## Chapter 11: Survival Analysis

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Time to Event
Survival analysis is the statistical analysis of time-to-event variables.
The event could be a heart attack, cancer remission, or death.

What is the probability that a participant survives 5 years?

Are there differences in survival between groups?

How do certain characteristics affect participants chances of survival?

Time to Event
Not all participants enroll when a study begins, so when a study ends, not all participants were enrolled for the same amount of time.

True survival time (failure time) is not known because the study ended before the event or participants dropped out.

The last observed follow-up is called the censored or censoring time.

Right censoring is when a participant does not have the event of interest during the study, last observed follow-up is less than the time to event.

Time to Event

## Patients experiences with Myocardial Infarction over 10 years



Some join up to 2 years after start, all are followed for 10 years from start.
3 Myocardial Infarction
1 death
2 drop out
4 completions


All enrolled at the same time, all are followed for 10 years from start.
3 Myocardial Infarction
1 death
2 drop out
4 completions


All enrolled at the same time, all are followed for 10 years from start.
3 Myocardial Infarction
1 death
2 drop out
4 completions

Survival Analysis analyzes not only the number of MI events, but also times.

### 11.1 Introduction to Survival Data

Survival analysis measures two pieces of information

1) Whether the event occurred, $1=y e s, 0=$ no
2) Last follow-up time, from enrollment.

The survival function is the probability a person survives past a time $t$.


$$
\begin{aligned}
& t=0.0: \text { survival probability }=1.00 \\
& t=2.0 \text { : survival probability }=0.83 \\
& t=8.5: \text { survival probability }=0.50 \text { (Median) } \\
& t=10.0: \text { survival probability }=0.47
\end{aligned}
$$

### 11.2 Estimating the Survival Function

There are several parametric and nonparametric ways to estimate survival Let's examine nonparametric step survival curves. Time on $x$ axis and survival (percentage) at risk on $y$ axis.


$$
\begin{aligned}
& t=0.0: \text { survival probability }=1.00 \\
& t=2.0 \text { : survival probability }=0.90 \\
& t=9.0 \text { : survival probability }=0.50 \text { (Median) } \\
& t=10.0 \text { : survival probability }=0.45
\end{aligned}
$$

### 11.2 Estimating the Survival Function

Example of 24 year study with 20 participants. Some die, many drop out, few finish.

### 11.2 Estimating the Survival Function - Actuarial

We can organize the data into a simple table.
Divide the 24 year study into 5 year intervals.
$0-4$ years
5-9 years
10-14 years
15-19 years
20-24 years.


4
14

### 11.2 Estimating the Survival Function - Actuarial


11.2 Estimating the Survival Function - Actuarial

|  | Participant | Year of Death | Last Contact |
| :---: | :---: | :---: | :---: |
| $0-4$ years | 14 | 1 |  |
|  | 8 |  | 2 |
|  | 2 | 3 |  |
| 5-9 years | 18 | 5 |  |
|  | 17 |  | 6 |
|  | 19 |  | 9 |
| 10-14 years | 15 |  | 10 |
|  | 3 |  | 11 |
|  | 13 |  | 12 |
|  | 6 |  | 13 |
|  | 7 | 14 |  |
| 15-19 years | 10 |  | 17 |
|  | 20 | 17 |  |
|  | 9 |  | 18 |
|  | 4 |  | 19 |
| 20-24 years | 12 |  | 21 |
|  | 16 | 23 |  |
|  | 1 |  | 24 |
|  | 5 |  | 24 |
|  | 11 |  | 24 |
| Time Data |  |  |  |

11.2 Estimating the Survival Function - Actuarial

We can organize the data into a simple table. Divide the 24 year study into 5 year intervals.

Count Alive at beginning of each interval. Count how many Deaths during interval. Number censored (dropped out) in each interval.

| Number <br> Interval <br> in Years | Number <br> Beginning <br> of Interval | of Deaths <br> During <br> Interval | Number <br> Censored |
| :---: | :---: | :---: | :---: |
| $0-4$ | 20 | 2 | 1 |
| $5-9$ | 17 | 1 | 2 |
| $10-14$ | 14 | 1 | 4 |
| $15-19$ | 9 | 1 | 3 |
| $20-24$ | 5 | 1 | 4 |


|  | Participant | Year of Death | Last Contact |
| :---: | :---: | :---: | :---: |
|  | 14 | 1 |  |
| $0-4$ years | 8 |  | 2 |
|  | 2 | 3 |  |
| 5-9 years | 18 | 5 |  |
|  | 17 |  | 6 |
|  | 19 |  | 9 |
|  | 15 |  | 10 |
|  | 3 |  | 11 |
| 10-14 years | 13 |  | 12 |
|  | 6 |  | 13 |
|  | 7 | 14 |  |
|  | 10 |  | 17 |
|  | 20 | 17 |  |
| 15-19 years | 9 |  | 18 |
|  | 4 |  | 19 |
|  | 12 |  | 21 |
|  | 16 | 23 |  |
| 20-24 years | 1 |  | 24 |
|  | 5 |  | 24 |
|  | 11 |  | 24 |
| Time Data |  |  |  |

## Life Tables (actuarial tables)

$N_{t}=$ number event free during interval $t$ (Number at risk)
$D_{t}=$ number who die during interval $t$
$C_{t}=$ number censored during interval $t$
$N_{t^{*}}=$ average number at risk during interval $t$

Deaths assumed to occur at end of the interval.
 $N_{t^{*}}=N_{t}-C_{t} / 2$

## Life Tables (actuarial tables)

$N_{t}=$ number event free during interval $t$ (Number at risk)
$D_{t}=$ number who die during interval $t$
$C_{t}=$ number censored during interval $t$
$N_{t^{*}}=$ average number at risk during interval $t$ $N_{t^{*}}=N_{t}-C_{t} / 2$
$q_{t}=$ prop. die in interval $t, q_{t}=D_{t} / N_{t^{*}}$


Time Data
$p_{t}=$ prop. survive in interval $t, p_{t}=1-q_{t}$
$S_{t}=$ prop. survive past interval $t, S_{t+1}=p_{t+1} S_{t}$
11.2 Estimating the Survival Function - Actuarial


# 11.2 Estimating the Survival Function - Actuarial 

| Interval in Years | Number at Risk During Interval, $N_{t}$ | Average Number at Risk During Interval, $\boldsymbol{N}_{\mathrm{t}}$. | Number of Deaths During Interval, $D_{t}$ |
| :---: | :---: | :---: | :---: |
| 0-4 | $20 \quad 20-(1 / 2)=19.5$ |  | 2 |
| 5-9 | 17 17-[2/2)=16.0 |  | 1 |
|  | $N_{t^{*}}=N_{t}-C_{t} / 2$ |  |  |
| Interval in Years | Number <br> Alive at Beginning of Interva | Number of Deaths During Interval | Number Censored |
| 0-4 | 20 | 2 | 1 |
| 5-9 | 17 | 1 | 2 |
| 10-14 | 14 | 1 | 4 |
| 15-19 | 9 | 1 | 3 |
| 20-24 | 5 | 1 | 4 |

# 11.2 Estimating the Survival Function - Actuarial 

| Interval in Years | Number at Risk During Interval, $N_{t}$ | Average Number at Risk During Interval, $N_{t}$. | Number of Deaths During Interval, $D_{t}$ | $\begin{aligned} & \text { Lost to } \\ & \text { Follow-Up, } \\ & C_{t} \end{aligned}$ | Proportion Dying During Interval, $q_{\text {t }}$ |  |  | $0-4$ years | 14 | 1 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | 8 |  | 2 |
|  |  |  |  |  |  | Among Those at Risk, Proportion Surviving Interval, $p_{t}$ | Survival Probability, $S_{t}$ |  | 5-9 years | 2 | 3 |  |
|  |  |  |  |  |  |  |  | 18 |  | 5 |  |
|  |  |  |  |  |  |  |  | 17 |  |  | 6 |
|  |  |  |  |  |  |  |  | 15 |  |  | 10 |
| 0-4 | 20 | 19.5 | 2 | 1 | 0.103 | 0.897 | 0.897 | 10-14 years | 15 |  | 10 |
| 5-9 | 17 | 16.0 | 1 | 2 | 0.063 | 0.937 | 0.840 |  | 13 |  | 12 |
| 10-14 | 14 | 12.0 | 1 | 4 | 0.083 | 0.917 | 0.770 |  | 6 |  | 13 |
| 15-19 | 9 | 7.5 | 1 | 3 | 0.133 | 0.867 | 0.688 |  | 7 | 14 |  |
| 20-24 | 5 | 3.0 | 1 | 4 | 0.333 | 0.667 | 0.446 | 15-19 years | 10 |  | 17 |
|  | $N_{t^{*}}=N_{t}-C_{t} / 2$ |  |  | $q_{t}=D_{t} / N_{t^{*}}$ |  | $p_{t}=1-q_{t}$ | $S_{t+1}=p_{t+1}$ |  | 20 | 17 | 18 |
|  |  |  |  | 4 |  |  |  |  | 19 |
|  | Number | Number | Number Censored |  |  | $N_{t}=\#$ event free during interval $t$ <br> (Number at risk) |  |  |  | 20-24 years | 12 |  | 21 |
|  | Alive at Beginning | of Deaths During |  | 16 | 23 |  |  |  |  |  |
| $\begin{aligned} & \text { Interval } \\ & \text { in Years } \end{aligned}$ | Beginning of Interval | During Interval |  | 16 1 | 23 |  |  |  |  | 24 |
| 0-4 | 20 | 2 | 1 | $D_{t}=\#$ who die during interval $t$ |  |  |  |  | 5 |  |  | 24 |
| $5-9$ $10-14$ | 17 | 1 | 2 |  |  |  |  |  | 11 |  |  | 24 |
| $10-14$ $15-19$ | 14 9 | 1 | 3 | $C_{t}=\#$ censored during interval $t$ |  |  |  |  | Time Data |  |  |
| 20-24 | 5 | 1 | 4 |  |  |  | rvalt |  |  |  |  |  |
|  |  |  |  | $N_{t^{*}}=$ avg. \# at risk during interval $t, N_{t^{*}}=N_{t}-C_{t} / 2$ |  |  |  |  |  |  |  |
|  |  |  |  | $q_{t}=$ prop. die in interval $t, q_{t}=D_{t} / N_{t^{*}}$ |  |  |  |  |  |  |  |
|  |  |  |  | $p_{t}=$ prop. survive in interval $t, p_{t}=1-q_{t}$ |  |  |  |  |  |  |  |
|  |  |  |  | $S_{t}=$ prop. survive past interval $t, S_{t+1}=p_{t+1} S_{t}$ |  |  |  |  |  |  |  |

11.2 Estimating the Survival Function - Actuarial

Participant Year of Death Last Contact


### 11.2 Estimating the Survival Function - Kaplan-Meier

Kaplan-Meier Survival Curve approach re-estimates the probability each time an event occurs. Re-estimates every death or censoring.

Assumes censoring is independent of the likelihood of developing the event of interest. You don't drop out because you don't think you will ever get the event or because you know you will get it. You drop out because you are too busy or move.

Survival probabilities are comparable in participants who are recruited earlier as well as later. How participants are recruited doesn't change.
11.2 Estimating the Survival Function - Kaplan-Meier


## Participant Year of Death Last Contact

| Participant | Year of Death | Last Contact |
| :---: | :---: | :---: |
| 14 | 1 |  |
| 8 |  | 2 |
| 2 | 3 |  |
| 18 | 5 |  |
| 17 |  | 6 |
| 19 |  | 9 |
| 15 |  | 10 |
| 3 |  | 11 |
| 13 |  | 12 |
| 6 |  | 13 |
| 7 | 14 |  |
| 10 |  | 17 |
| 20 | 17 |  |
| 9 |  | 18 |
| 4 |  | 19 |
| 12 |  | 21 |
| 16 | 23 |  |
| 1 |  | 24 |
| 5 |  | 24 |
| 11 |  | 24 |
| Time Data |  |  |
| $S=S N_{t}-D_{t}$ |  |  |

11.2 Estimating the Survival Function - Kaplan-Meier

| Time, years | Survival Probability, $S_{t+1}=S_{t} \times\left[\left(N_{t+1}-D_{t+1}\right] / N_{t+1}\right]$ |
| :---: | :---: |
| 0 | $1^{+}$ |
| 1 | $1 \times[(20-1) / 20]=0.950$ |
| 2 | $0.950 \times[[19-0] / 19]=0.950$ |
| 3 | $0.950 \times[(18-1] / 18]=0.897$ |
| 5 | $0.897 \times[[17-1] / 17]=0.844$ |
| 6 | 0.844 |
| 9 | 0.844 |
| 10 | 0.844 |
| 11 | 0.844 |
| 12 | 0.844 |
| 13 | 0.844 |
| 14 | 0.760 |
| 17 | 0.676 |
| 18 | 0.676 |
| 19 | 0.676 |
| 21 | 0.676 |
| 23 | 0.507 |
| 24 | 0.507 |



Participant Year of Death Last Contact


## Biostatistical Methods

11.2 Estimating the Survival Function - Kaplan-Meier

Survival
Participant Year of Death Last Contact



### 11.2 Estimating the Survival Function - Kaplan-Meier Survival




Participant Year of Death Last Contact

| Participant | Year of Death | Last Contact |
| :---: | :---: | :---: |
| 14 | 1 |  |
| 8 |  | 2 |
| 2 | 3 |  |
| 18 | 5 |  |
| 17 |  | 6 |
| 19 |  | 9 |
| 15 |  | 10 |
| 3 |  | 11 |
| 13 |  | 12 |
| 6 |  | 13 |
| 7 | 14 |  |
| 10 |  | 17 |
| 20 | 17 |  |
| 9 |  | 18 |
| 4 |  | 19 |
| 12 |  | 21 |
| 16 | 23 |  |
| 1 |  | 24 |
| 5 |  | 24 |
| 11 |  | 24 |
| Time Data |  |  |
| $S_{t+1}=S_{t} \frac{\mid \mathbf{V}_{t}-\boldsymbol{D}_{t}}{\boldsymbol{N} t}$ |  |  |

11.2 Estimating the Survival Function - Kaplan-Meier

Time, Probability, Failure Probability,


Participant Year of Death Last Contact


### 11.3 Comparing Survival Curves

There are methods for comparing equivalence of survival curves.
An example is one survival curve for a group receiving a medication and another survival curve for another group receiving a placebo.

We might be comparing survival curves for men vs. women or between two demographic groups.

Here present version of log-rank test statistic linked to $\chi^{2}$ test. Compares observed events to expected events at each time point.

### 11.3 Comparing Survival Curves

Example: Small clinical trial to compare chemo Before vs. After surgery.
Chemotherapy Before Surgery

| Month of Death | Month of Last Contact |  | Month of Death |
| :---: | :---: | :---: | :---: |
|  | 8 | 33 | Month of Last Contact |
| 8 | 32 | 28 | 48 |
| 26 | 20 | 41 | 48 |
| 14 | 40 |  | 25 |
| 21 |  | 37 |  |
| 27 |  | 48 |  |
|  |  | 25 |  |

We can perform a hypothesis test to see if the two treatments result in equivalent outcomes.

### 11.3 Comparing Survival Curves

Example: We can perform a hypothesis test for equivalence.
Chemo Before Surgery

| Time, <br> months | Number at <br> Risk, $N_{t}$ | Number of <br> Deaths, $D_{t}$ | Number <br> Censored, $C_{t}$ | Survival Probability, <br> $\left.S_{t+1}=S_{t} \times\\| \\|_{t+1}-D_{t+1} / N_{t+1}\right]$ | Time, <br> months | Number at <br> Risk, $N_{t}$ | Number of <br> Deaths, $D_{t}$ | Number <br> Censored, $C_{t}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | | Survival Probability, |
| :---: |
| $S_{t+1}=S_{t} \times \\| N_{t+1} \times D_{t+1} / 1 / N_{t+1} \mid$ |

Plot the survival curves.

### 11.3 Comparing Survival Curves

Example: We can perform a hypothesis test for equivalence.
Step 1: Hypotheses and significance. $\alpha=0.05 \quad \sum_{t=1}^{T} O_{i j}=$ Observed Deaths in Group $j$ $\mathrm{H}_{0}$ : The two survival curves are identical. $\mathrm{H}_{1}$ : The two survival curves are not identical.
$\sum_{t=1}^{T} E_{i j}=$ Expected Deaths in Group $j$
Step 2: Test Statistic (log-rank test)

$$
\chi^{2}=\sum_{j=1}^{2}\left(\sum_{t=1}^{T} O_{i j}-\sum_{t=1}^{T} E_{i j}\right)^{2} / \sum_{t=1}^{T} E_{i j}
$$

$$
d f=k-1
$$

Step 3: Decision Rule Reject if $\chi^{2}>\chi^{2}{ }_{\alpha, \text { df }}=3.84$.
Step 4: Compute Test Statistic Next slide.



### 11.3 Comparing Survival Curves

## Step 4: Compute Test Statistic

|  |  |  |  |  | tatis |  | Expected | Expected |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Death Time, months | Number at Risk in Group 1, $N_{1 t}$ | Number at Risk in Group 2, $N_{2 t}$ | Total Number at Risk, $N_{t}$ | Number of Events in Group 1, $0_{1 t}$ | Number of Events in Group 2, $\mathrm{O}_{2 t}$ | Total Number of Events, 0 t | Events in Group 1, $E_{1}$ $=N_{1 t} \times\left[0 / N_{t}\right]$ | $\begin{gathered} \text { Events in } \\ \text { Group 2, } \\ =E_{2 t}=\left[N_{2 t} \times\left(N_{t}\right)\right. \end{gathered}$ |
| 8 | 10 | 10 | 20 | 1 | 0 | 1 | 0.500 | 0.500 |
| 12 | 8 | 10 | 18 | 1 | 0 | 1 | 0.444 | 0.556 |
| 14 | 7 | 10 | 17 | 1 | 0 | 1 | 0.412 | 0.588 |
| 21 | 5 | 10 | 15 | 1 | 0 | 1 | 0.333 | 0.667 |
| 26 | 4 | 8 | 12 | 1 | 0 | 1 | 0.333 | 0.667 |
| 27 | 3 | 8 | 11 | 1 | 0 | 1 | 0.273 | 0.727 |
| 28 | 2 | 8 | 10 | 0 | 1 | 1 | 0.200 | 0.800 |
| 33 | 1 | 7 | 8 | 0 | 1 | 1 | 0.125 | 0.875 |
| 41 | 0 | 5 | 5 | 0 | 1 | 1 | 0.000 | 1.000 |
| $\begin{aligned} & \chi^{2}=\sum_{j=1}^{2} \frac{\left(\sum_{t=1}^{T} O_{i j}-\sum_{t=1}^{T} E_{i j}\right)^{2}}{\sum_{t=1}^{T} E_{i j}}=\frac{(6-2.620)^{2}}{2.620}+\frac{(3-6.380)^{2}}{6.380} \\ & \chi^{2}=4.360+1.791=6.151 \end{aligned}$ |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |

Expected Expected Number of Number of Events in Events in Group 2, $E_{2 t}$
$\left(0, / N_{t}\right)$ $\sum_{t=1}^{T} O_{i j}=$ Observed Deaths in Group $j$ $\sum_{t=1}^{T} E_{i j}=$ Expected Deaths in Group $j$


- Chemo before surgery
-*- Chemo after surgery


### 11.3 Comparing Survival Curves

Example: We can perform a hypothesis test for equivalence.
Step 1: Hypotheses and significance. $\alpha=0.05 \quad \sum_{t=1}^{T} O_{j i}=$ Observed Deaths in Group $j$ $\mathrm{H}_{0}$ : The two survival curves are identical. $\mathrm{H}_{1}$ : The two survival curves are not identical.
$\sum_{t=1}^{T} E_{i j}=$ Expected Deaths in Group $j$
Step 2: Test Statistic (log-rank test)

$$
\chi^{2}=\sum_{j=1}^{2}\left(\sum_{t=1}^{T} O_{i j}-\sum_{t=1}^{T} E_{i j}\right)^{2} / \sum_{t=1}^{T} E_{i j}
$$

$$
d f=k-1
$$

Step 3: Decision Rule Reject if $\chi^{2}>\chi^{2}{ }_{\alpha, d f}=3.84$.
Step 4: Compute Test Statistic
$\chi^{2}=6.15$
Step 5: Conclusion


Reject $\mathrm{H}_{0}$ because 6.16>3.84.
$\rightarrow$ Chemo before surgery

-     -         - Chemo after surgery

The survival function is the probability a person survives past a time $t$.

## Actuarial Life Table

$N_{t}=\underset{\text { (Number at risk) }}{\text { \# event }}$ during interval $t$
$D_{t}=\#$ who die in interval $t$
$C_{t}=$ \# censored in interval $t$
$N_{t^{*}}=$ avg. \# at risk in interval $t, N_{t^{*}}=N_{t}-C_{t} / 2$
$q_{t}=$ prop. die in interval $t, q_{t}=D_{t} / N_{t^{*}}$
$p_{t}=$ prop. survive in interval $t, p_{t}=1-q_{t}$
$S_{t}=$ prop. survive past interval $t$
Can plot $S_{t}$ vs. $t$.

## Kaplan-Meier Life Table

$$
S_{t+1}=S_{t} \frac{N_{t}-D_{t}}{N_{t}}
$$

$$
S E\left(S_{t}\right)=S_{t} \sqrt{\sum \frac{D_{t}}{N_{t}\left(N_{t}-D_{t}\right)}}
$$

## Chi-Square Test

$$
\chi^{2}=\sum_{j=1}^{2} \frac{\left(\sum_{t=1}^{T} O_{i j}-\sum_{t=1}^{T} E_{i j}\right)^{2}}{\sum_{t=1}^{T} E_{i j}} d f=k-1
$$

Cox Proportional Hazards Model $h(t)=h_{0}(t) \exp \left(b_{1} x_{1}+b_{2} x_{2}+\ldots+b_{p} x_{p}\right)$

## Questions?

Homework 11

## Read Chapter 11.

Problems 12, 14. (Both interpreting graphs.)



