

# Bayesian Statistics

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# Outline

**Background**

**Likelihood Distribution**

**Prior Distribution**

**Posterior Distribution**

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**Homework**

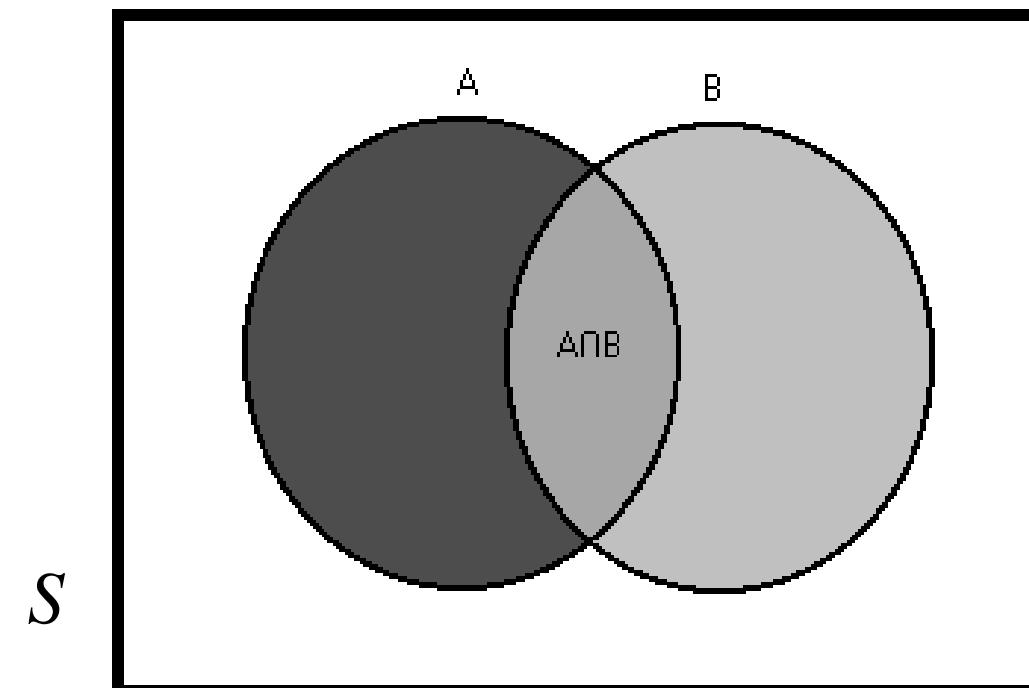
# Background

We learned about the conditional probability of  $B$  given  $A$ .

If  $A$  and  $B$  are events in  $S$ , and  $P(A)>0$ , then the *conditional probability of  $B$  given  $A$*  written is,

$$P(B | A) = \frac{P(A \cap B)}{P(A)}$$

$$P(B | A) = \frac{P(A | B)P(B)}{P(A)}$$



# Background

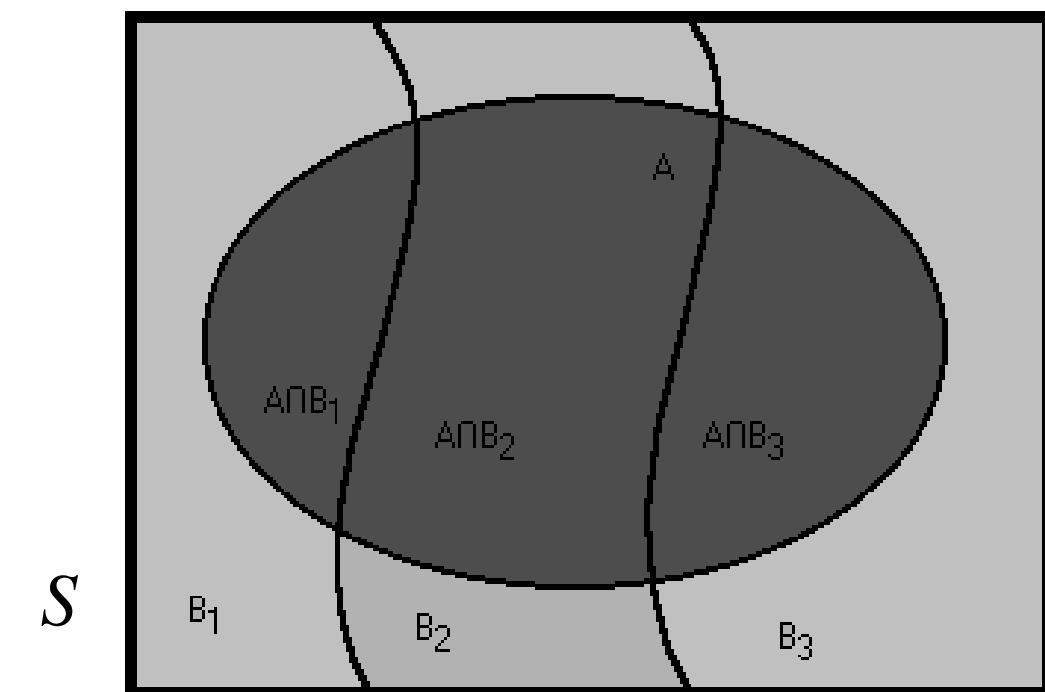
We extended to more  $A$  events,  $A_1, A_2, \dots$

Let  $B_1, B_2, \dots$  be a partition of the sample space, and let  $B$  be any set.

Then for each  $i=1,2,\dots,$

$$P(B_i | A) = \frac{P(A | B_i)P(B_i)}{P(A)}$$

$$P(A) = \sum_{i=1}^{\infty} P(A | B_i)P(B_i)$$



# Background

**Example:** Medical Test.  $P(\text{have disease}|\text{test positive})$ .

$T^+$ : The event that the test is positive.

$T^-$ : The event that the test is negative.

$D^+$ : The event that the person truly has disease.

$D^-$ : The event that the person truly does not have disease.

The sensitivity of test is  $P(T^+|D^+)=0.99$ .

The specificity of test is  $P(T^-|D^-)=0.99$ .

If the proportion of population that truly has disease is  $10^{-6}$ .

$$P(D^-|T^+) = \frac{P(T^+|D^-)P(D^-)}{P(T^+)} = 0.99990101$$

$$P(T^+) = P(T^+|D^+)P(D^+) + P(T^+|D^-)P(D^-)$$

# Likelihood

## Simple Normal Samples Model

Assume that we have  $y_i = \mu + \varepsilon_i$ , where  $\varepsilon_i$  are iid  $N(0, \sigma^2)$ , for  $i=1, \dots, n$ .

This means that given  $\mu$  and  $\sigma^2$ , the PDF of  $y_i$  is

$$f(y_i | \mu, \sigma^2) = (2\pi\sigma^2)^{-1/2} \exp\left[-\frac{1}{2\sigma^2}(y_i - \mu)^2\right]$$

and since these are independent observations, we wrote

$$\begin{aligned} f(y_1, \dots, y_n | \mu, \sigma^2) &= (2\pi\sigma^2)^{-n/2} \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right] \\ &= L(\mu, \sigma^2) \end{aligned}$$

## Prior

In MLE, we sort of heuristically turned things around.

We took  $f(y_1, \dots, y_n | \mu, \sigma^2)$  which was a (probability) function of the data  $y_1, \dots, y_n$  given  $\mu$  and  $\sigma^2$  and changed it into a function  $L(\mu, \sigma^2)$  of  $\mu$  and  $\sigma^2$  (given the data  $y_1, \dots, y_n$ ).

Why and how did this happen?

Truthfully  $L$  is the probability of getting data  $y_1, \dots, y_n$  given  $\mu$  and  $\sigma^2$  and not probability of  $\mu$  and  $\sigma^2$  given data!

# Prior

What happened to the rules of probability? i.e. Bayes' Rule

$$P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{P(A|B)P(B)}{P(A)}$$

Did we just throw out what we have learned?

To be correct, shouldn't we instead write

$$f(\underbrace{\mu, \sigma^2}_{B} | \underbrace{y_1, \dots, y_n}_{A}) = \frac{f(\underbrace{y_1, \dots, y_n}_{A} | \mu, \sigma^2) f(\mu, \sigma^2)}{f(y_1, \dots, y_n)}$$

$$A \rightarrow y_1, \dots, y_n \quad B \rightarrow \mu, \sigma^2$$

# Prior

What happened to the rules of probability? i.e. Bayes' Rule

$$P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{P(A|B)P(B)}{P(A)}$$

Did we just throw out what we have learned?

To be correct, shouldn't we instead write

$$f(\mu, \sigma^2 | y_1, \dots, y_n) = \frac{f(y_1, \dots, y_n | \mu, \sigma^2) f(\mu, \sigma^2)}{f(y_1, \dots, y_n)}$$

distribution of  $y$ 's given  $\mu$  and  $\sigma^2$

distribution of  $\mu$  and  $\sigma^2$

distribution of  $\mu$  and  $\sigma^2$  given  $y$ 's

marginal distribution of  $y$ 's

$A \rightarrow y_1, \dots, y_n \quad B \rightarrow \mu, \sigma^2$

# Prior

$$f(\mu, \sigma^2 | y_1, \dots, y_n) = \frac{f(y_1, \dots, y_n | \mu, \sigma^2) f(\mu, \sigma^2)}{f(y_1, \dots, y_n)}$$

We have  $f(y_1, \dots, y_n | \mu, \sigma^2)$ , the PDF of random variables given  $\mu$ , and  $\sigma^2$ .

We need  $f(\mu, \sigma^2)$ , the PDF of the parameters.

Given  $f(\mu, \sigma^2)$ , we can get  $f(y_1, \dots, y_n)$  by integration

$$f(y_1, \dots, y_n) = \int_{\sigma^2=0}^{\infty} \int_{\mu=-\infty}^{\infty} f(y_1, \dots, y_n | \mu, \sigma^2) f(\mu, \sigma^2) d\mu d\sigma^2.$$

(but it is just a proportionality constant often neglected).

$$f(\mu, \sigma^2 | y_1, \dots, y_n) \propto f(y_1, \dots, y_n | \mu, \sigma^2) f(\mu, \sigma^2)$$

# Prior

The distribution  $f(\mu, \sigma^2)$  is called the prior distribution.

It is arrived at by quantifying expert opinion or using previous data.

There is a way of generating a distributional form for a prior distribution

then all we need are its parameters.

## Prior

Although any distribution that depends on certain parameters  $\theta$

can be used as a prior distribution, we can obtain a “nice” one

called a natural conjugate prior distribution.

Then all we need to do is assess the parameters  $\theta$  for this distribution

either by expert opinion or from previous data.

# Prior

A common joint distribution for the mean  $\mu$  and variance  $\sigma^2$  when data is normal is the natural conjugate prior distribution,

$$f(\mu, \sigma^2) = f(\mu | \sigma^2) f(\sigma^2)$$

$$f(\mu | \sigma^2) = (2\pi\sigma^2 / \alpha)^{-1/2} e^{-\frac{(\mu - \mu_0)^2}{2\sigma^2/\alpha}}$$

$$f(\sigma^2) = \frac{\kappa^{\frac{\nu-2}{2}} (\sigma^2)^{-\frac{\nu}{2}}}{\Gamma\left(\frac{\nu-2}{2}\right) 2^{(\nu-2)/2}} e^{-\frac{\kappa}{2\sigma^2}}$$

←  
inverse gamma distribution

$$\mu_0, \alpha, \nu, \kappa$$

Need to be assessed.

# Prior

The hyperparameters  $(\mu_0, \alpha, \nu, \kappa)$  need to be assessed.

One way is from previous similar study data:

Parameters of prior are called hyperparameters.

i.e.  $n_0$  observations with sample mean  $\bar{y}_0$  and

sample variance  $s_0^2$  use

$$\mu_0 = \bar{y}_0 \quad \nu = n_0 + 1$$

$$\alpha = n_0 \quad \kappa = (n_0 - 1)s_0^2 .$$

# Prior

The likelihood of the previous observations is

$$f(y_1, \dots, y_{n_0} | \mu, \sigma^2) = (2\pi\sigma^2)^{-n_0/2} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^{n_0} (y_i - \mu)^2}$$

$$f(\bar{y}_0, s_0^2 | \mu, \sigma^2) = \left( \frac{n_0}{2\pi\sigma^2} \right)^{\frac{1}{2}} e^{-\frac{n_0}{2\sigma^2} (\mu - \bar{y}_0)^2} \cdot \frac{1}{(2\pi\sigma^2 / n_0)^{(n_0-1)/2}} e^{-\frac{(n_0-1)}{2\sigma^2} s_0^2}.$$

$$\mu_0 = \bar{y}_0$$

$$\alpha = n_0$$

$$\nu = n_0 + 1$$

$$\kappa = (n_0 - 1)s_0^2$$

$$f(\mu | \sigma^2) = \frac{1}{(2\pi\sigma^2/\alpha)^{\frac{1}{2}}} e^{-\frac{\alpha(\mu - \mu_0)^2}{2\sigma^2}}$$

$$f(\sigma^2) = \frac{\kappa^{\frac{\nu-2}{2}} (\sigma^2)^{-\frac{\nu}{2}}}{\Gamma\left(\frac{\nu-2}{2}\right) 2^{(\nu-2)/2}} e^{-\frac{\kappa}{2\sigma^2}}$$

# Posterior

We can now form the posterior distribution

$$f(\overbrace{\mu, \sigma^2}^B | \overbrace{y_1, \dots, y_n}^A) = \frac{f(\overbrace{y_1, \dots, y_n}^A | \overbrace{\mu, \sigma^2}^B) f(\overbrace{\mu, \sigma^2}^B)}{f(\overbrace{y_1, \dots, y_n}^A)}$$

$$f(\overbrace{\mu, \sigma^2}^B) = f(\mu | \sigma^2) f(\sigma^2)$$

$$f(\overbrace{y_1, \dots, y_n}^A | \overbrace{\mu, \sigma^2}^B) = (2\pi\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2}$$

$$f(\overbrace{\mu, \sigma^2}^B) = (2\pi\sigma^2/\alpha)^{-1/2} e^{-\frac{(\mu - \mu_0)^2}{2\sigma^2/\alpha}} \frac{\kappa^{\frac{\nu-2}{2}} (\sigma^2)^{-\frac{\nu}{2}}}{\Gamma(\frac{\nu-2}{2}) 2^{(\nu-2)/2}} e^{-\frac{\kappa}{2\sigma^2}}$$

# Posterior

We can neglect  $f(y_1, \dots, y_n)$  since doesn't have  $(\mu, \sigma^2)$   
and other constants

$$f(\overbrace{\mu, \sigma^2}^B | \overbrace{y_1, \dots, y_n}^A) = \frac{f(\overbrace{y_1, \dots, y_n}^A | \overbrace{\mu, \sigma^2}^B) f(\overbrace{\mu, \sigma^2}^B)}{f(\overbrace{y_1, \dots, y_n}^A)}$$

$$f(\overbrace{\mu, \sigma^2}^B | \overbrace{y_1, \dots, y_n}^A) \propto f(\overbrace{y_1, \dots, y_n}^A | \overbrace{\mu, \sigma^2}^B) f(\overbrace{\mu, \sigma^2}^B)$$

$$f(\overbrace{\mu, \sigma^2}^B | \overbrace{y_1, \dots, y_n}^A) \propto (\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2} (\sigma^2)^{-\frac{1}{2}} e^{-\frac{(\mu - \mu_0)^2}{2\sigma^2/\alpha}} (\sigma^2)^{-\frac{\nu}{2}} e^{-\frac{\kappa}{2\sigma^2}}$$

$$f(\overbrace{\mu, \sigma^2}^B | \overbrace{y_1, \dots, y_n}^A) \propto (\sigma^2)^{-\frac{(n+\nu+1)}{2}} e^{-\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (y_i - \mu)^2 + \alpha(\mu - \mu_0)^2 + \kappa \right]}$$

# Posterior

Now that we have a distribution  $f(\mu, \sigma^2 | y_1, \dots, y_n)$ , we need to estimate the  $(\mu, \sigma^2)$  parameters from it.

We can obtain (marginal) means

$$E(\mu | y_1, \dots, y_n) = \int_{\mu} \int_{\sigma^2} f(\mu, \sigma^2 | y_1, \dots, y_n) d\sigma^2 d\mu$$

$$E(\sigma^2 | y_1, \dots, y_n) = \int_{\sigma^2} \int_{\mu} f(\mu, \sigma^2 | y_1, \dots, y_n) d\mu d\sigma^2$$

or modes

$$\left. \frac{\partial f(\mu, \sigma^2 | y_1, \dots, y_n)}{\partial \mu} \right|_{\hat{\mu}, \hat{\sigma}^2} = 0 \quad \text{and} \quad \left. \frac{\partial f(\mu, \sigma^2 | y_1, \dots, y_n)}{\partial \sigma^2} \right|_{\hat{\mu}, \hat{\sigma}^2} = 0 .$$

# Estimation

The  $(\mu, \sigma^2)$  that maximize the posterior distribution are maximum *a posteriori* (MAP) estimates

$$f(\mu, \sigma^2 | y's, \mu_0, \alpha, \nu, \kappa) = C(\sigma^2)^{-\frac{(\nu+n+1)}{2}} e^{-\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (y_i - \mu)^2 + \alpha(\mu - \mu_0)^2 + \kappa \right]}$$

$$\ln(f(\mu, \sigma^2 | y's, \mu_0, \alpha, \nu, \kappa)) = -\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (y_i - \mu)^2 + \alpha(\mu - \mu_0)^2 + \kappa \right] - \frac{(\nu+n+1)}{2} \ln(\sigma^2) + C$$



does not depend on  $\mu$  or  $\sigma^2$

$$LP = \ln(f(\mu, \sigma^2 | y's, \mu_0, \alpha, \nu, \kappa))$$

# Estimation

Maximum *a posteriori* (MAP) estimates

$$LP(\mu, \sigma^2) = -\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (y_i - \mu)^2 + (\mu - \mu_0)^2 + \kappa \right] - \frac{(\nu+n+1)}{2} \ln(\sigma^2) + C$$

$$\left. \frac{\partial LP(\mu, \sigma^2)}{\partial \mu} \right|_{\hat{\mu}, \hat{\sigma}^2} = -\frac{1}{2\hat{\sigma}^2} \left[ \sum_{i=1}^n 2(y_i - \hat{\mu})(-1) + 2\alpha(\hat{\mu} - \mu_0) \right] = 0$$

$$\left. \frac{\partial LP(\mu, \sigma^2)}{\partial \sigma^2} \right|_{\hat{\mu}, \hat{\sigma}^2} = -\frac{\nu + n + 1}{2} \frac{2}{\hat{\sigma}^2} - \frac{-1}{2(\hat{\sigma}^2)^2} \left[ \sum_{i=1}^n (y_i - \hat{\mu})^2 + \alpha(\hat{\mu} - \mu_0)^2 + \kappa \right] = 0$$

# Estimation

Solving for  $\mu$  and  $\sigma^2$  yields MAP estimates

$$\frac{\partial \text{LP}(\mu, \sigma^2)}{\partial \mu} \Bigg|_{\hat{\mu}, \hat{\sigma}^2} = -\frac{1}{2\hat{\sigma}^2} \left[ \sum_{i=1}^n 2(y_i - \hat{\mu})(-1) + 2\alpha(\hat{\mu} - \mu_0) \right] = 0$$

$$\frac{\partial \text{LP}(\mu, \sigma^2)}{\partial \sigma^2} \Bigg|_{\hat{\mu}, \hat{\sigma}^2} = -\frac{\nu + n + 1}{2} \frac{2}{\hat{\sigma}^2} - \frac{-1}{2(\hat{\sigma}^2)^2} \left[ \sum_{i=1}^n (y_i - \hat{\mu})^2 + \alpha(\hat{\mu} - \mu_0)^2 + \kappa \right] = 0$$

$$\hat{\mu} = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0 \quad \leftarrow \text{Convex combination of prior and likelihood means.}$$

$$\hat{\sigma}^2 = \frac{1}{\nu + n + 1} \left[ \sum_{i=1}^n (y_i - \hat{\mu})^2 + \alpha(\hat{\mu} - \mu_0)^2 + \kappa \right] \quad \leftarrow \text{Can simplify with algebra}$$

# Estimation

Marginal Posterior mean (MPM) for  $\mu$

$$f(\mu | \cdot) \propto \left[ 1 + \frac{1}{u} \left( \frac{\mu - \hat{\mu}}{\tau} \right)^2 \right]^{-(u+1)/2}$$

$$f(\mu | \cdot) = \frac{\Gamma\left(\frac{u+1}{2}\right)}{\Gamma\left(\frac{u}{2}\right)} \frac{(\tau^2)^{-1/2}}{(u\pi)^{1/2}} \frac{1}{\left[ 1 + \frac{1}{u} \left( \frac{\mu - \hat{\mu}}{\tau} \right)^2 \right]^{(u+1)/2}}$$

$$E(\mu | \cdot) = \int_{\mu} \mu f(\mu | \cdot) d\mu = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0$$

$$\hat{\mu} = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0$$

$$u = n + v - 2$$

$$\tau^2 = \frac{\alpha \mu_0^2 + \kappa + n \bar{y}^2 - \frac{(n \bar{y} + \alpha \mu_0)^2}{n + \alpha}}{(n + \alpha)(n + v - 2)}$$

Scaled Student-t

# Estimation

Marginal Posterior mean (MPM) for  $\mu$

$$f(\mu | \cdot) \propto \left[ (n + \alpha)(\mu - \mu_0)^2 + \alpha\mu_0^2 + \kappa + n\bar{y}^2 - \frac{(n\bar{y} + \alpha\mu_0)^2}{n + \alpha} \right]^{-\alpha}$$

$$f(\mu | \cdot) \propto \left[ 1 + \frac{1}{u} \left( \frac{\mu - \hat{\mu}}{\tau} \right)^2 \right]^{-(u+1)/2}$$

$$\hat{\mu} = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0$$

$$u = n + v - 2$$

$$\tau^2 = \frac{\alpha\mu_0^2 + \kappa + n\bar{y}^2 - \frac{(n\bar{y} + \alpha\mu_0)^2}{n + \alpha}}{(n + \alpha)(n + v - 2)}$$

# Estimation

Marginal Posterior mean (MPM) for  $\mu$

$$f(\mu | \cdot) \propto \left[ 1 + \frac{1}{u} \left( \frac{\mu - \hat{\mu}}{\tau} \right)^2 \right]^{-(u+1)/2}$$

$$f(\mu | \cdot) = \frac{\Gamma\left(\frac{u+1}{2}\right)}{\Gamma\left(\frac{u}{2}\right)} (\tau^2)^{-1/2} \frac{1}{(u\pi)^{1/2} \left[ 1 + \frac{1}{u} \left( \frac{\mu - \hat{\mu}}{\tau} \right)^2 \right]^{(u+1)/2}}$$

$$E(\mu | \cdot) = \int_{\mu} \mu f(\mu | \cdot) d\mu = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0$$

Scaled Student-t

$$\hat{\mu} = \frac{n}{\alpha + n} \bar{y} + \frac{\alpha}{\alpha + n} \mu_0$$

$$u = n + v - 2$$

$$\tau^2 = \frac{\alpha \mu_0^2 + \kappa + n \bar{y}^2 - \frac{(n \bar{y} + \alpha \mu_0)^2}{n + \alpha}}{(n + \alpha)(n + v - 2)}$$

# Student-t Distribution

The Student- $t$  Distribution, can be generalized to have location and scale parameters, so that  $x \sim t(\nu, \delta, \tau)$  if

$$f(x|\nu, \delta, \tau) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)} \frac{(\tau^2)^{-1/2}}{\sqrt{\nu\pi}} \frac{1}{\left(1 + \frac{1}{\nu} \left(\frac{x-\delta}{\tau}\right)^2\right)^{(\nu+1)/2}} \quad x \in \mathbb{R}$$

where  $\nu = 1, 2, \dots$ .

$$E(x) = \delta \quad \nu > 1 \quad \text{var}(x) = \frac{\nu}{\nu - 2} \tau^2 \quad \nu > 2$$

# Marginal Posterior Mean Estimation

Posterior mean (MPM) for  $\sigma^2$

$$f(\mu, \sigma^2 | \cdot) = C(\sigma^2)^{-\frac{(v+n+1)}{2}} e^{-\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (y_i - \mu)^2 + \alpha(\mu - \mu_0)^2 + \kappa \right]} \quad \delta^2 = \frac{2\sigma^2}{n + \alpha}$$

$$f(\mu, \sigma^2 | \cdot) \propto (\sigma^2)^{-\frac{(n+v+1)}{2}} e^{-\frac{1}{2\sigma^2} \left[ \left( \frac{n+\alpha}{2} \right) (\mu - \hat{\mu})^2 + g \right]}$$

$$f(\sigma^2 | \cdot) \propto (\sigma^2)^{-\frac{(n+v)}{2}} e^{-\frac{g}{2\sigma^2}} (2\pi\delta^2)^{1/2} \int_{\mu} (2\pi\delta^2)^{-1/2} e^{-\frac{(\mu - \hat{\mu})^2}{2\delta^2}} d\mu$$

$$f(\sigma^2 | \cdot) = \frac{(g/2)^{(n+v-2)/2}}{\Gamma\left(\frac{n+v-2}{2}\right)} (\sigma^2)^{-\frac{(n+v-2)}{2}-1} e^{-\frac{g}{2\sigma^2}} \quad \begin{matrix} 1 \\ \leftarrow \\ \text{Inverse Gamma} \end{matrix}$$

$$g = n\bar{y}^2 - (n + \alpha)\hat{\mu}^2 + \alpha\mu_0^2 + \kappa$$

# Marginal Posterior Mean Estimation

## Marginal Posterior mean (MPM)

$$f(\sigma^2 | \cdot) = \frac{(g/2)^{(n+\nu-2)/2}}{\Gamma\left(\frac{n+\nu-2}{2}\right)} (\sigma^2)^{-\frac{(n+\nu-2)}{2}-1} e^{-\frac{g}{2\sigma^2}}$$

$$E(\sigma^2 | \cdot) = \frac{g}{n + \nu - 4}$$

$$\text{var}(\sigma^2 | \cdot) = \frac{(g/2)^2}{(\gamma-1)^2(\gamma-2)}$$

$$\gamma = \frac{n + \nu - 2}{2}$$

$$g = n\bar{y}^2 - (n + \alpha)\hat{\mu}^2 + \alpha\mu_0^2 + \kappa$$

# Discussion

Can do any Statistical Model Bayesian:

Bayesian Regression

Bayesian Time Series

Bayesian ANOVA

Bayesian Classification

Bayesian Multivariate Regression

Bayesian Image Reconstruction

## Discussion

# Questions?

## Homework 12

1. You talk to an expert on MU undergraduate heights at the University and they help you assess hyperparameters  $\mu_0=65$ ,  $\alpha=10$ ,  $\nu=11$ ,  $\kappa=36$ . You take a random sample of  $n=35$  heights to get  $\bar{y}=68.3$ ,  $\bar{y}^2=4675.4$ , and  $s=3.48$ .
  - a. Calculate the exact marginal posterior means and variances of  $\mu$  and  $\sigma^2$ .
  - b. Calculate the marginal posterior means and variances via numerical rectangle integration.
  - c. Calculate the marginal posterior mean height via a simulation integration technique.

# Homework 12

2. Now use a Laplace (double exponential) prior distribution for the mean and still an inverse gamma distribution for the variance.

(You do not have conjugate priors now.) Pick your own  $\lambda$ .

- Calculate the marginal posterior mean height via Numerical rectangle integration.
- Calculate the marginal posterior mean height via a simulation integration technique.

$$f(y_1, \dots, y_n | \mu, \sigma^2) = (2\pi\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2}$$

$$f(\mu | \sigma^2) = \frac{\lambda}{4\sigma^2} e^{-\frac{\lambda|\mu - \mu_0|}{2\sigma^2}}$$

$$f(\sigma^2) = \frac{\kappa^{\frac{\nu-2}{2}} (\sigma^2)^{-\frac{\nu}{2}}}{\Gamma\left(\frac{\nu-2}{2}\right) 2^{(\nu-2)/2}} e^{-\frac{\kappa}{2\sigma^2}}$$

\*If you'd like you can use  
 $|\mu - \mu_0| = \sqrt{(\mu - \mu_0)^2 + .001}$

LASSO