### Class 18

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## Agenda:

Recap Chapter 9.1

Lecture Chapter 9.2

### Recap Chapter 9.1

### 9: Inferences Involving One Population

### 9.1 Inference about the Mean $\mu$ ( $\sigma$ Unknown)

In Chapter 8, we performed hypothesis tests on the mean by

- 1) assuming that  $\overline{x}$  was normally distributed (n "large"),
- 2) assuming the hypothesized mean  $\mu_0$  were true,
- 3) assuming that  $\sigma$  was known, so that we could form

$$z^* = \frac{\overline{x} - \mu_0}{\sigma / \sqrt{n}}$$
 which with 1) – 3) has standard normal dist.

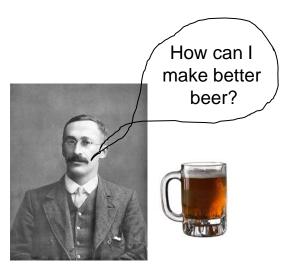
### 9: Inferences Involving One Population 9.1 Inference about the Mean $\mu$ ( $\sigma$ Unknown)

However, in real life, we never know  $\sigma$  for

$$z^* = \frac{\overline{x} - \mu_0}{\sigma / \sqrt{n}}$$

so we would like to estimate  $\sigma$  by s, then use

$$t^* = \frac{\overline{x} - \mu_0}{s / \sqrt{n}} .$$

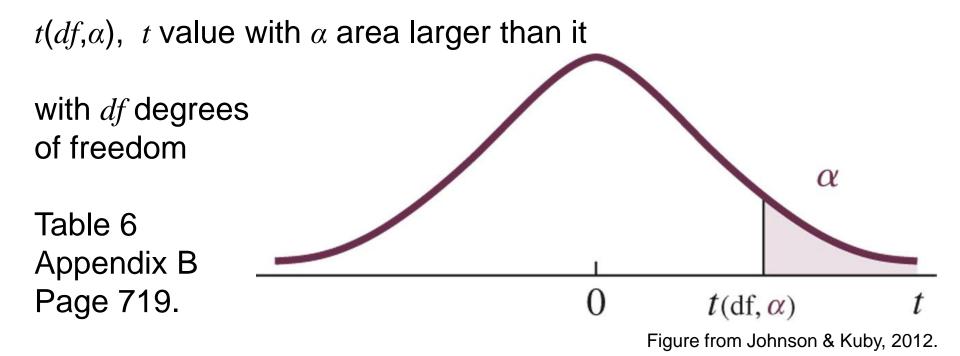


But  $t^*$  does not have a standard normal distribution.

It has what is called a Student t-distribution.  $\leftarrow$  Gosset

# 9: Inferences Involving One Population 9.1 Inference about the Mean $\mu$ ( $\sigma$ Unknown) Using the t-Distribution Table

Finding critical value from a Student *t*-distribution, df=n-1



 $t(df, \alpha)$ 

9: Inferences Involving One Population

9.1 Inference about the Mean  $\mu$  ( $\sigma$  Unknown)

**Example:** Find the value of t(10,0.05),

df=10,  $\alpha$ =0.05.

	0.25	0.10	0.05	0.025	0.01	0.005
Area in Two Tails						
df	0.50	0.20	0.10	0.05	0.02	0.01
3 4 5	0.765 0.741 0.727	1.64 1.53 1.48	2.35 2.13 2.02	3.18 2.78 2.57	4.54 3.75 3.36	5.84 4.60 4.03
6 7 8 9 10	0.718 0.711 0.706 0.703 0.700	1.44 1.41 1.40 1.38 1.37	1.94 1.89 1.86 1.83	2.45 2.36 2.31 2.26 2.23	3.14 3.00 2.90 2.82 2.76	3.71 3.50 3.36 3.25 3.17
Ė						
35 40 50 70 100	0.682 0.681 0.679 0.678 0.677	1.31 1.30 1.30 1.29 1.29	1.69 1.68 1.68 1.67 1.66	2.03 2.02 2.01 1.99 1.98	2.44 2.42 2.40 2.38 2.36	2.72 2.70 2.68 2.65 2.63
df > 100	0.675	1.28	1.65	1.96	2.33	2.58

Table 6
Appendix B
Page 719.

0

Go to 0.05 One Tail column and down to 10 df row.

Figures from Johnson & Kuby, 2012.

### 9: Inferences Involving One Population

### 9.1 Inference about the Mean $\mu$ ( $\sigma$ Unknown)

#### **Confidence Interval Procedure**

Discussed a confidence interval for the  $\mu$  when  $\sigma$  was known,

#### **Confidence Interval for Mean:**

$$\overline{x} - z(\alpha/2) \frac{\sigma}{\sqrt{n}}$$
 to  $\overline{x} + z(\alpha/2) \frac{\sigma}{\sqrt{n}}$  (8.1)

now, with sigma unknown, the CI for the mean is

#### **Confidence Interval for Mean:**

$$\overline{x} - t(df, \alpha/2) \frac{s}{\sqrt{n}}$$
 to  $\overline{x} + t(df, \alpha/2) \frac{s}{\sqrt{n}}$  (9.1)

### 9: Inferences Involving One Population

### 9.1 Inference about the Mean $\mu$ ( $\sigma$ Unknown)

#### Recap 9.1:

Essentially have new critical value,  $t(df,\alpha)$  to look up

in a table when  $\sigma$  is unknown. Used same way as before.

#### σ assumed known

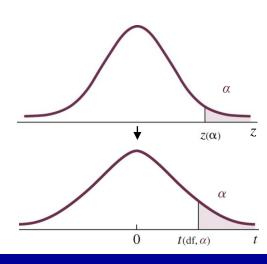
$$\overline{x} \pm z(\alpha/2) \frac{\sigma}{\sqrt{n}}$$

$$z^* = \frac{\overline{x} - \mu_0}{\sigma / \sqrt{n}}$$

#### <u>σ assumed unknown</u>

$$\overline{x} \pm t(df, \alpha/2) \frac{s}{\sqrt{n}}$$

$$t^* = \frac{\overline{x} - \mu_0}{s / \sqrt{n}}$$



### Lecture Chapter 9.2

We talked about a Binomial experiment with two outcomes. Each performance of the experiment is called a trial. Each trial is independent.

$$P(x) = \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x}$$

$$n = 1, 2, 3, ...$$
  
 $0 \le p \le 1$   
 $x = 0, 1, ..., n$ 

n = number of trials or times we repeat the experiment.

x = the number of successes out of n trials.

p =the probability of success on an individual trial.

When we perform a binomial experiment we can estimate the probability of heads as

### **Sample Binomial Probability**

i.e. number of H out of n flips

$$p' = \frac{x}{n} \tag{9.3}$$

where x is the number of successes in n trials.

This is a point estimate. Recall the rule for a CI is point estimate ± some amount

#### **Background**

In Statistics, if we have a random variable x with

$$mean(x) = \mu$$
 and  $variance(x) = \sigma^2$ 

then the mean and variance of cx where c is a constant is

If x has a binomial distribution then

$$mean(cx) = cnp$$
 and  $variance(cx) = c^2 np(1-p)$ .

#### **Background**

With 
$$p' = \frac{x}{n}$$
, the constant is  $c = \frac{1}{n}$ , and

$$\operatorname{mean}\left(\frac{x}{n}\right) = \left(\frac{1}{n}\right)\operatorname{mean}(x) = \left(\frac{1}{n}\right)np = p$$

so the variance of 
$$p' = \frac{x}{n}$$
 is variance  $\left(\frac{x}{n}\right) = \frac{\sigma^2}{n^2} = \frac{p(1-p)}{n}$ 

standard error of 
$$p' = \frac{x}{n}$$
 is  $\sigma_{p'} = \sqrt{\frac{\sigma^2}{n^2}} = \sqrt{\frac{p(1-p)}{n}}$ .

That is where 1. and 2. in the green box below come from

If a random sample of size n is selected from a large population with p=P(success), then the sampling distribution of p' has:

- 1. A mean  $\mu_{p}$  equal to p
- 2. A standard error  $\sigma_{p'}$  equal to  $\sqrt{\frac{p(1-p)}{n}}$
- 3. An approximately normal distribution if *n* is sufficiently "large."

# 6: Normal Probability Distributions 6.5 Normal Approximation of the Binomial Distribution

If we flip the coin a large number of times

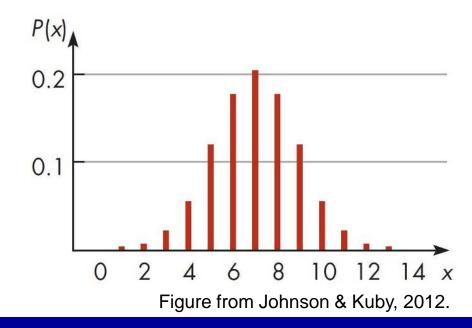
$$P(x) = \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x}$$

x = # of heads when we flip a coin n times

$$n=14$$
  $p=1/2$ 

It gets tedious to find the n=14 probabilities!

$$x = 0, ...., n$$



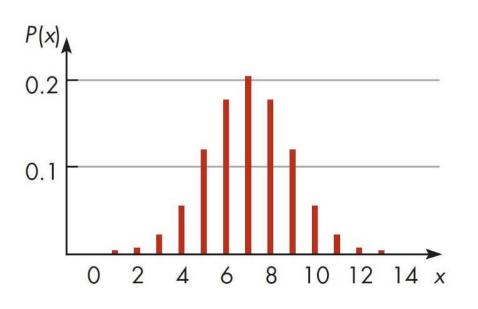
### 6: Normal Probability Distributions

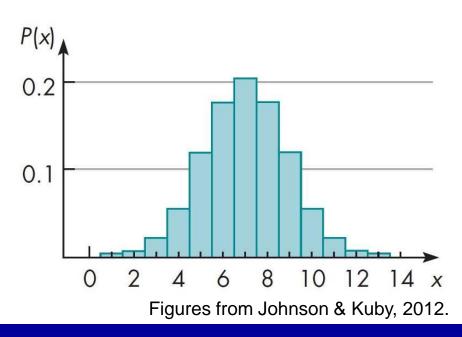
### 6.5 Normal Approximation of the Binomial Distribution

It gets tedious to find the n=14 probabilities!

$$n=14$$
  $p=1/2$ 

So what we can do is use a histogram representation,

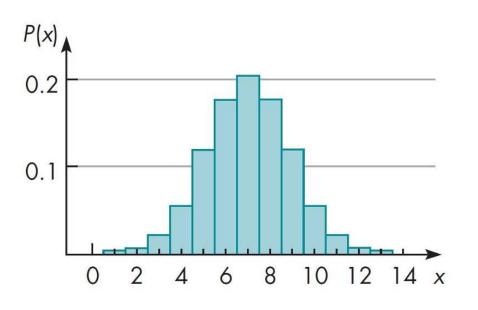


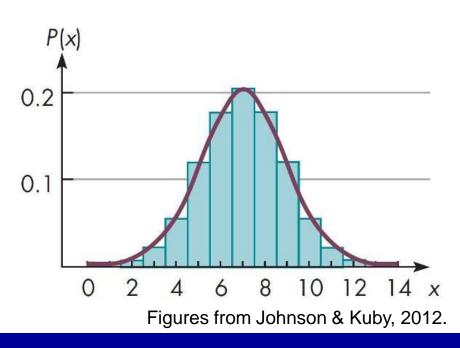


# 6: Normal Probability Distributions 6.5 Normal Approximation of the Binomial Distribution

So what we can do is use a histogram representation, p=1/2

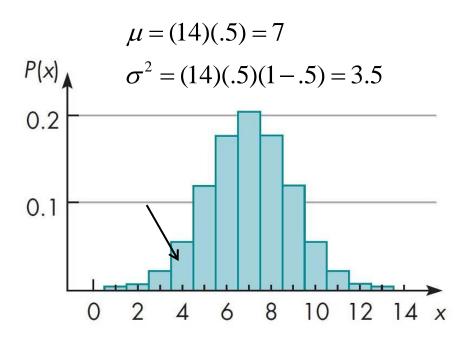
Then approximate binomial probabilities with normal areas.

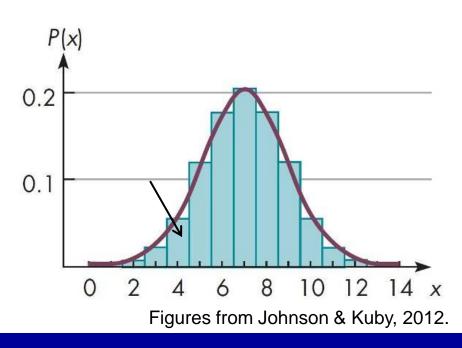




# 6: Normal Probability Distributions 6.5 Normal Approximation of the Binomial Distribution

Approximate binomial probabilities with normal areas. Use a normal with  $\mu = np$ ,  $\sigma^2 = np(1-p)$ 





### 6: Normal Probability Distributions

### 6.5 Normal Approximation of the Binomial Distribution

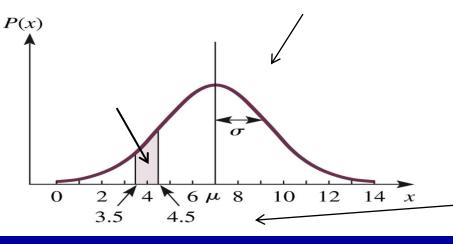
n=14, p=1/2

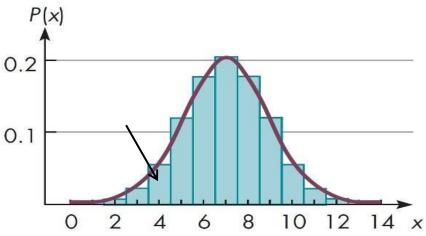
We then approximate binomial probabilities with normal areas.

P(x=4) from the binomial formula

is approximately P(3.5 < x < 4.5)

from the normal with  $\mu = 7$ ,  $\sigma^2 = 3.5$ 





the ±.5 is called a "continuity correction"

Figures from Johnson & Kuby, 2012.

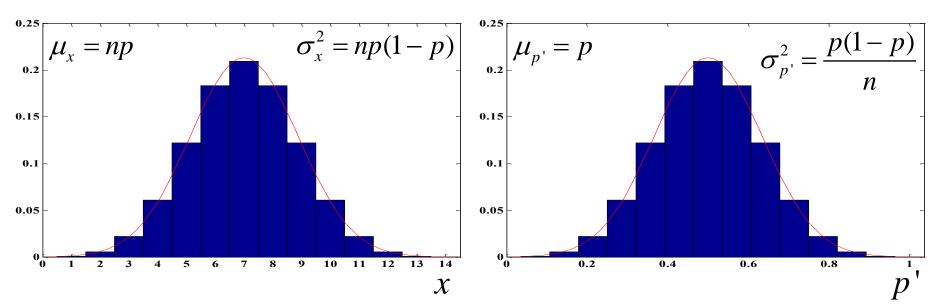
In practice, using these guidelines will ensure normality of x:

1. The sample size n is greater than 20.

- page 435
- 2. The product np and n(1-p) are both greater than 5.
- 3. The sample consists of less than 10% of the population.
  - 1.  $n \ge 20$ , 2.  $np \ge 5$  and  $n(1-p) \ge 5$ , 3.  $\frac{n}{N} < .10$ .

But we're not using x, we're scaling it and using  $p' = \frac{x}{n}$ .

It turns out that  $p' = \frac{x}{n}$  also has an approx. normal distribution.



n=14, p=1/2

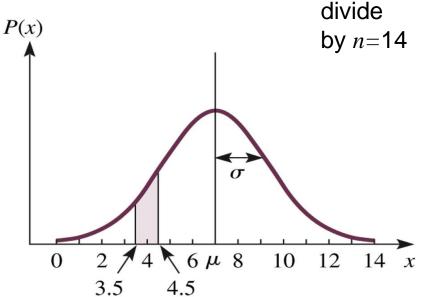
Now we can determine probabilities with normal areas.

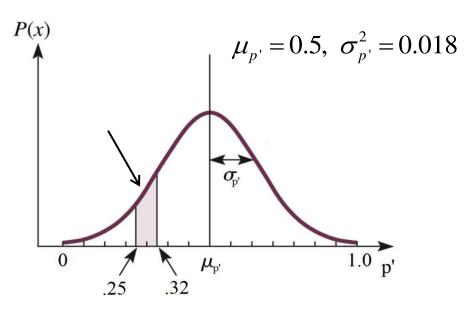
 $p' = \frac{x}{n}$ 

$$P(3.5 < x < 4.5) = 0.0594$$



$$P(3.5/14 < p' < 4.5/14) = 0.0594$$





Need to convert to z's.

Figure left from and right modified Johnson & Kuby, 2012.

n=14, p=1/2

Now we can determine probabilities with normal areas.  $p' = \frac{x}{n}$ 

$$p' = \frac{x}{n}$$

For x

$$P(3.5-7 < x-np < 4.5-7)$$

$$P\left(\frac{3.5-7}{\sqrt{3.5}} < \frac{x-np}{\sqrt{np(1-p)}} < \frac{4.5-7}{\sqrt{3.5}}\right)$$

For p'

$$P(.25 - .5 < p' - p < .32 - .5)$$

$$P\left(\frac{3.5-7}{\sqrt{3.5}} < \frac{x-np}{\sqrt{np(1-p)}} < \frac{4.5-7}{\sqrt{3.5}}\right) \qquad P\left(\frac{.25-.5}{\sqrt{\frac{.5(1-.5)}{14}}} < \frac{p'-p}{\sqrt{\frac{p(1-p)}{n}}} < \frac{.32-.5}{\sqrt{\frac{.5(1-.5)}{14}}}\right)$$

Now we can look up areas.

For a confidence interval, we would use

#### **Confidence Interval for a Proportion**

$$p'-z(\alpha/2)\sqrt{\frac{p'q'}{n}} \quad \text{to} \quad p'+z(\alpha/2)\sqrt{\frac{p'q'}{n}}$$
where  $p'=\frac{x}{n}$  and  $q'=(1-p')$ . (9.6)

Since we didn't know the true value for p, we estimate it by p'.

This is of the form  $point estimate \pm some amount$ .

#### **Example:**

Dana randomly selected n=200 cars and found x=17 convertibles. Find the 90% CI for the proportion of cars that are convertibles.

$$p' = \frac{x}{n} = \frac{17}{200}$$

$$\alpha = 0.1$$

$$z(\alpha / 2) = z(0.1 / 2) = 1.65$$

$$p' \pm z(\alpha / 2) \sqrt{\frac{p' q'}{n}}$$

$$\frac{17}{200} \pm 1.65 \sqrt{\frac{(17/200)(1 - 17/200)}{200}}$$

$$0.052 \text{ to } 0.118$$

# 9: Inferences Involving One Population 9.2 Inference about the Binomial Probability of Success Determining the Sample Size

Using the error part of the CI, we determine the sample size n.

#### **Maximum Error of Estimate for a Proportion**

$$E = z(\alpha/2)\sqrt{\frac{p'(1-p')}{n}}$$
(9.7)

#### Sample Size for 1- $\alpha$ Confidence Interval of p

$$n = \frac{\left[z(\alpha/2)\right]^2 p * (1-p*)}{E^2}$$
 From prior data, experience, gut feelings, séance. Or use 1/2. (9.8)

where  $p^*$  and  $q^*$  are provisional values used for planning.

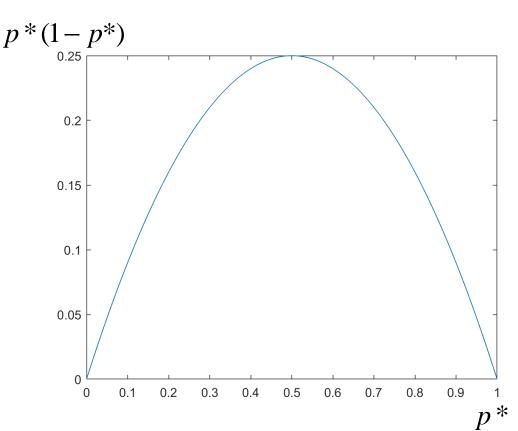
# 9: Inferences Involving One Population 9.2 Inference about the Binomial Probability of Success Determining the Sample Size

Why use  $p^* = 1/2$ ?

It makes  $p^*(1-p^*)$  largest and hence makes

$$n = \frac{[z(\alpha/2)]^2 p * (1-p^*)}{E^2}$$

the largest.



# 9: Inferences Involving One Population 9.2 Inference about the Binomial Probability of Success Determining the Sample Size

#### **Example:**

A supplier claims bolts are approx. 5% defective. Determine the sample size n if we want our estimate within ±0.02 with 90% confidence.

$$n = \frac{[z(\alpha/2)]^2 p^* (1-p^*)}{E^2} \qquad 1-\alpha = 0.90 \qquad E = 0.02$$
$$z(0.1/2) = 1.65 \qquad p^* = 0.05$$

$$n = \frac{[1.65]^2 (0.05)(1 - 0.05)}{(0.02)^2} = \frac{0.12931875}{0.0004} = 323.4 \implies n = 324$$

### **Chapter 9: Inferences Involving One Population**

Questions?

Homework: Read Chapter 9.2 WebAssign Homework Chapter 9 # 75, 89