^2)$

$$\text{cov}(y) = \sigma^2 I_{2p \times 2p}.$$

## Statistics: Expectation and Covariance.

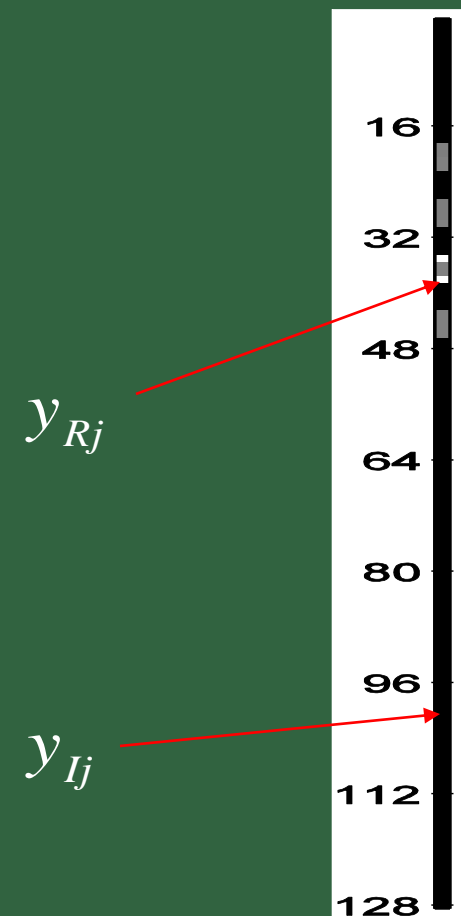
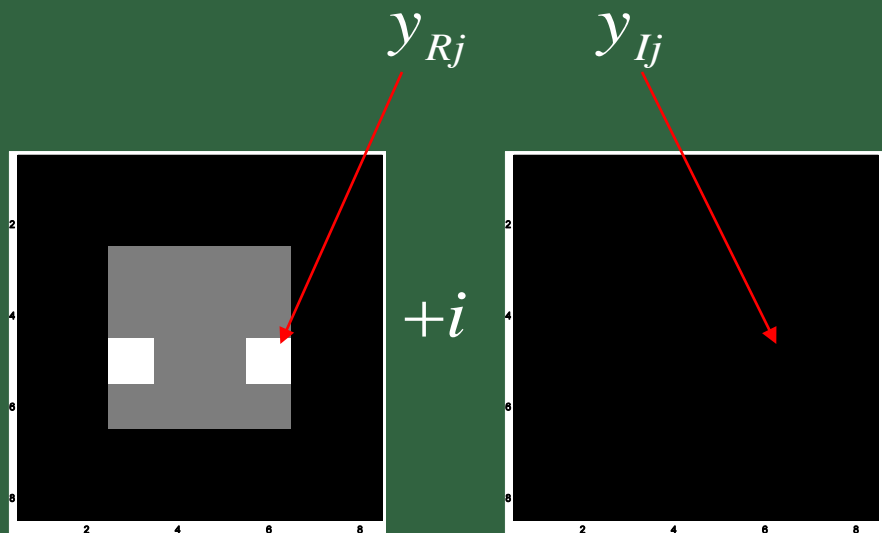
When we use normal distribution from thermal noise

$$s = s_0 + \varepsilon, \quad \varepsilon \sim N(0, \sigma_k^2 I)$$

$$s \sim N(s_0, \sigma_k^2 I), \quad \text{then } y \sim N(\Omega s_0, \sigma^2 I).$$

$2p \times 1$        $2p \times 1$        $2p \times 2p$        $2p \times 1$        $2p \times 1$        $2p \times 2p$

This means that if we choose a voxel, say  $j$



## Statistics: Expectation and Covariance.

from  $y \sim N(\Omega s_0, \sigma^2 I)$ , the distribution of  $y_{Rj}$  and  $y_{Ij}$  is

$$\begin{pmatrix} y_{Rj} \\ y_{Ij} \end{pmatrix} \sim N \left( \begin{pmatrix} \mu_{Rj} \\ \mu_{Ij} \end{pmatrix}, \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix} \right) \quad \text{where} \quad \begin{aligned} \mu_{Rj} &= \omega_j s_0 && \swarrow \text{j}^{\text{th}} \text{ row of } \Omega \\ \mu_{Ij} &= \omega_{p+j} s_0 && \nwarrow \text{(p+j)}^{\text{th}} \text{ row of } \Omega \end{aligned}$$

$$y_{Cj} = y_{Rj} + iy_{Ij}$$

the pdf is

$$p(y_{Rj}, y_{Ij}) = \frac{1}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ (y_{Rj} - \mu_{Rj})^2 + (y_{Ij} - \mu_{Ij})^2 \right] \right\}$$

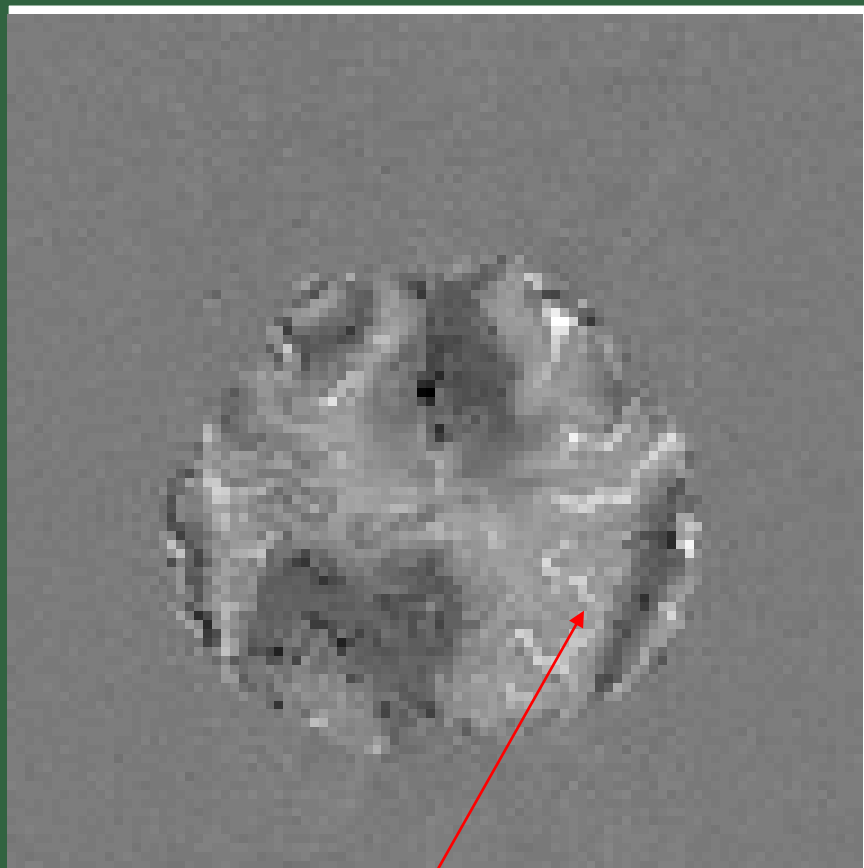
product of two normal pdfs

$$\text{with phase coupled means} \quad \mu_{Rj} = \rho_j \cos \theta_j$$

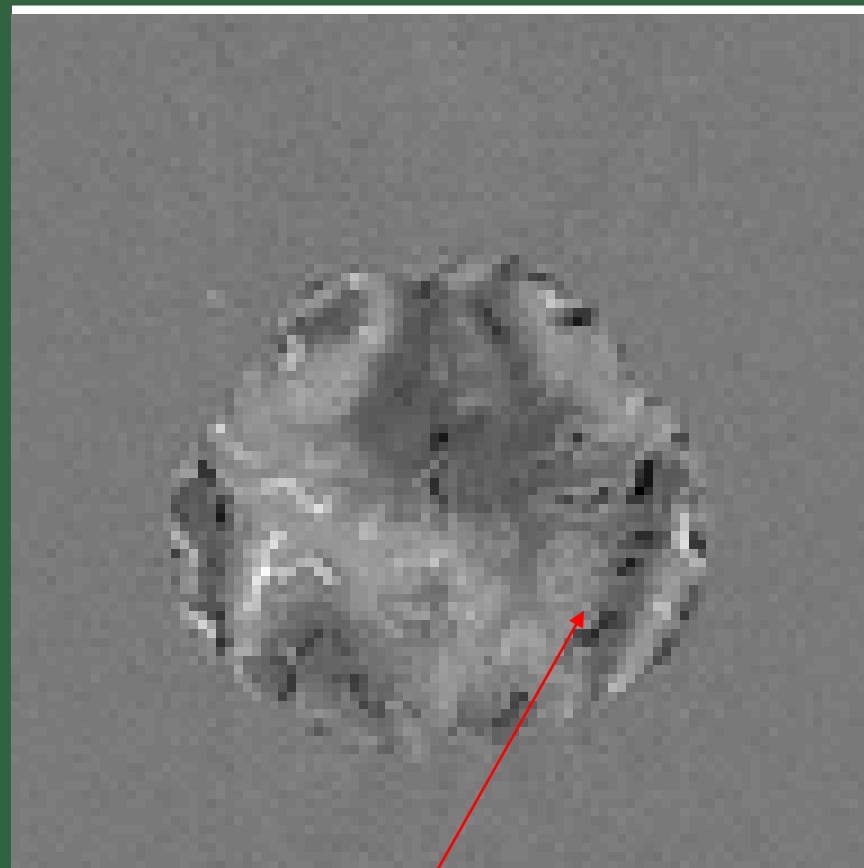
$$\mu_{Ij} = \rho_j \sin \theta_j$$

# Statistics:

## Real Image

voxel  $j$  $y_{Rj}$ 

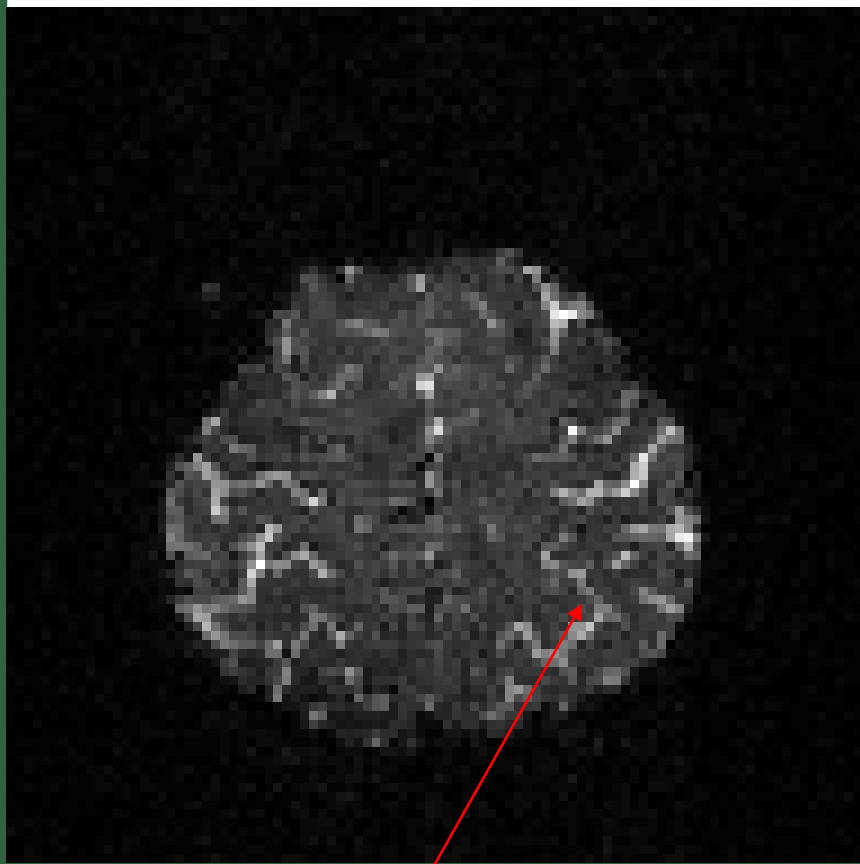
## Imaginary Image

 $y_{Ij}$



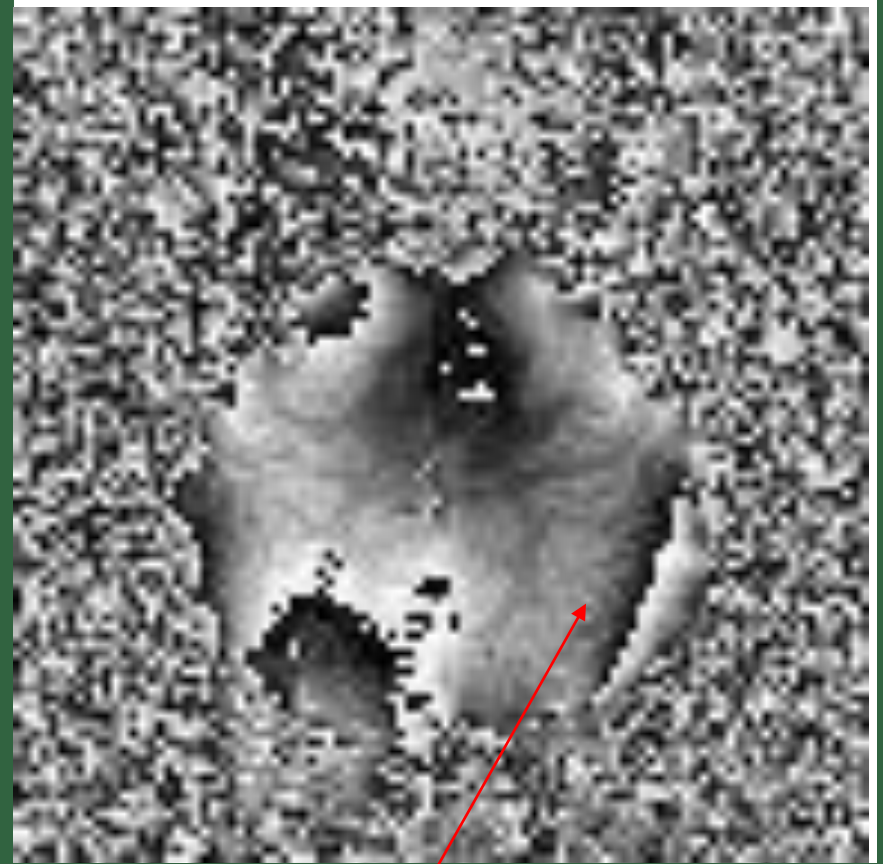
## Statistics:

## Magnitude Image

voxel  $j$ 

$$m_j = \sqrt{y_{Rj}^2 + y_{Ij}^2}$$

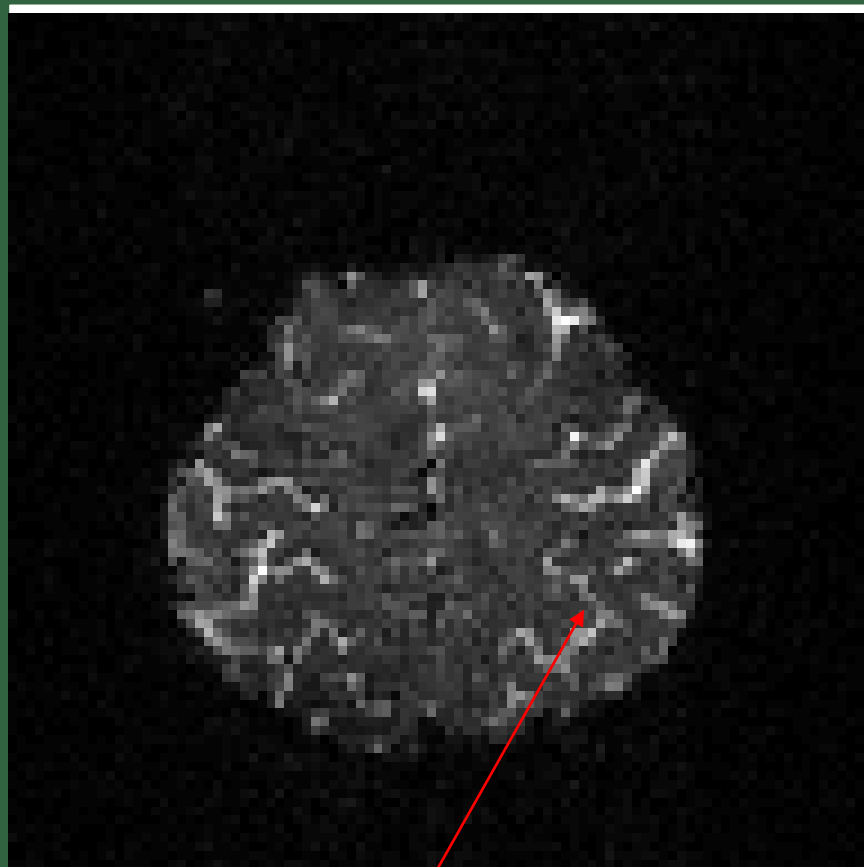
## Phase Image



$$\varphi_j = \tan^{-1}(y_{Ij} / y_{Rj})$$

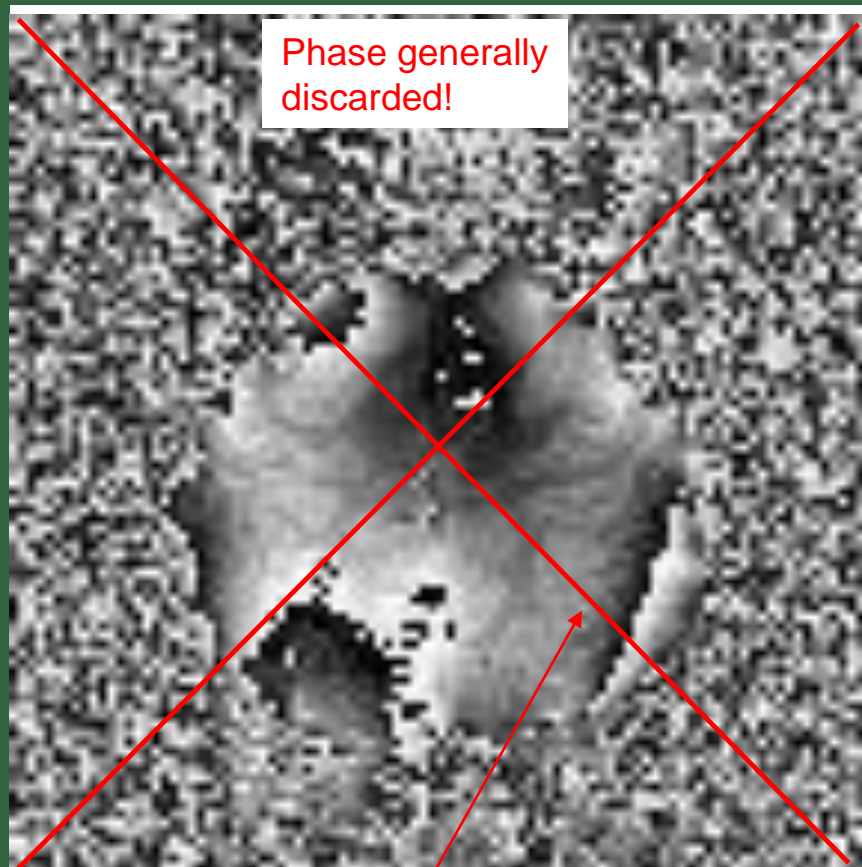
## Statistics:

## Magnitude Image

voxel  $j$ 

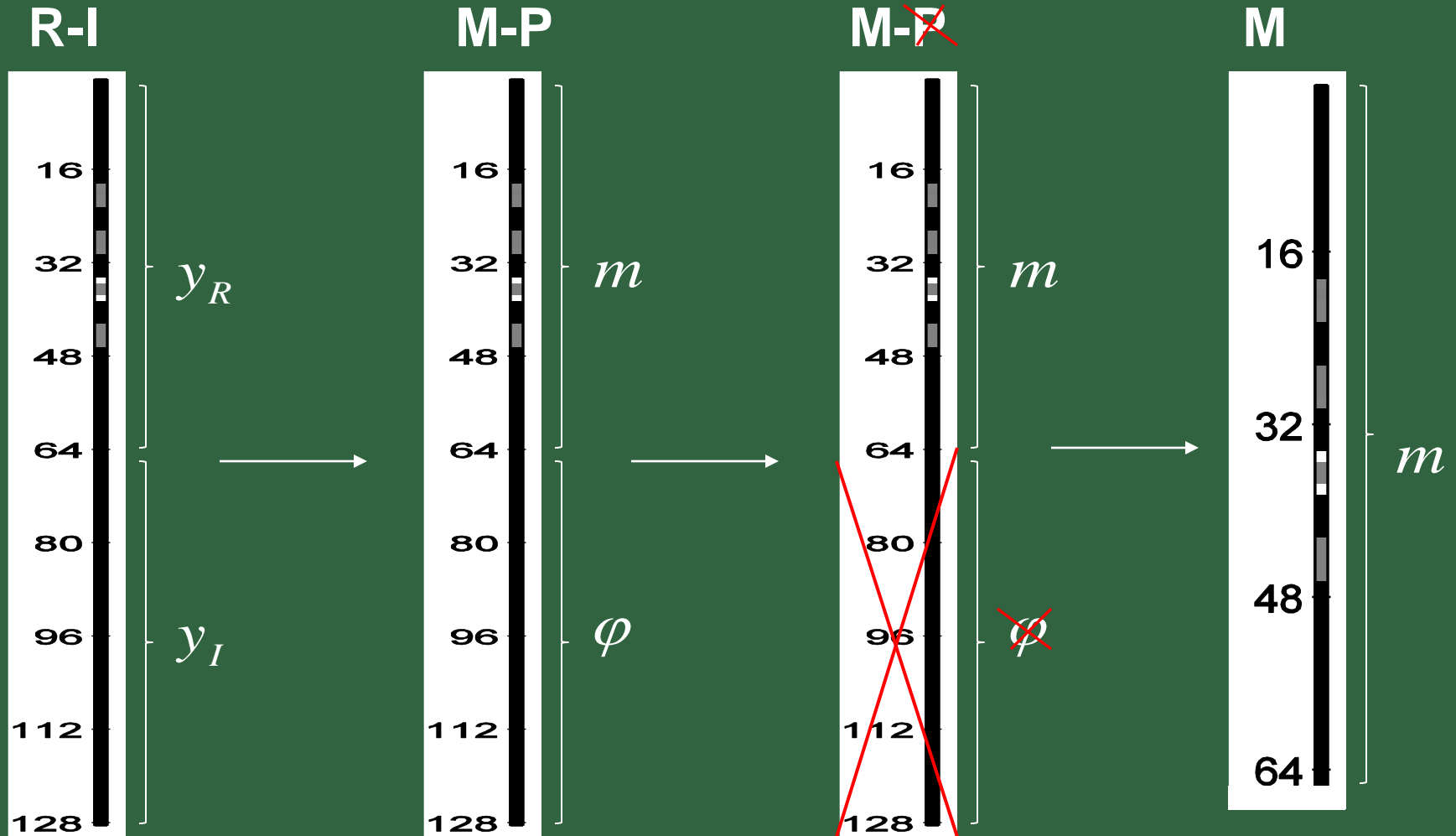
$$m_j = \sqrt{y_{Rj}^2 + y_{Ij}^2}$$

## Phase Image

Phase generally  
discarded!

$$\varphi_j = \tan^{-1}(y_{Ij} / y_{Rj})$$

# Statistics:



$$m_j = \sqrt{y_{Rj}^2 + y_{Ij}^2}$$

## Statistics:

Get  $p(m_j)$  from  $p(y_{Rj}, y_{Ij})$ .

$$\mu_{Rj} = \rho_j \sin \theta_j$$

$$\mu_{Ij} = \rho_j \cos \theta_j$$

Convert from  $y_{Rj}, y_{Ij}$  to  $m_j, \varphi_j$ .

$$p(y_{Rj}, y_{Ij}) = \frac{1}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ (y_{Rj} - \rho_j \cos \theta_j)^2 + (y_{Ij} - \rho_j \sin \theta_j)^2 \right] \right\}$$

$$p(m_j, \varphi_j) = \frac{m_j}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ m_j^2 + \rho_j^2 - 2m_j\rho_j \cos(\varphi_j - \theta_j) \right] \right\}$$

$$p(m_j) = \frac{m_j}{\sigma^2} \exp \left\{ -\frac{m_j^2 + \rho_j^2}{2\sigma^2} \right\} \underbrace{I_0 \left( \frac{\rho_j m_j}{\sigma^2} \right)}$$

zeroth order modified Bessel function of first kind

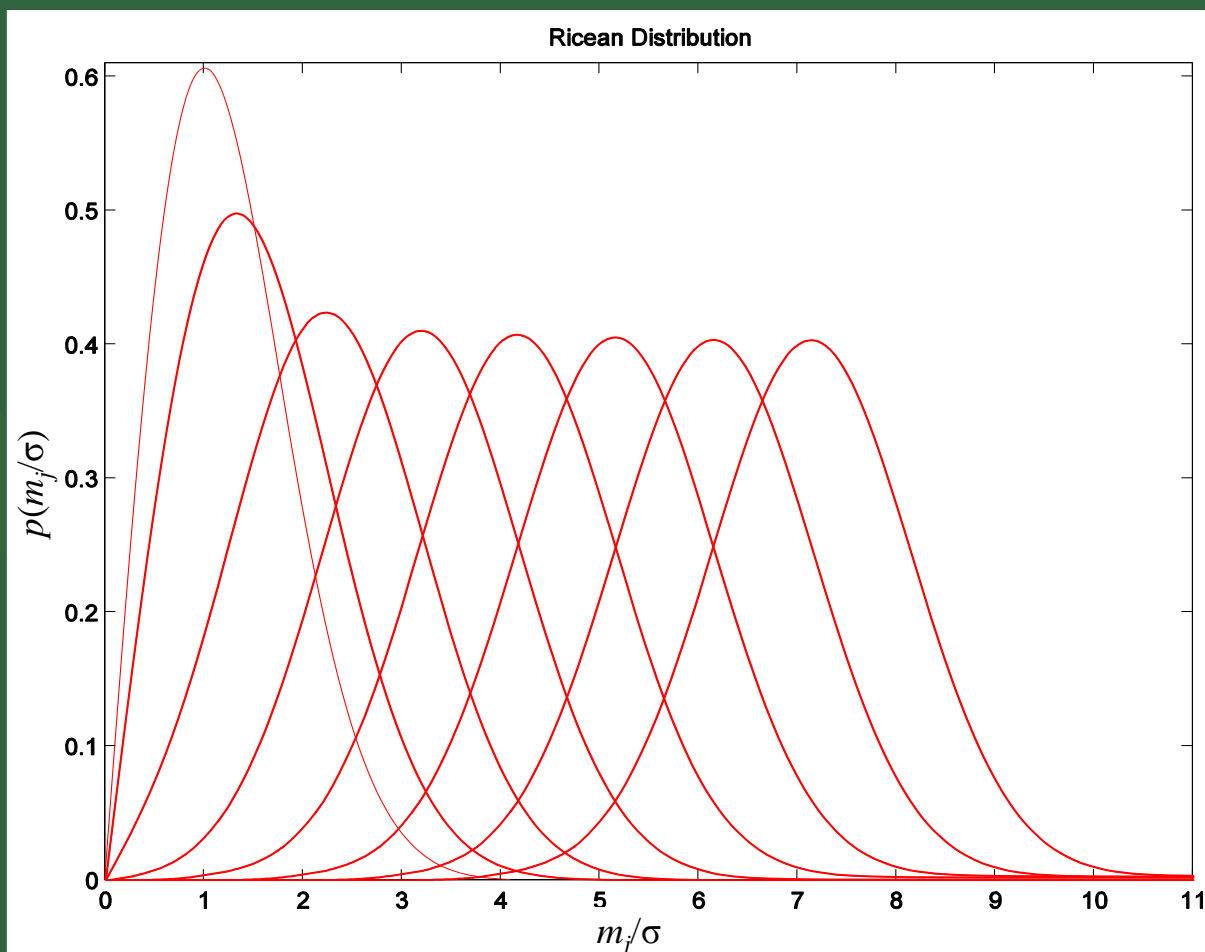
$$\frac{1}{2\pi} \int_{\varphi_j = -\pi}^{\pi} e^{\frac{\rho_j m_j}{\sigma^2} \cos(\varphi_j - \theta_j)} d\varphi_j$$

Rice, S.O., Bell Syst. Tech. 23:282, 1944.  
 Gudbjartsson, Patz. MRM 34:910-914, 1995.  
 Rowe and Logan: NIMG, 23:1078-1092, 2004.

# Statistics:

$$p(m_j) = \frac{m_j}{\sigma^2} \exp\left\{-\frac{m_j^2 + \rho_j^2}{2\sigma^2}\right\} I_0\left(\frac{\rho_j m_j}{\sigma^2}\right)$$

$$SNR = \frac{\rho_j}{\sigma^2}$$



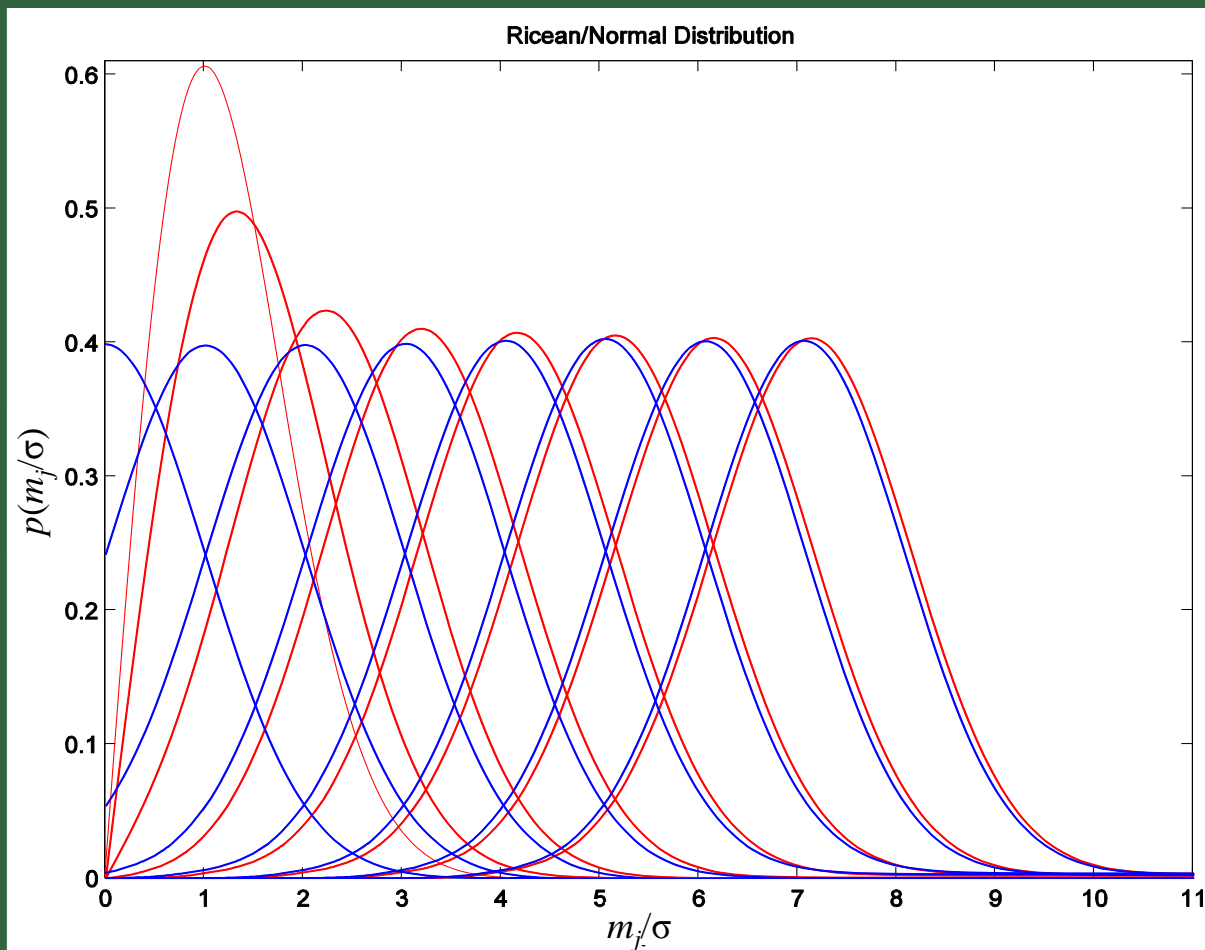
The magnitude, does not have a normal distribution!

Ricean Distribution!

## Statistics:

$$p(m_j) = \frac{m_j}{\sigma^2} \exp\left\{-\frac{m_j^2 + \rho_j^2}{2\sigma^2}\right\} I_0\left(\frac{\rho_j m_j}{\sigma^2}\right)$$

$$SNR = \frac{\rho_j}{\sigma^2}$$



The magnitude, does not have a normal distribution!

Ricean Distribution!

Ricean  $\rightarrow$  Normal  
as the SNR  $\uparrow$

## Statistics:

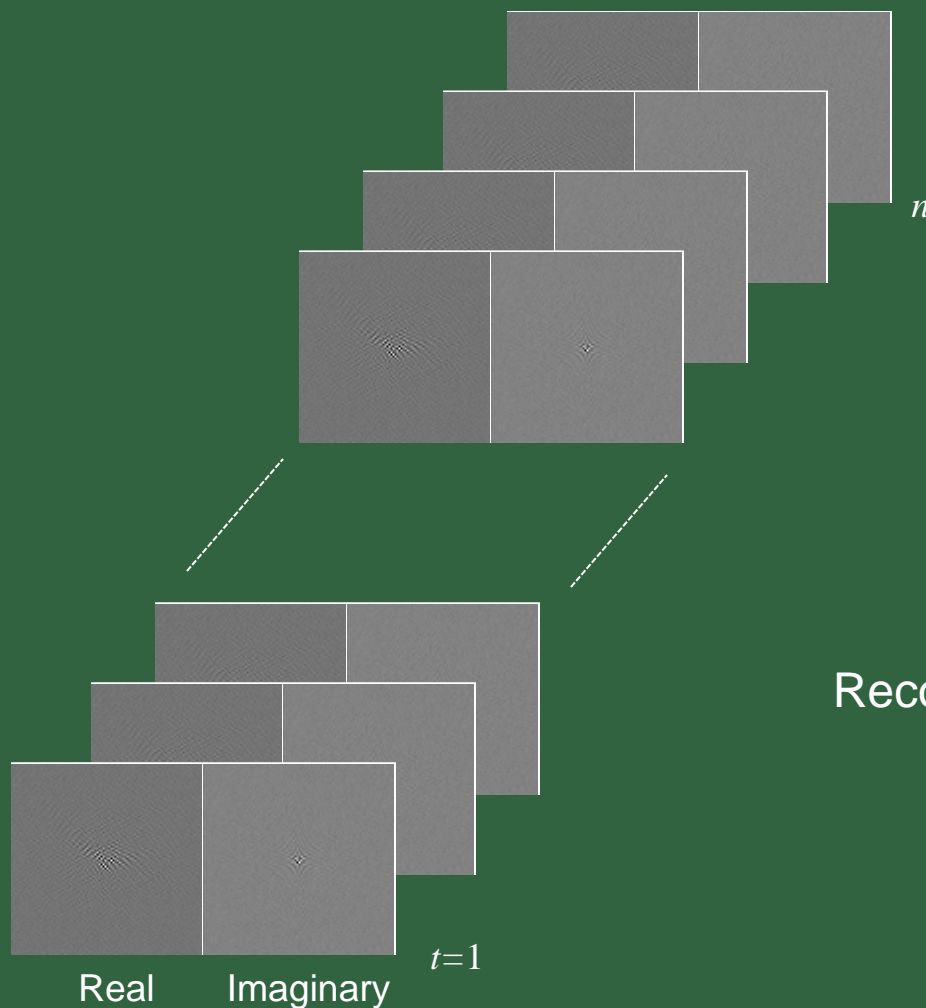
The high SNR normality of  $m_j$  can be seen as

$$\begin{aligned}
 m_j &= \left[ (y_{Rj})^2 + (y_{Ij})^2 \right]^{1/2} \\
 &= \left[ (\rho_j \cos \theta_j + \eta_{Rj})^2 + (\rho_j \sin \theta_j + \eta_{Ij})^2 \right]^{1/2} \\
 &= \left[ \rho_j^2 + (\eta_{Rj}^2 + \eta_{Ij}^2) + 2\rho_j(\eta_{Rj} \cos \theta_j + \eta_{Ij} \sin \theta_j) \right]^{1/2} \\
 &= \rho_j \left[ 1 + 2 \frac{(\eta_{Rj} \cos \theta_j + \eta_{Ij} \sin \theta_j)}{\rho_j} + \frac{(\eta_{Rj}^2 + \eta_{Ij}^2)}{\rho_j^2} \right]^{1/2} \\
 &\approx \rho_j + \varepsilon_j
 \end{aligned}$$

where  $\varepsilon_j = \eta_{Rj} \cos \theta_j + \eta_{Ij} \sin \theta_j$   
 $\varepsilon_j \sim N(0, \sigma^2)$   
 $\sqrt{1+u^2} \approx 1+u/2, \quad |u| \ll 1$

## Statistics:

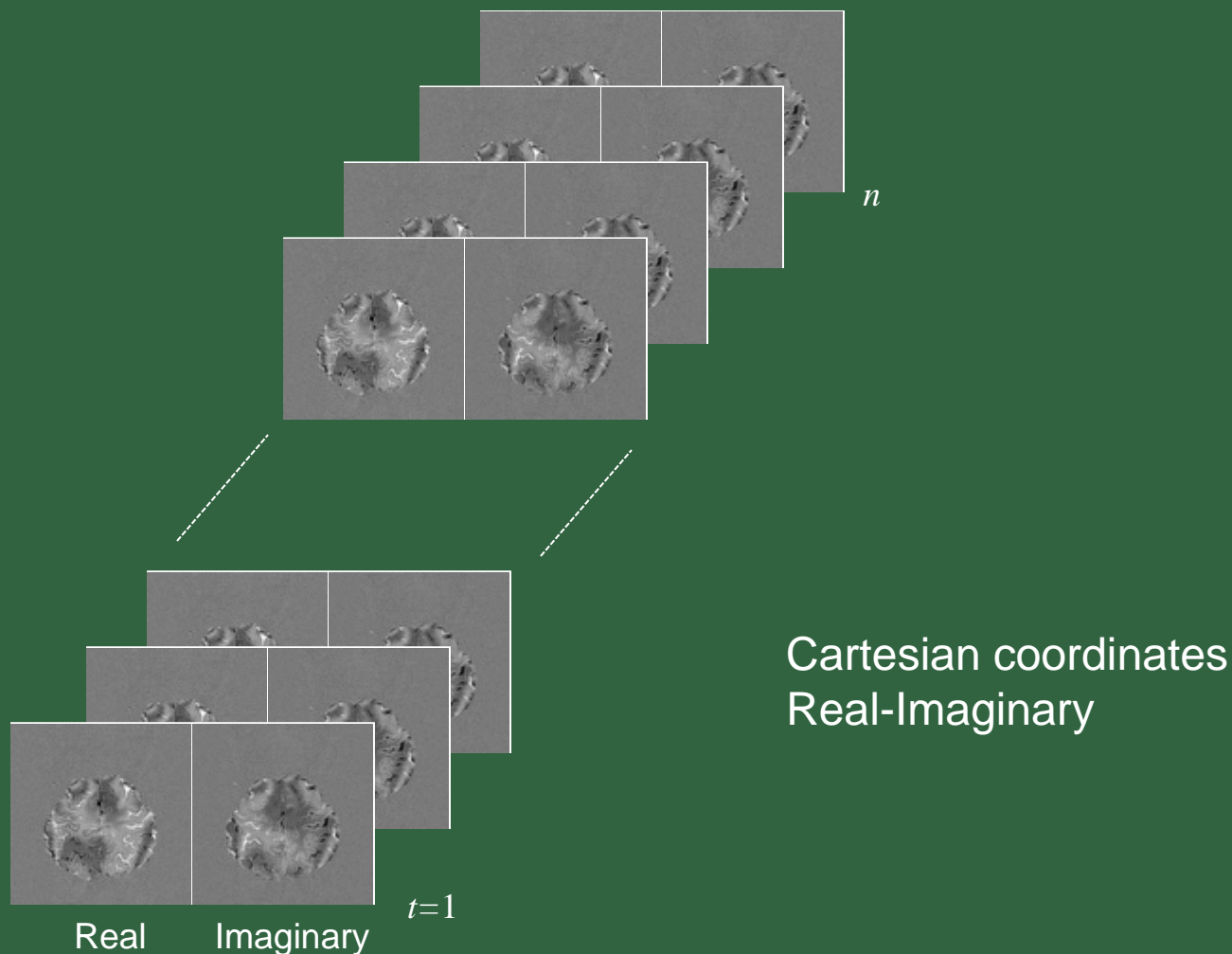
We take  $n$   $k$ -space arrays under different signal conditions





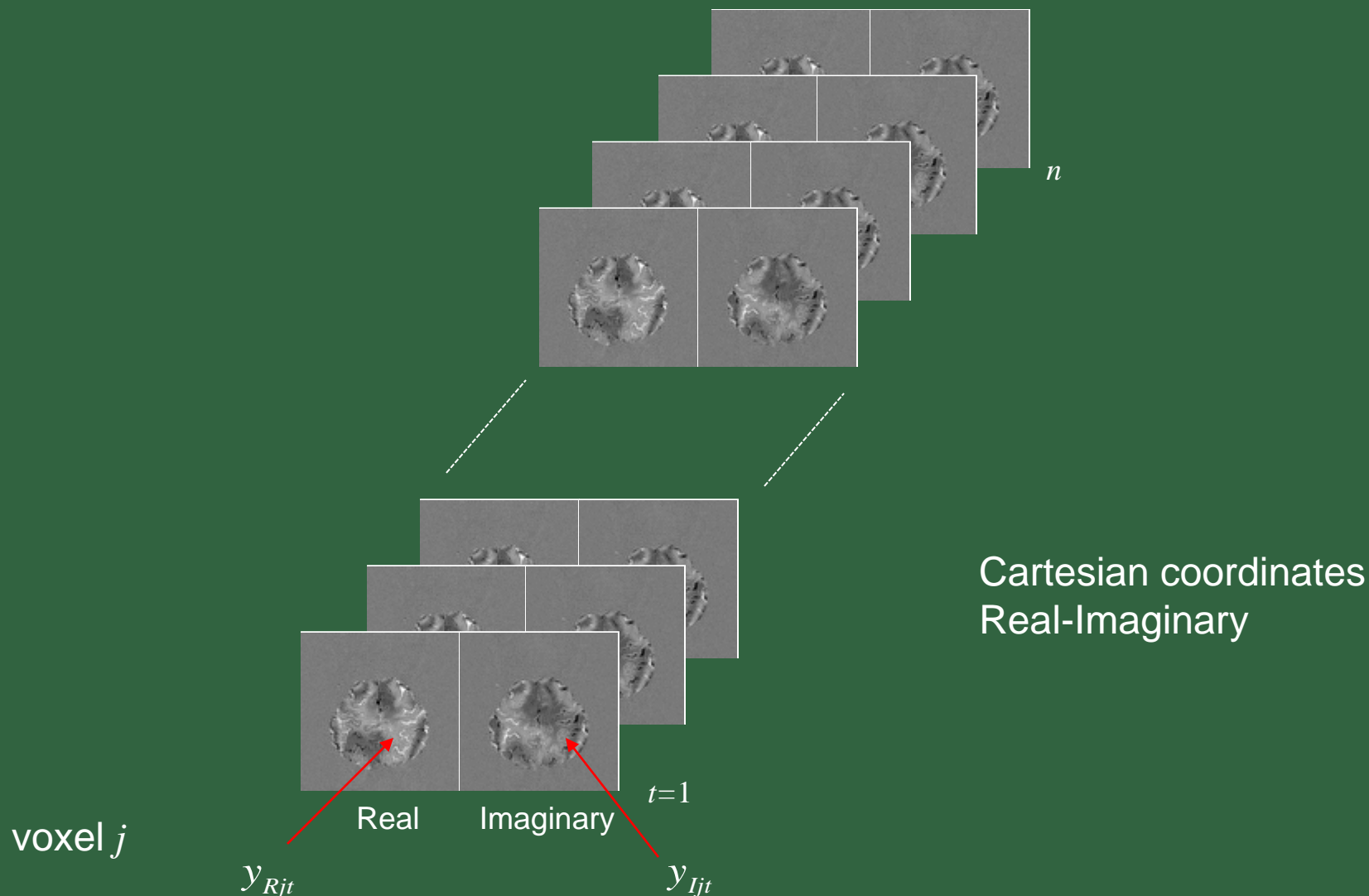
## Statistics:

We get  $n$  images under different signal conditions



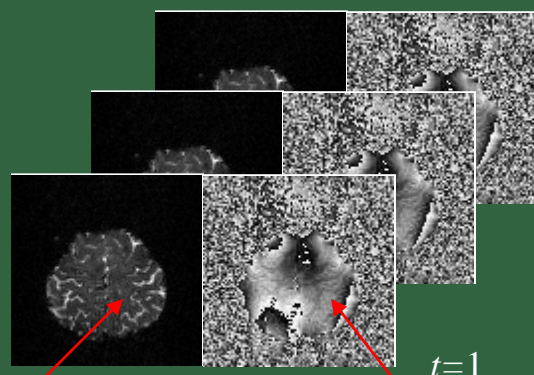
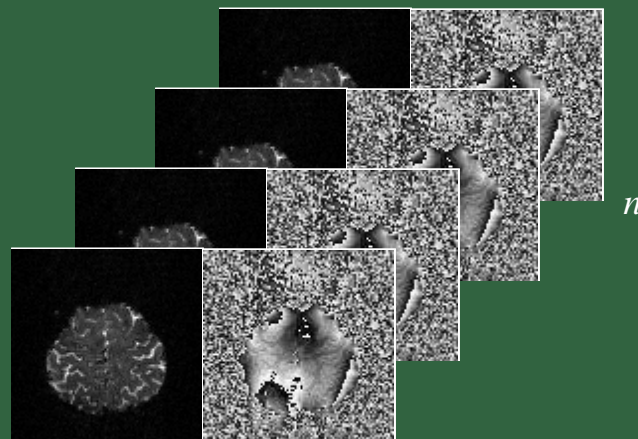
## Statistics:

We get  $n$  images under different signal conditions



## Statistics:

We get  $n$  images under different signal conditions



Polar Coordinates  
Magnitude-Phase

voxel  $j$

$m_{jt}$

Magnitude

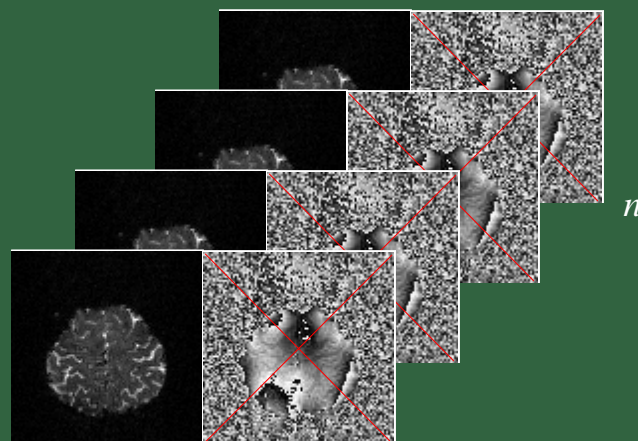
Phase

$t=1$

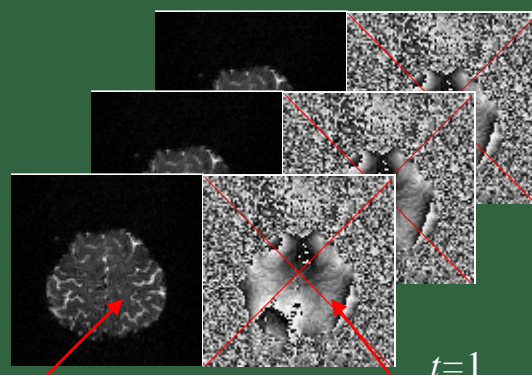
$\varphi_{jt}$

## Statistics:

We get  $n$  images under different signal conditions



Phase discarded  
in a lot of MRI  
especially fMRI



Polar Coordinates  
Magnitude-Phase

voxel  $j$

$m_{jt}$

Magnitude

Phase

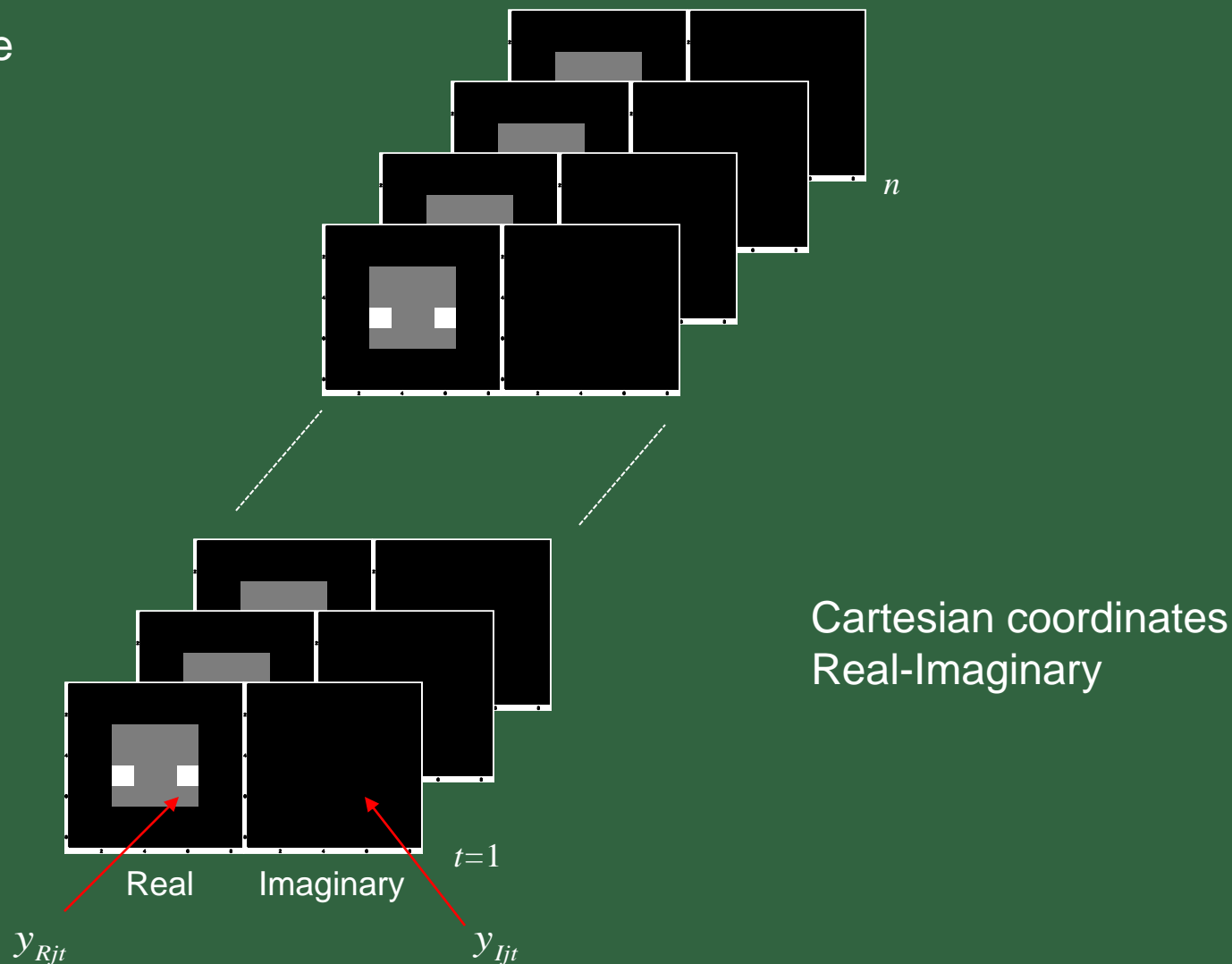
$t=1$

$\varphi_{jt}$

# Statistics:

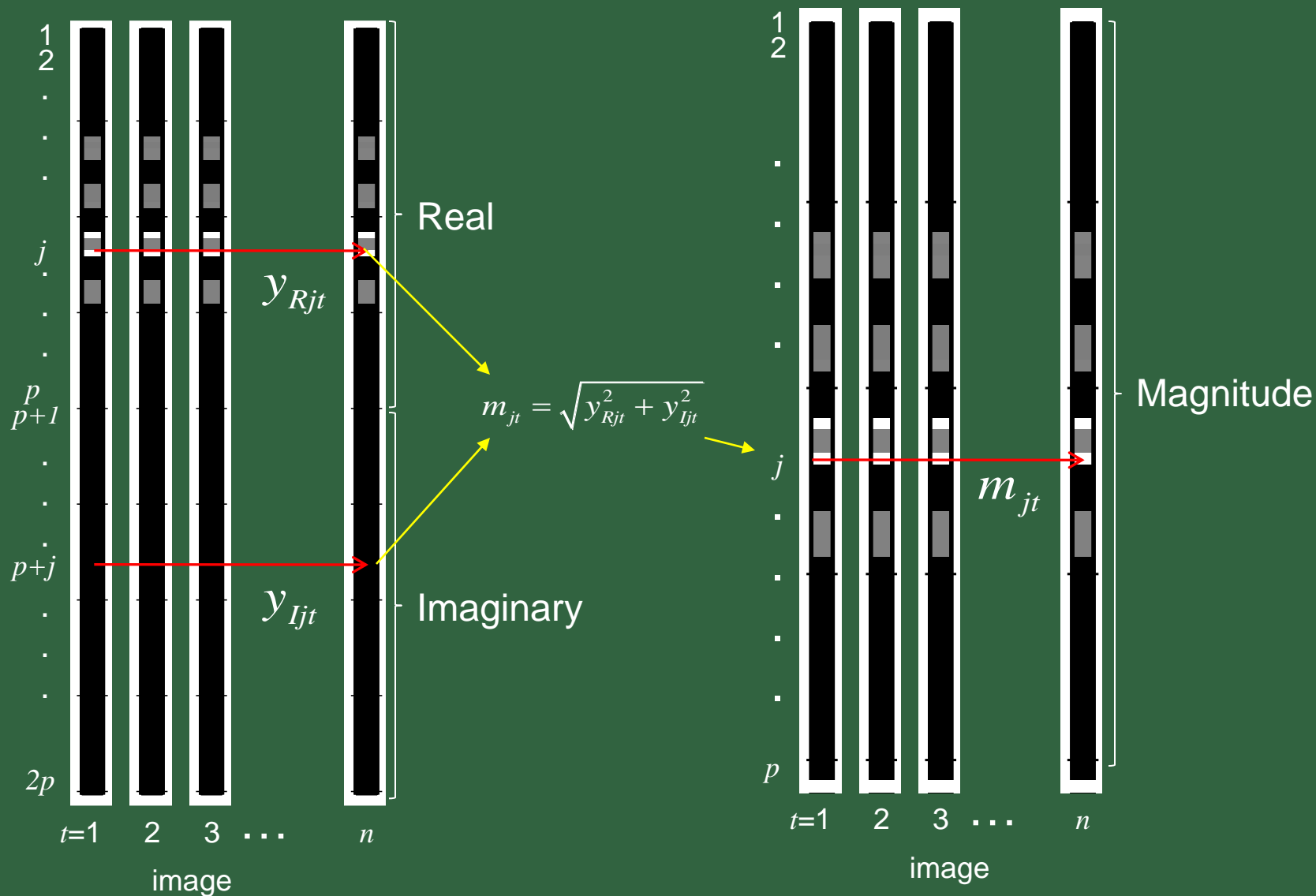
We take  $n$  images under different signal conditions

Toy Example



**Statistics:**  $y_t = \Omega s_t$

We take  $n$  images under different signal conditions



## Statistics: Bivariate Normal to Ricean

The distribution of measurement  $t$  in voxel  $j$  is:

$$p(y_{Rjt}, y_{Ijt}) = \frac{1}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ (y_{Rjt} - \rho_{jt} \cos \theta_{jt})^2 + (y_{Ijt} - \rho_{jt} \sin \theta_{jt})^2 \right] \right\}$$

$$p(m_{jt}) = \frac{m_{jt}}{\sigma^2} \exp \left\{ -\frac{m_{jt}^2 + \rho_{jt}^2}{2\sigma^2} \right\} I_0 \left( \frac{\rho_{jt} m_{jt}}{\sigma^2} \right) \quad t = 1, \dots, n$$

The goal is to estimate a functional form  $\rho_{jt} = f(x_t / \beta_j)$  for the magnitude and possibly  $\theta_{jt} = g(u_t / \gamma_j)$  for the phase from the data  $y_{Cj1}, \dots, y_{Cjn}$  or  $m_{j1}, \dots, m_{jn}$  in each voxel.

$x$  is a vector of known “dial” settings  
 $\beta$  is a vector of unknown parameters.

$$\mu_{Rjt} = \omega_j s_{0t} = \rho_{jt} \cos \theta_{jt}$$

$$\mu_{Ijt} = \omega_{p+j} s_{0t} = \rho_{jt} \sin \theta_{jt}$$

$\omega_j$  is  $j^{\text{th}}$  row of  $\Omega$

$\omega_{p+j}$  is  $(p+j)^{\text{th}}$  row of  $\Omega$

**Estimation:**

Types of functions to estimate:

Data  $y_{C1}, \dots, y_{Cn}$  or  $m_1, \dots, m_n$  in each voxel. No  $j$  subscript.

	$f(x   \beta)$	$x$	$\beta$
1.	$\rho$	1	$\rho$
2.	$\rho \exp(-TE / T_2)$	$TE$	$\rho, T_2$
3.	$S_0 \exp(-br'Dr)$	$b, r$	$S_0, D$
4.	$\rho(1 - 2 \exp(t / T_1))$	$t$	$\rho, T_1$
5.	$x' \beta$	$x'$	$\beta$



## Estimation: Ricean

Estimate parameters of function from magnitude data:

$$p(m_t) = \frac{m_t}{\sigma^2} \exp \left\{ -\frac{m_t^2 + (f(x_t | \beta))^2}{2\sigma^2} \right\} I_0 \left( \frac{f(x_t | \beta)m_t}{\sigma^2} \right) \quad t = 1, \dots, n$$

$$L = \frac{\prod_{t=1}^n m_t}{\sigma^{2n}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{t=1}^n [m_t^2 + (f(x_t | \beta))^2] \right\} \prod_{t=1}^n I_0 \left( \frac{f(x_t | \beta)m_t}{\sigma^2} \right)$$

$$LL = -n \log(\sigma^2) + \sum_{t=1}^n \log(m_t) - \frac{1}{2\sigma^2} \sum_{t=1}^n [m_t^2 + (f(x_t | \beta))^2] + \sum_{t=1}^n \log \left[ I_0 \left( \frac{f(x_t | \beta)m_t}{\sigma^2} \right) \right]$$

Maximize  $LL$ :  $\frac{\partial LL}{\partial \beta} = 0$  and  $\frac{\partial LL}{\partial \sigma^2} = 0$  Under  $H_1$  and  $H_0$

## Estimation: Large SNR Normal

Ricean

$$p(m_t) = \frac{m_t}{\sigma^2} \exp \left\{ -\frac{m_t^2 + (f(x_t | \beta))^2}{2\sigma^2} \right\} I_0 \left( \frac{f(x_t | \beta)m_t}{\sigma^2} \right)$$

Normal as SNR  $\uparrow$

$$p(m_t) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2\sigma^2} [m_t - f(x_t | \beta)]^2 \right\}$$

Then use usual least squares estimation.

$$\text{Maximize } LL: \frac{\partial LL}{\partial \beta} = 0 \text{ and } \frac{\partial LL}{\partial \sigma^2} = 0 \quad \text{Under } H_1 \text{ and } H_0$$

$$LL = -2n \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n [m_t - f(x_t | \beta)]^2$$

## Estimation: Large SNR Normal

$$LL = -2n \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n [m_t - f(x_t | \beta)]^2$$

Under  $H_1$ :

$$\frac{\partial LL}{\partial \beta} = -\frac{2}{\sigma^2} \sum_{t=1}^n [m_t - f(x_t | \beta)] \frac{\partial f(x_t | \beta)}{\partial \beta}$$

Under  $H_0$ : add Lagrange constraint  $h(\beta, \sigma^2)$  to  $LL$

$$\frac{\partial LL}{\partial \beta} = -\frac{2}{\sigma^2} \sum_{t=1}^n [m_t - f(x_t | \beta)] \frac{\partial f(x_t | \beta)}{\partial \beta} + \frac{\partial h(\beta, \sigma^2)}{\partial \beta}$$

Under  $H_0$  and  $H_1$ :

$$\frac{\partial LL}{\partial \sigma^2} = -\frac{2n}{\sigma^2} - \frac{1}{\sigma^4} \sum_{t=1}^n [m_t - f(x_t | \beta)]^2 \left( + \frac{\partial h(\beta, \sigma^2)}{\partial \sigma^2} \right)$$

May require numerical maximization depending on  $f(x_t/\beta)$ .

## Estimation: Large SNR Normal

GLM: Does not require numerical maximization.  $X$  known

$$\text{Under } H_1: \hat{\beta} = (X'X)^{-1} X'm \quad \hat{\sigma}^2 = (y - X\hat{\beta})'(y - X\hat{\beta}) / n$$

$$\text{Under } H_0: h(\beta, \sigma^2) = 2\psi' C\beta / \sigma^2$$

$$\tilde{\beta} = \Psi(X'X)^{-1} X'm \quad \tilde{\sigma}^2 = (y - X\tilde{\beta})'(y - X\tilde{\beta}) / n$$

$$\Psi = I - (X'X)^{-1} C'[C(X'X)^{-1}C']^{-1}C$$

Insert back into likelihoods and take ratio.

$$\lambda = L(\tilde{\beta}, \tilde{\sigma}^2) / L(\hat{\beta}, \hat{\sigma}^2)$$

This is how we get our usual  $t$  and  $F$  statistics.

## Estimation: Large SNR Normal

DTI: Requires numerical maximization.  $b$  and  $r_t$  known

$$LL = -2n \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n [m_t - S_0 \exp(-br_t' D r_t)]^2$$

Under  $H_1$ :

$$\frac{\partial LL}{\partial S_0} = -\frac{2}{\sigma^2} \sum_{t=1}^n [m_t - S_0 \exp(-br_t' D r_t)] \frac{\partial S_0 \exp(-br_t' D r_t)}{\partial S_0}$$

$$\hat{S}_0 | \hat{D} = \left[ \sum_{t=1}^n m_t \exp(-br_t' \hat{D} r_t) \right] / \left[ \sum_{t=1}^n \exp(-2br_t' \hat{D} r_t) \right]$$

$$\frac{\partial LL}{\partial D} = 0 \quad \text{Does not yield a closed form solution.}$$

Need numerical maximization with say  
Newton-Raphson or Levenberg-Marquardt.

## Estimation: Large SNR Normal

Numerical maximization.

$$LL = -\frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t - \sqrt{f(x_t | \beta)^2 + \sigma^2} \right]^2$$

$$\beta^{(0)} : \sum_{t=1}^n [m_t - f(x_t | \beta)]^2 \quad \text{Minimized by Levenberg-Marquardt}$$

$$(\sigma^2)^{(0)} = \sum_{t=1}^n [m_t - f(x_t | \beta^{(0)})]^2 / n$$

$$\beta^{(r+1)} : \sum_{t=1}^n \left[ m_t - \sqrt{f(x_t | \beta)^2 + (\sigma^2)^{(r)}} \right]^2 \quad \text{Minimize by Levenberg-Marquardt}$$

$$(\sigma^2)^{(0)} : \quad LL = -\frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t - \sqrt{f(x_t | \beta)^2 + \sigma^2} \right]^2$$

Minimized by Newton-Raphson

$\beta^{(0)}, (\sigma^2)^{(0)}, \beta^{(1)}, (\sigma^2)^{(1)}, \dots, \beta^{(r+1)}, (\sigma^2)^{(r+1)}$  sequence converges to MLE!

## Estimation: Small SNR Ricean

$$LL = -n \log(\sigma^2) + \sum_{t=1}^n \log(m_t) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + (f(x_t | \beta))^2 \right] \\ + \sum_{t=1}^n \log \left[ I_0 \left( f(x_t | \beta) m_t / \sigma^2 \right) \right]$$

Under  $H_1$ :  $A_t = f(x_t | \beta) m_t / \sigma^2$

$$\frac{\partial LL}{\partial \beta} = \frac{1}{\sigma^2} \sum_{t=1}^n \left[ m_t I_1(A_t) / I_0(A_t) - f(x_t | \beta) \right] \frac{\partial f(x_t | \beta)}{\partial \beta}$$

Under  $H_0$ :

$$\frac{\partial LL}{\partial \beta} = \frac{1}{\sigma^2} \sum_{t=1}^n \left[ m_t I_1(A_t) / I_0(A_t) - f(x_t | \beta) \right] \frac{\partial f(x_t | \beta)}{\partial \beta} + \frac{\partial h(\beta, \sigma^2)}{\partial \beta}$$

Under  $H_1$  and  $H_0$

$$\frac{\partial LL}{\partial \sigma^2} = \frac{1}{2\sigma^4} \left[ m_t^2 + (f(x_t | \beta))^2 - 2m_t A_t f(x_t | \beta) - 2n\sigma^2 \right] \left( + \frac{\partial h(\beta, \sigma^2)}{\partial \sigma^2} \right)$$

No closed form solution. Requires numerical maximization!

**Estimation:** Small SNR Ricean

EM Algorithm. Easier and convenient. Does not need phase.

Take magnitude variates  $m_1, \dots, m_n$  that are Ricean distributed

$$p(m_t) = \frac{m_t}{\sigma^2} \exp\left\{-\frac{m_t^2 + f(x_t | \beta)^2}{2\sigma^2}\right\} I_0\left(\frac{f(x_t | \beta)m_t}{\sigma^2}\right)$$

Introduce latent phase variables  $\phi_1, \dots, \phi_n$  such that

$$p(m_t, \phi_t) = \frac{m_t}{2\pi\sigma^2} \exp\left[-(m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos \phi_t) / 2\sigma^2\right]$$

and

$$LL = -n \log(2\pi\sigma^2) + \sum_{t=1}^n \log(m_t) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos(\phi_t) \right]$$



## Estimation: Small SNR EM Algorithm. Iterative.

$$LL = -n \log(2\pi\sigma^2) + \sum_{t=1}^n \log m_t - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos \phi_t \right]$$

E Step: Let  $Y_m = (m_1, \dots, m_n)$ ,  $Y_\phi = (\phi_1, \dots, \phi_n)$ ,  $Y_x = (x_1, \dots, x_n)$

given  $\beta^{(r)}, (\sigma^2)^{(r)}$ : Initial values from normal GLM

$$E[L_c(\beta, \sigma^2 | Y_m, Y_\phi, Y_x) | Y_m, Y_x, \beta^{(r)}, (\sigma^2)^{(r)}] =$$

$$-n \log(\sigma^2)^{(r)} - \frac{1}{2(\sigma^2)^{(r)}} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta^{(r)})^2 - 2m_t f(x_t | \beta^{(r)}) A_t^{(r)} \right]$$

$$A_t^{(r)} = f(x_t | \beta^{(r)}) m_t / (\sigma^2)^{(r)}$$

with respect to  $p(Y_\phi | Y_m, Y_x, \beta^{(r)}, (\sigma^2)^{(r)}) = \prod_{t=1}^n p(\phi_t | m_t, \beta^{(r)}, (\sigma^2)^{(r)})$

# Estimation: Small SNR

## EM Algorithm. Iterative.

M Step:

given  $\beta^{(r)}, (\sigma^2)^{(r)}$  :

$$(\sigma^2)^{(r+1)} = \frac{1}{2n} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta^{(r)})^2 - 2m_t f(x_t | \beta^{(r)}) A_t^{(r)} \right]$$

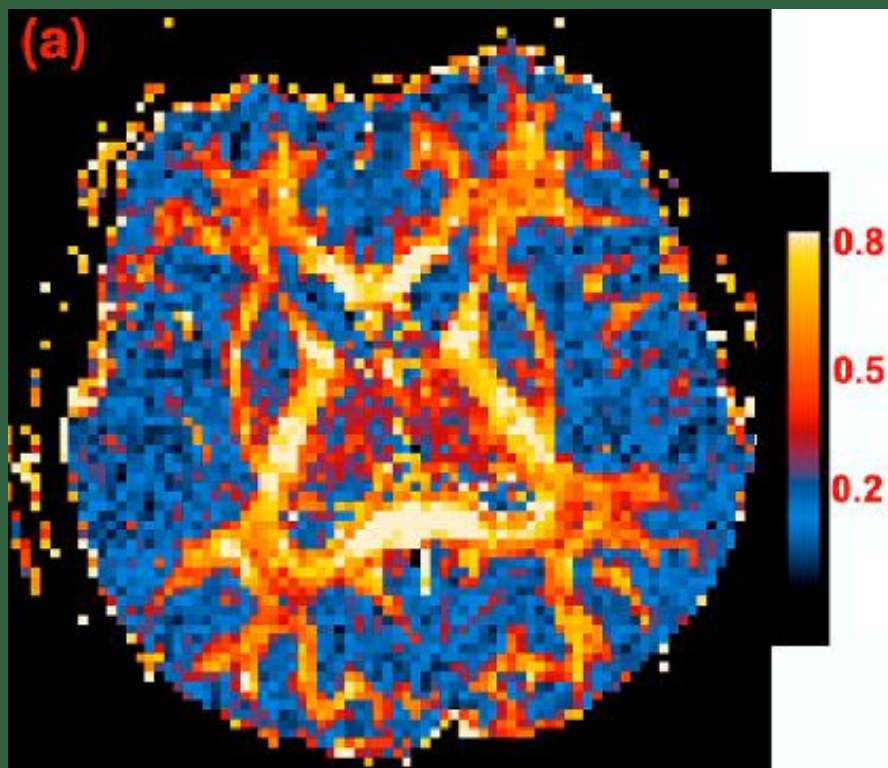
$$A_t^{(r)} = f(x_t | \beta^{(r)}) m_t / (\sigma^2)^{(r)}$$

$$\beta^{(r+1)} : \text{minimize } \sum_{t=1}^n \left[ f(x_t | \beta)^2 - m_t A_t^{(r)} \right]^2 \quad \text{given } (\sigma^2)^{(r+1)}$$

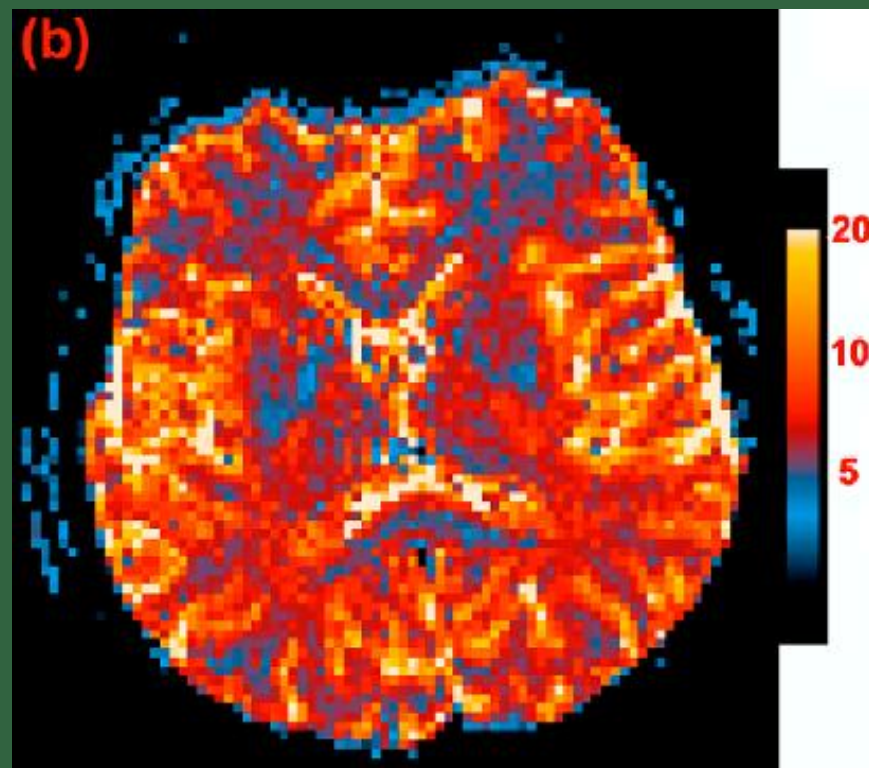
$\beta^{(0)}, (\sigma^2)^{(0)}, \beta^{(1)}, (\sigma^2)^{(1)}, \dots, \beta^{(r+1)}, (\sigma^2)^{(r+1)}$  sequence converges to MLE!

# Estimation: Small SNR EM Algorithm.

$$f(S_0, D | r, b) = S_0 \exp(-br'Dr)$$



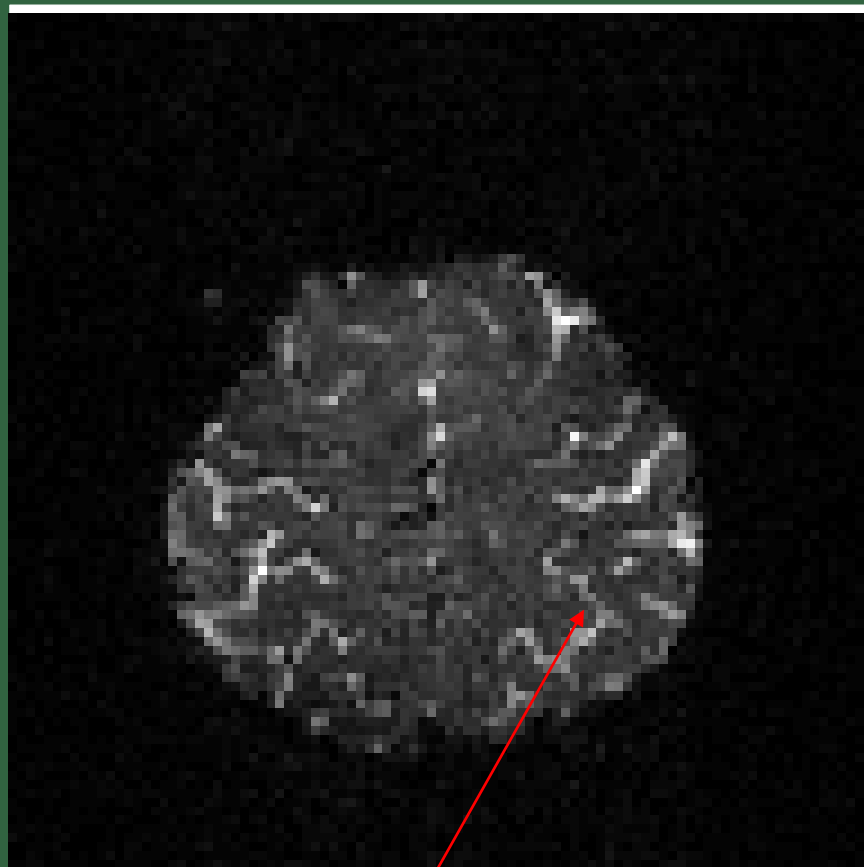
Fractional Anisotropy, FA



Signal-to-Noise Ratio,  $S_0/\sigma^2$

# Estimation: Bivariate Normal

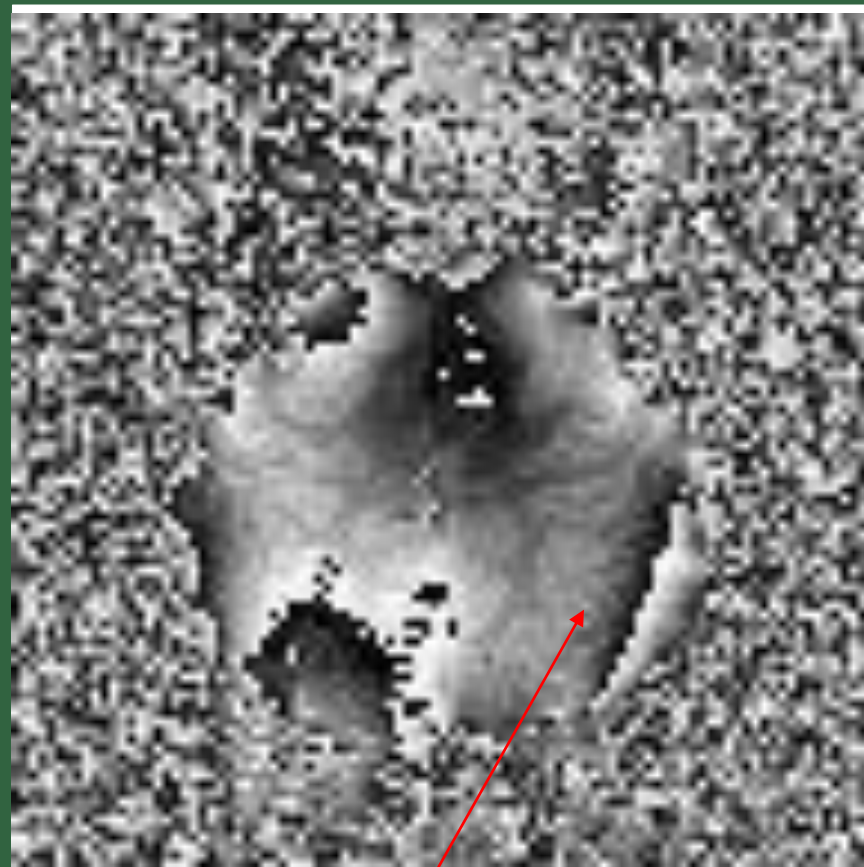
## Magnitude Image



voxel  $j$

$$m_j = \sqrt{y_{Rj}^2 + y_{Ij}^2}$$

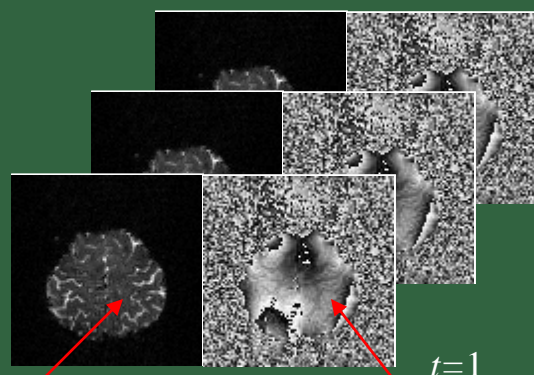
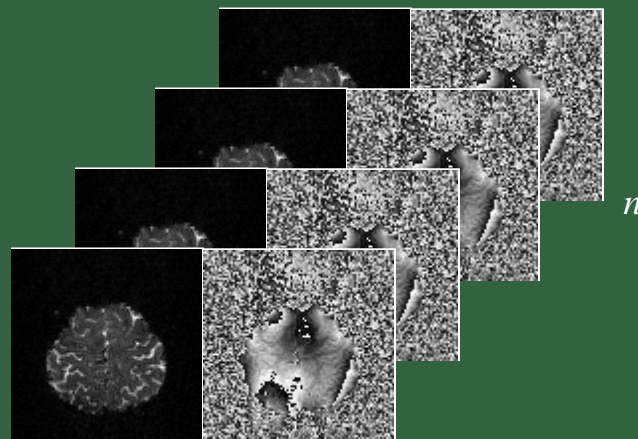
## Phase Image



$$\varphi_j = \tan^{-1}(y_{Ij} / y_{Rj})$$

## Statistics:

We get  $n$  images under different signal conditions



voxel  $j$

$m_{jt}$

Magnitude

Phase

$t=1$

$\varphi_{jt}$

Polar Coordinates  
Magnitude-Phase

## Estimation: All SNRs Bivariate Normal

$$\begin{pmatrix} y_{Rt} \\ y_{It} \end{pmatrix} = \begin{pmatrix} \rho_t \cos \theta_t \\ \rho_t \sin \theta_t \end{pmatrix} + \begin{pmatrix} \eta_{Rt} \\ \eta_{It} \end{pmatrix}, \quad \begin{pmatrix} \eta_{Rt} \\ \eta_{It} \end{pmatrix} \sim N(\mathbf{0}, \Sigma)$$

$$p(y_{Rt}, y_{It}) = \frac{1}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ (y_{Rt} - \rho_t \cos \theta_t)^2 + (y_{It} - \rho_t \sin \theta_t)^2 \right] \right\}$$

$$p(m_t, \varphi_t) = \frac{m_t}{2\pi\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} \left[ m_t^2 + \rho_t^2 - 2m_t\rho_t \cos(\varphi_t - \theta_t) \right] \right\}$$

$$\rho_t = f(x_t | \beta) \text{ and } \theta_t = g(u_t | \gamma)$$

$$LL = -n \log(2\pi\sigma^2) + \sum_{t=1}^n \log(m_t)$$

$$- \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos(\varphi_t - g(u_t | \gamma)) \right]$$

## Estimation: All SNR Bivariate Normal

$$LL = -n \log(2\pi\sigma^2) + \sum_{t=1}^n \log(m_t) \\ - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos(\varphi_t - g(u_t | \gamma)) \right]$$

$$\frac{\partial LL}{\partial \beta} = \frac{1}{\sigma^2} \sum_{t=1}^n \left[ m_t \cos(\varphi_t - g(u_t | \gamma)) - f(x_t | \beta) \right] \frac{\partial f(x_t | \beta)}{\partial \beta}$$

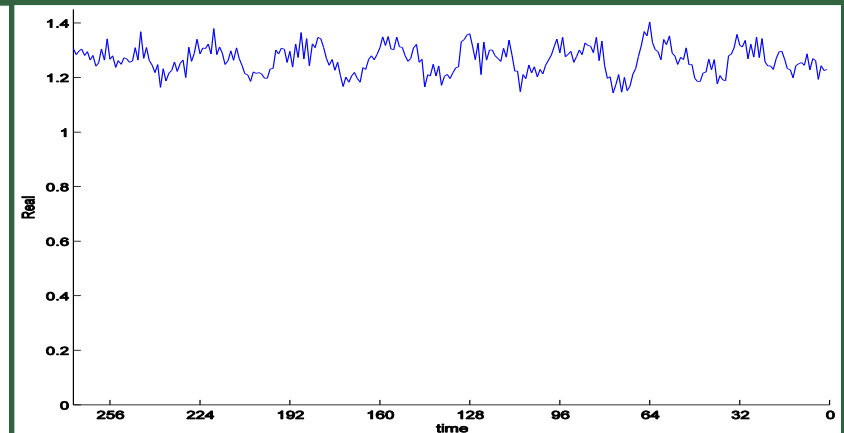
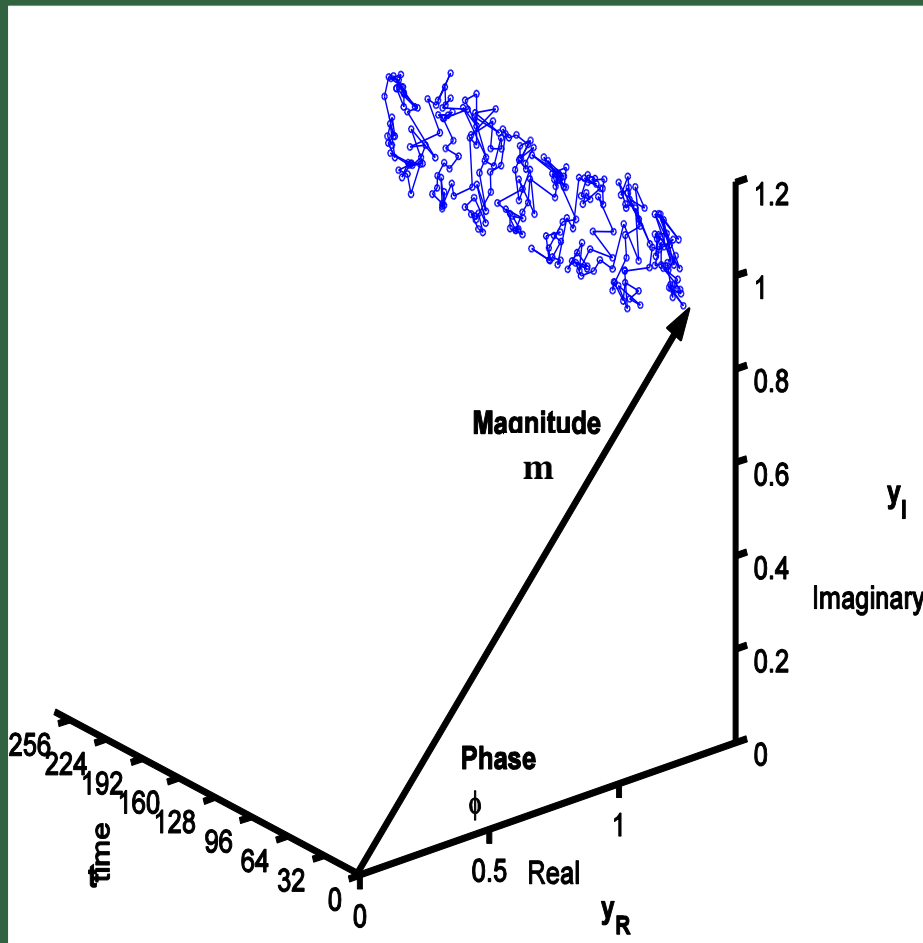
$$\frac{\partial LL}{\partial \gamma} = \frac{1}{\sigma^2} \sum_{t=1}^n \left[ m_t f(x_t | \beta) \sin(\varphi_t - g(u_t | \gamma)) \right] \frac{\partial g(u_t | \gamma)}{\partial \gamma}$$

$$\frac{\partial LL}{\partial \sigma^2} = \frac{n}{\sigma^2} + \frac{1}{2\sigma^4} \sum_{t=1}^n \left[ m_t^2 + f(x_t | \beta)^2 - 2m_t f(x_t | \beta) \cos(\varphi_t - g(u_t | \gamma)) \right]$$

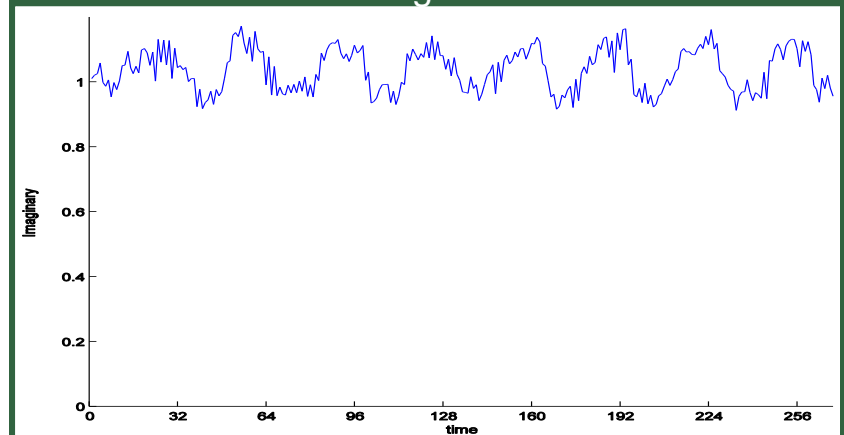
$\sigma^2$  can be uniquely solved for given  $\beta, \gamma$

# Estimation:

Time series are complex, bivariate with phase coupled means.



Real: Task related changes!



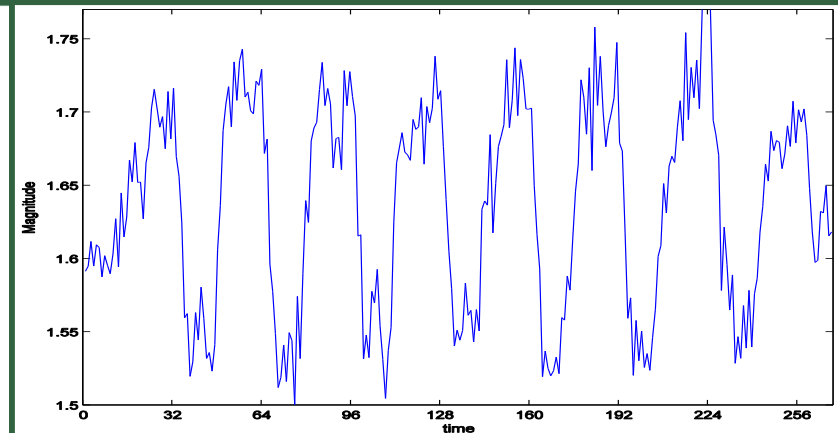
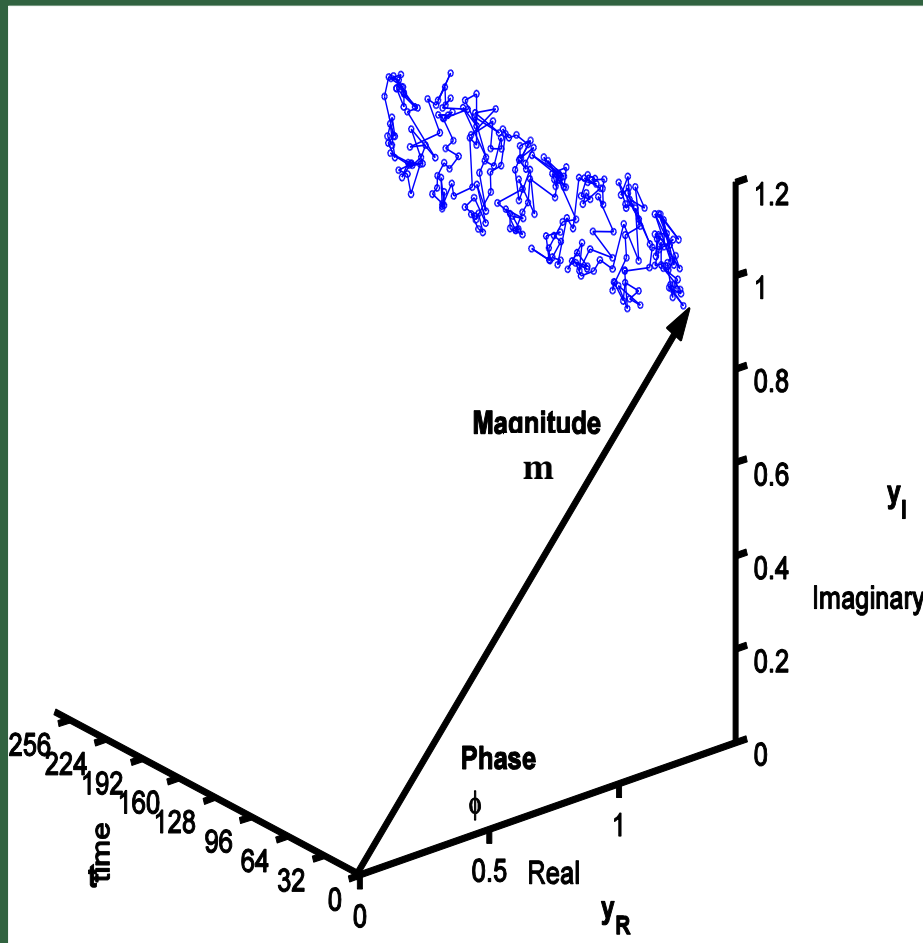
Imaginary: Task related changes!

The  $y_R$  and  $y_I$  time courses have related info! From actual human data!

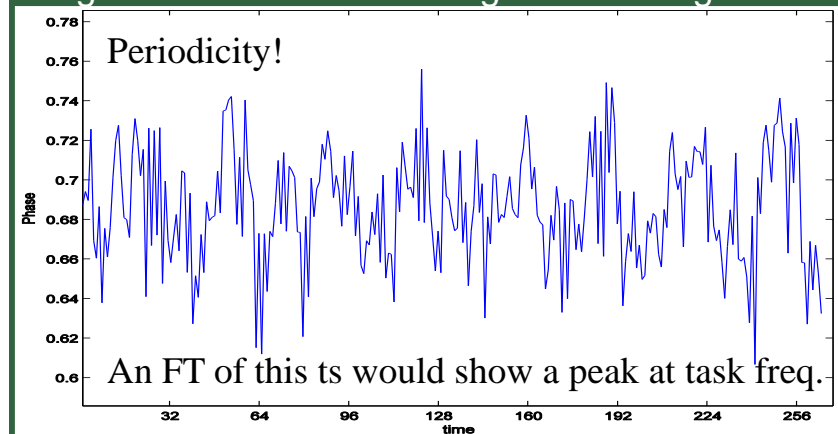


# Estimation:

Time series are complex, bivariate with phase coupled means.



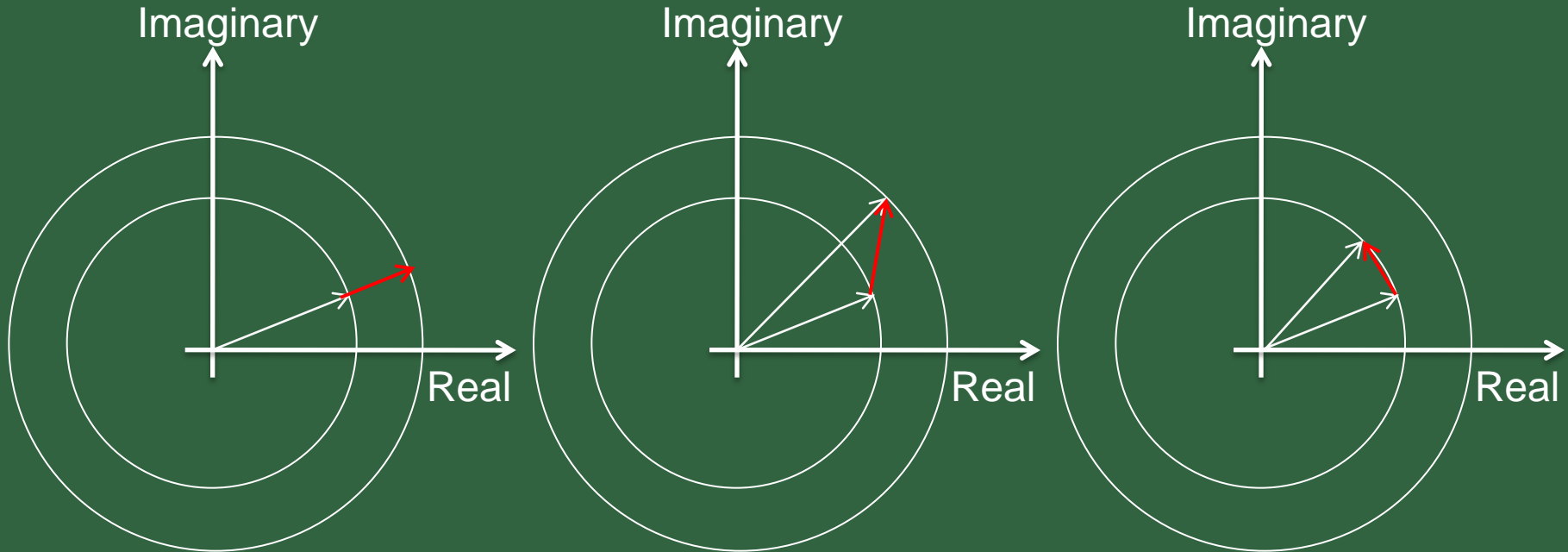
Magnitude: Task related magnitude changes!



Phase: Task related phase changes!

## Estimation:

Magnitude and or phase change.



$$\rho_t = x_t' \beta \text{ and } \theta_t = u_t' \gamma$$

<sup>1</sup>Rowe and Logan: NIMG, 23:1078-1092, 2004.

<sup>3</sup>Rowe: NIMG, 25:1310-1324, 2005b.

<sup>5</sup>Rowe and Logan: NIMG 24:603-606, 2005.

<sup>7</sup>Rowe: MRM, to appear, 2009.

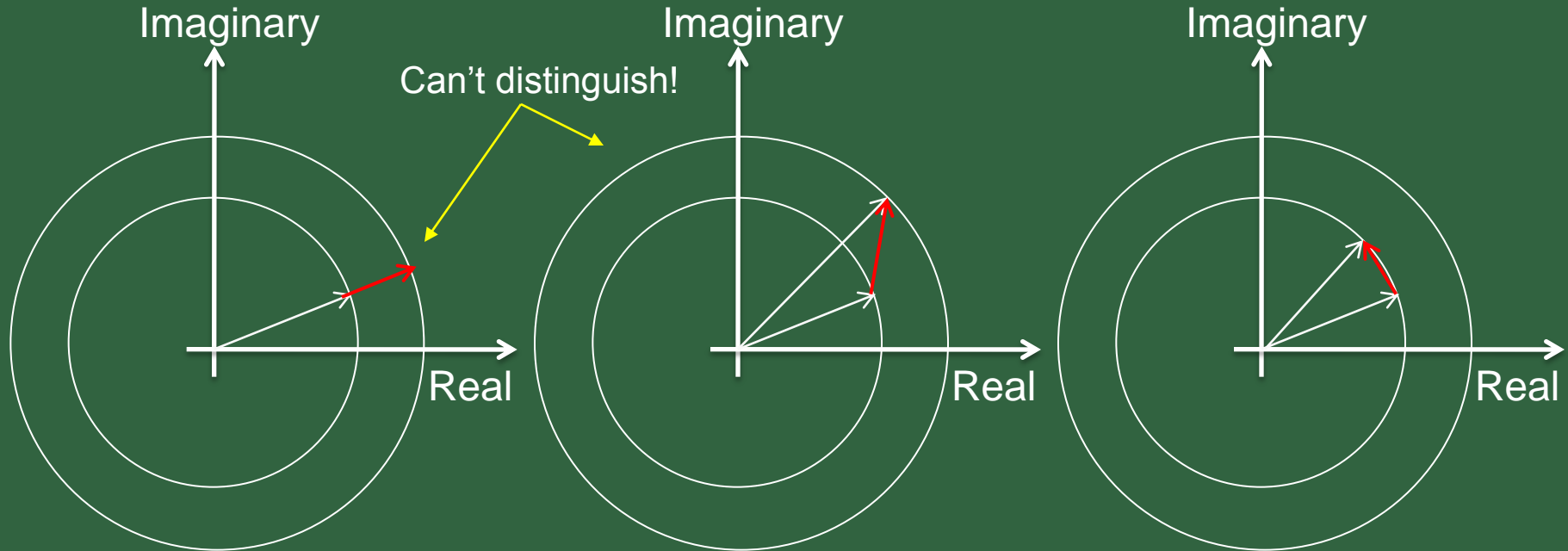
<sup>2</sup>Rowe: NIMG 25:1124-1132, 2005a.

<sup>4</sup>Bandettini et al.: MRM, 30:161-173, 1993.

<sup>6</sup>Rowe, et al.: JNeuroSciMeth, 161:331-341, 2007.

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<sup>7</sup>Rowe: MRM, to appear, 2009.

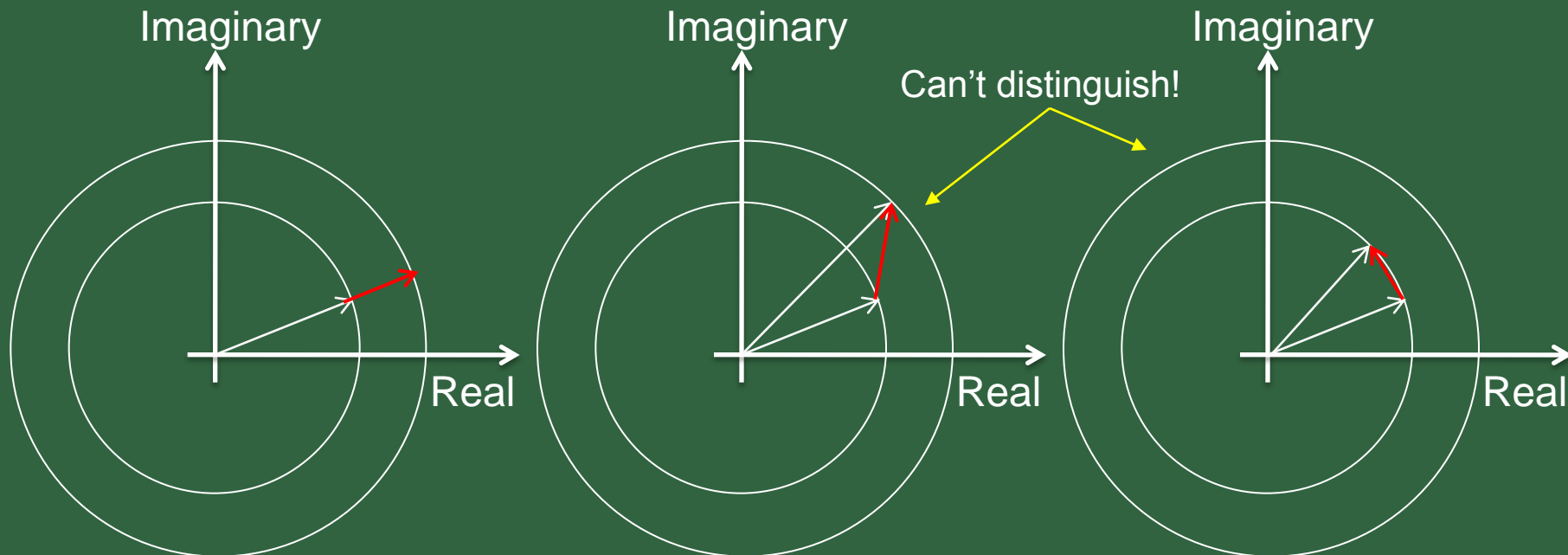
<sup>2</sup>Rowe: NIMG 25:1124-1132, 2005a.

<sup>4</sup>Bandettini et al.: MRM, 30:161-173, 1993.

<sup>6</sup>Rowe, et al.: JNeuroSciMeth, 161:331-341, 2007.

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<sup>6</sup>Rowe, et al.: JNeuroSciMeth, 161:331-341, 2007.

# Estimation: All SNR

GLM:

$$\rho_t = x_t' \beta \quad \text{and} \quad \theta_t = u_t' \gamma \quad H_0 : \mathcal{C} \beta = 0 \quad \text{vs.} \quad H_1 : \mathcal{C} \beta \neq 0$$

$$\mathcal{D} \gamma = 0 \quad \mathcal{D} \gamma \neq 0$$

$$LL = -n \log(2\pi\sigma^2) + \sum_{t=1}^n \log(m_t) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left[ m_t^2 + (x_t' \beta)^2 - 2m_t x_t' \beta \cos(\varphi - u_t' \gamma) \right]$$

Maximize  $LL$ : under  $H_1$  and  $H_0$

$\beta^{(0)}$  : initial value

$\varphi_*^{(r)}$  has elements  $\varphi_t \sqrt{m_t x_t' \hat{\beta}^{(r)}}$

$$\hat{\gamma}^{(r)} = (\hat{\mathbf{Z}}_{(r)}' \hat{\mathbf{Z}}_{(r)})^{-1} \hat{\mathbf{Z}}_{(r)}' \hat{\varphi}_*^{(r)}$$

$\hat{\mathbf{Z}}_{(r)}'$  has rows  $u_t' \sqrt{m_t x_t' \hat{\beta}^{(r)}}$

$m_*^{(r)}$  has elements  $m_t \cos(\varphi_t - u_t' \hat{\gamma}^{(r)})$

$$\hat{\beta}^{(r+1)} = (X' X)^{-1} X' m_*^{(r)}$$

$$(\hat{\sigma}^2)^{(r+1)} = \frac{1}{2n} \sum_{t=1}^n \left[ (m - X \hat{\beta}^{(r+1)})' (m - X \hat{\beta}^{(r+1)}) + 2(m - \hat{m}_*^{(r+1)})' X \hat{\beta}^{(r+1)} \right]$$

Rowe: NIMG, 25:1310-1324, 2005.

Rowe: MRM, to appear, 2009.

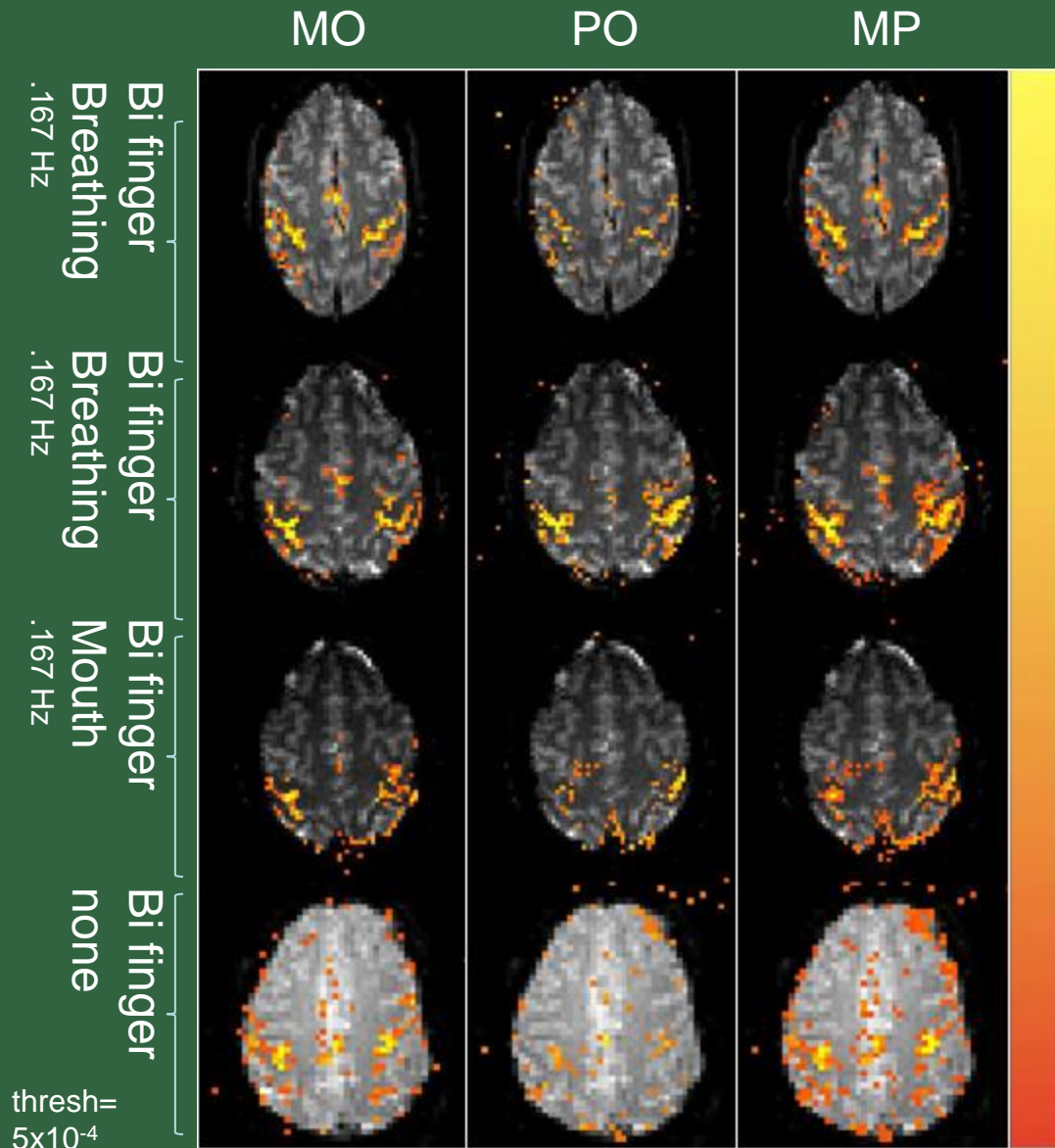
# Estimation: GLM:

20s off+16x(8 s on 8 s off), 276 TRs  
 12 axial slices, 96 x 96, FOV = 24 cm  
 TH = 2.5 mm, TR = 1 s, TE = 34.6 ms  
 FA = 45°, BW = 125 kHz, ES = .708 ms

20s off+16x(8 s on 8 s off), 276 TRs  
 10 axial slices, 96 x 96, FOV = 24 cm  
 TH = 2.5 mm, TR = 1 s, TE = 42.8 ms  
 FA = 45°, BW = 125 kHz, ES = .768 ms

20s off+16x(8 s on 8 s off), 276 TRs  
 10 axial slices, 96 x 96, FOV = 24 cm,  
 TH = 2.5 mm, TR = 1 s, TE = 42.8 ms  
 FA = 45°, BW = 125 kHz, ES = .768 ms

20s off+10x(8 s on 8 s off), 180 TRs  
 9 axial slices, 64 x 64, FOV = 24 cm  
 TH = 3.8 mm, TR = 1 s, TE = 26.0 ms  
 FA = 45°, BW = 125 kHz, ES = .680 ms



Rowe: NIMG, 25:1310-1324, 2005.  
 Rowe: MRM, to appear, 2009.

Hahn, Nencka, Rowe: NIMG, 742-752, 2009.  
 Hahn, Nencka, Rowe: In progress.

$$\Delta B_i = \frac{\arg \left( I_i \sum_{j=1}^n \left( \frac{I_j^*}{|I_j|} \right) \right)}{\gamma TE}$$

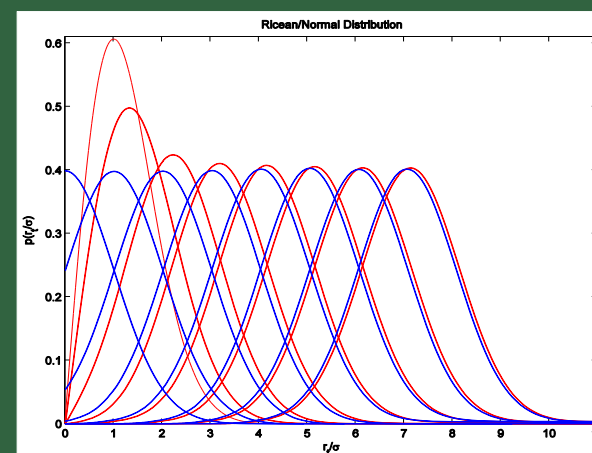
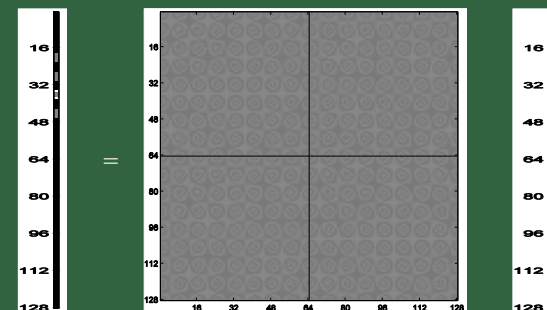
## Discussion:

- Not clear how much improvement from Ricean distribution.
- Improvements will show below SNR=5. High  $b$ -values.
- Other factors hinder it.
  - Dynamic field changes
  - Image Warping
  - Motion
  - Image Processing
- Should also use phase for complete data model.
- More biological info extracted with use of phase.

## Discussion:

1. Image Reconstruction
2. Statistics-Ricean & Normal
3. Estimation-Ricean & Normal
4. Estimation-Bivariate Normal
5. Discussion

Further research is needed ....



$$\begin{pmatrix} y_{Rt} \\ y_{It} \end{pmatrix} = \begin{pmatrix} \rho_t \cos \theta_t \\ \rho_t \sin \theta_t \end{pmatrix} + \begin{pmatrix} \eta_{Rt} \\ \eta_{It} \end{pmatrix}$$

$$\rho_t = f(x_t | \beta)$$

$$\theta_t = g(u_t | \gamma)$$



# Thank You

Questions?